

FORMULAE
Version July, 2017
Probability and Statistics
401-2604-00, Spring 2017

1. STANDARD DISCRETE DISTRIBUTIONS

- (1) **Bernoulli distribution with success parameter $p \in (0, 1)$.** $X \in \{0, 1\}$ and

$$P(X = 1) = p, \quad EX = p, \quad \text{Var}(X) = p(1 - p).$$

- (2) **Binomial distribution with n trials and success parameter $p \in (0, 1)$.**
 $X \in \{0, 1, \dots, n\}$

$$P(X = k) = \binom{n}{k} p^k (1 - p)^{n-k}, \quad k = 0, 1, \dots, n,$$

$$EX = np, \quad \text{Var}(X) = np(1 - p).$$

- (3) **Poisson distribution with parameter $\lambda > 0$.** $X \in \{0, 1, \dots\}$

$$P(X = k) = \frac{\lambda^k}{k!} e^{-\lambda}, \quad k = 0, 1, \dots,$$

$$EX = \lambda, \quad \text{Var}(X) = \lambda.$$

2. STANDARD CONTINUOUS DISTRIBUTIONS

- (4) **Gaussian distribution with mean μ and variance σ^2 .** $X \in \mathbb{R}$,

$$f_X(x) := \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{1}{2}\left(\frac{x - \mu}{\sigma}\right)^2\right], \quad x \in \mathbb{R}.$$

Denoted by $X \sim \mathcal{N}(\mu, \sigma^2)$.

$$EX = \mu, \quad \text{var}(X) = \sigma^2.$$

$$X \sim \mathcal{N}(\mu, \sigma^2) \Leftrightarrow Z := \frac{X - \mu}{\sigma} \sim \mathcal{N}(0, 1).$$

$\mathcal{N}(0, 1)$ is called the **standard normal (or Gaussian)**.

- (5) **The standard normal distribution function.**

$$\Phi(x) := \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-z^2/2} dz, \quad x \in \mathbb{R}.$$

Let Φ^{-1} be its inverse function. Then,

$$\Phi^{-1}(0.9) = 1.28, \quad \Phi^{-1}(0.95) = 1.64, \quad \Phi^{-1}(0.975) = 1.96.$$

- (6) **Exponential distribution with parameter $\lambda > 0$.** $X \in \mathbb{R}_+ := [0, \infty)$,

$$f_X(x) = \frac{1}{\lambda} e^{-x/\lambda}, \quad x \geq 0.$$

$$EX = \lambda, \quad \text{Var}(X) = \lambda^2.$$

Note: in many textbooks λ is replaced by $1/\lambda$.

- (7) **Gamma distribution with parameters α, λ .** $X \in \mathbb{R}_+ := [0, \infty)$,

$$f_X(x) = \frac{1}{\lambda^\alpha \Gamma(\alpha)} x^{\alpha-1} e^{-x/\lambda}, \quad x \geq 0.$$

Here $\Gamma(\alpha)$ is the Gamma function and for integer values $\Gamma(m) = (m-1)!$.

$$EX = \alpha\lambda, \quad \text{Var}(X) = \alpha\lambda^2.$$

Note: in many textbooks λ is replaced by $1/\lambda$.

- (8) **Beta distribution with parameters r, s .** $X \in [0, 1]$,

$$f_X(x) = \frac{\Gamma(r+s)}{\Gamma(r)\Gamma(s)} x^{r-1} (1-x)^{s-1}, \quad x \in [0, 1].$$

$$EX = \frac{r}{r+s}, \quad \text{Var}(X) = \frac{rs}{(r+s)^2 (1+r+s)}.$$

- (9) **Chi-Square (χ^2) distribution.**

The χ^2 distribution with m degrees of freedom is the Gamma distribution with parameters $\alpha = m/2$, $\lambda = 2$. Denoted by $\chi^2(m)$. In particular,

$$X \sim \mathcal{N}(0, 1) \quad \Rightarrow \quad X^2 \sim \chi^2(1),$$

$$X_j \sim \mathcal{N}(0, 1), \quad j = 1, \dots, m, \text{ i.i.d.} \quad \Rightarrow \quad \sum_{j=1}^m X_j^2 \sim \chi^2(m),$$

- (10) **Student distribution.**

If $Z \sim \mathcal{N}(0, 1)$, $Y \sim \chi^2(m)$, $Z \perp Y$, then,

$$T := \frac{Z}{\sqrt{Y/m}},$$

has a student distribution with m degrees of freedom.

Its density is given by

$$f_T(t) = \frac{\Gamma((m+1)/2)}{\sqrt{m\pi} \Gamma(m/2)} \left(1 + \frac{t^2}{m}\right)^{-(m+1)/2}, \quad t \in \mathbb{R}.$$

- (11) **Studentizing.** Let $\{X_i\}_{i=1}^n$ be i.i.d. with $\mathcal{N}(\mu, \sigma^2)$ distribution. Let $\bar{X}_n := \sum_{i=1}^n X_i/n$ and set

$$S_n^2 := \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X}_n)^2.$$

Then, $\bar{X}_n \perp S_n^2$ and

$$\frac{\sqrt{n} [\bar{X}_n - \mu]}{S_n}$$

has a Student distribution with $n-1$ degrees of freedom.

3. BOREL-CANTELLI

- (12) **Infinitely often, i.o.** For a given countable sequence of events $\{A_n\}_{n=1}^\infty$, the set $\{A_n \text{ i.o.}\}$ is defined by

$$\{A_n \text{ i.o.}\} := \bigcap_{n=1}^\infty \bigcup_{m=n}^\infty A_m.$$

- (13) **Borel-Cantelli Lemma 1.** Suppose that $\{A_n\}_{n=1}^\infty$ satisfy

$$\sum_n P(A_n) < \infty.$$

Then, $P(\{A_n \text{ i.o.}\}) = 0$.

- (14) **Borel-Cantelli Lemma 2.** Suppose that $\{A_n\}_{n=1}^\infty$ are mutually independent and satisfy

$$\sum_i P(A_n) = \infty.$$

Then, $P(\{A_n \text{ i.o.}\}) = 1$.

4. LIMIT THEOREMS

- (15) **Law of Large Numbers.** Let $\{X_i\}_{i=1}^\infty$ be an i.i.d. sequence. Set $\mu := EX_i$, $\sigma^2 := \text{Var}(X_i) < \infty$ for any i , and

$$\bar{X}_n := \frac{1}{n} \sum_{i=1}^n X_i.$$

Then the weak law of large numbers states that \bar{X}_n converges to μ in probability and the strong law states that the convergence is almost surely.

- (16) **Central Limit Theorem.** Let $\{X_i\}_{i=1}^\infty$ be an i.i.d. sequence. Set $\mu := EX_i$, $\sigma^2 := \text{Var}(X_i) < \infty$ for any i , and $\bar{X}_n := \sum_{i=1}^n X_i/n$. Let

$$Z_n := \frac{\sqrt{n}}{\sigma} [\bar{X}_n - \mu].$$

The distribution of Z_n converges to the standard Gaussian, i.e., for all $z \in \mathbb{R}$

$$\lim_{n \rightarrow \infty} P(Z_n \leq z) = \Phi(z),$$

where Φ is the standard normal distribution function.

5. INEQUALITIES

- (17) **Jensen's Inequality.** For a convex function $g : \mathbb{R} \rightarrow \mathbb{R}$,

$$g(EX) \leq Eg(X).$$

- (18) **Markov's Inequality.** For a non-negative random variable X and constant $a > 0$

$$P(X \geq a) \leq \frac{EX}{a}.$$

Generalized Chebyshev's Inequality. For a random variable $X \in \mathbb{R}$ and a non-negative increasing function g and a real number a with $g(a) > 0$,

$$P(X \geq a) \leq \frac{Eg(X)}{g(a)}.$$

6. MOMENTS

(19) **Variance and Standard Deviation.**

$$\begin{aligned} \text{variance of } X &= \text{Var}(X) := EX^2 - (EX)^2 = E(X - EX)^2, \\ \text{standard deviation of } X &= \sigma_X := \sqrt{\text{Var}(X)}. \end{aligned}$$

(20) **Covariance and Correlation.**

$$\begin{aligned} \text{covariance of } X \text{ and } Y &= \text{Cov}(X, Y) := EXY - (EX)(EY) \\ &= E(X - EX)(Y - EY), \\ \text{correlation of } X \text{ and } Y &= \text{Corr}(X, Y) := \frac{\text{Cov}(X, Y)}{\sigma_X \sigma_Y} \in [-1, 1]. \end{aligned}$$

We have

$$\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y) + 2\text{Cov}(X, Y).$$

(21) **Moment Generating Function.** The moment generating function of a random variable X is

$$\Psi(t) := \mathbb{E}[e^{tX}], \quad t \in \mathbb{R}.$$

The value of $\Psi(t)$ could be $+\infty$. When it is finite for t near the origin,

$$EX^k = \frac{d^k}{dt^k} \Psi(0), \quad k = 1, 2, \dots$$

7. CONFIDENCE INTERVAL FOR THE MEAN OF A NORMAL DISTRIBUTION

(22) **Two sided Confidence Interval.** Let X_1, \dots, X_n be i.i.d. $\mathcal{N}(\mu, \sigma^2)$ and

$$\bar{X}_n := \sum_{i=1}^n X_i/n, \quad S_n^2 := \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X}_n)^2.$$

Set

$$\begin{aligned} A &:= \bar{X}_n - T_{n-1}^{-1}(1 - \alpha/2) \frac{S_n}{\sqrt{n}}, \\ B &:= \bar{X}_n + T_{n-1}^{-1}(1 - \alpha/2) \frac{S_n}{\sqrt{n}}, \end{aligned}$$

where T_m is the c.d.f. of the Student distribution with m degrees of freedom, T_m^{-1} is its inverse function and $0 < \alpha < 1$. The interval (A, B) is a two sided $(1 - \alpha)$ -confidence interval for μ .

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Set

$$\begin{aligned} \underline{A} &:= \bar{X}_n - T_{n-1}^{-1}(1 - \alpha) \frac{S_n}{\sqrt{n}}, \\ \underline{B} &:= \bar{X}_n + T_{n-1}^{-1}(1 - \alpha) \frac{S_n}{\sqrt{n}}. \end{aligned}$$

where T_m is the c.d.f. of the Student distribution with m degrees of freedom, T_m^{-1} is its inverse function and $0 < \alpha < 1$.

The interval (\underline{A}, ∞) is an upper $(1 - \alpha)$ -confidence interval for μ .

The interval $(-\infty, \underline{B})$ is a lower $(1 - \alpha)$ -confidence interval for μ .

8. VARIOUS

- (24) **Change of Variables Formula.** Let $X = (X_1, \dots, X_n)$ have a continuous joint distribution $f_X(x)$ for $x \in \mathbb{R}^n$. Let $Y = AX$ for some non-singular square matrix A . Then, the probability distribution function of Y is given by

$$f_Y(y) = \frac{1}{|\det(A)|} f_X(A^{-1}y), \quad y \in \mathbb{R}^n.$$

- (25) **Continuous Bayes Theorem.** Consider two random variables X and θ , where θ has density $w(\cdot)$ and given $\theta = \vartheta$, the random variable X has density $f(x | \vartheta)$. Then the density of θ given $X = x$ is

$$w(\vartheta | x) = \frac{f(x | \vartheta) w(\vartheta)}{f(x)},$$

where

$$f(x) = \int f(x|\vartheta) w(\vartheta) d\vartheta.$$