

Exercise 6.1

Notation: $P^{T+\delta} = Q^{T+\delta}$
 $P^A = Q^A$

- (a) Recall and prove the numeraire change theorem.
- (b) Show that the forward Libor rate $(L(t, T, T+\delta))_t^1$ is a martingale under the forward probability $P^{T+\delta}$. Derive the arbitrage free price of a caplet with payoff $C^{caplet} = \delta(L(t, T, T+\delta) - K)^+$.
- (c) Using the formula

$$S_{T_0, T_N}(t) = \frac{P(t, T_0) - P(t, T_N)}{\sum_{k=0}^N \delta P(t, T_k)}$$

show that the swap rate $(S_{T_0, T_N}(t))_t$ is a martingale under the annuity probability P^A where the annuity is $A(t) = \sum_{k=0}^N \delta P(t, T_k)$. Derive the arbitrage free price of a swaption with payoff $C^{swaption} = A(T) (S_{T_0, T_N}(t) - K)^+$.

¹the forward zero-coupon $P(t, T, T+\delta) = \frac{1}{1+\delta L(t, T, T+\delta)}$ and $P(t, T, T+\delta) = \frac{P(t, T+\delta)}{P(t, T)}$.

a) Change of numeraire Thm: $\mathcal{M}_e(S^1) = \{\tilde{Q} \mid \frac{d\tilde{Q}}{dQ} = \frac{X_t^1}{X_0^1}, \text{ for } Q \in \mathcal{M}_e(S^0)\}$

where $X_t^d = \frac{S_t^d}{S_0^d}$

Proof: " \supset ": \leadsto take \tilde{Q} defined by $\frac{d\tilde{Q}}{dQ} = \frac{X_t^1}{X_0^1}$, for $Q \in \mathcal{M}_e(S^0)$

$\leadsto \tilde{Q} \ll Q$ holds by def. of \tilde{Q}

\leadsto since $S_t^0 > 0, S_t^1 > 0 \forall t \Rightarrow \frac{X_t^1}{X_0^1} = \left(\frac{S_t^0}{S_0^0}\right)^{-1} \cdot \left(\frac{S_t^1}{S_0^1}\right) > 0$

Claim 1: $Q \ll \tilde{Q}$

Proof 1: let $A \in \mathcal{F}$ with $\tilde{Q}(A) = 0$

$\Leftrightarrow 0 = \tilde{Q}(A) = \int_A \frac{X_t^1}{X_0^1} d\tilde{Q} \leadsto$ if $Q(A) > 0 \Rightarrow$ also $\int_A \frac{X_t^1}{X_0^1} dQ > 0$ has to hold (see Measure Theory) \square (Claim 1)

$\Rightarrow \tilde{Q} \sim Q \sim P \Rightarrow \tilde{Q} \sim P$, so \tilde{Q} is an equivalent measure

$\leadsto \tilde{Q}(\Omega) = E_Q\left(\frac{X_t^1}{X_0^1}\right) = E_Q\left(E_Q\left(\frac{X_t^1}{X_0^1} \mid \mathcal{F}_0\right)\right) = E_Q\left(\frac{X_0^1}{X_0^1}\right) = 1 \Rightarrow \tilde{Q}$ is equiv. prob.-meas.

Claim 2: $\left(\frac{S_t^1}{S_t^0}\right)_{t \geq 0}$ is a \tilde{Q} -martingale $\forall 0 \leq i \leq d \leadsto$ with Claim 2 we have that $\tilde{Q} \in \mathcal{M}_e(S^1)$

Proof 2: $\leadsto 0 \leq r \leq t \leq T \leadsto E_{\tilde{Q}}\left(\frac{S_t^1}{S_t^0} \mid \mathcal{F}_r\right) = \left(E_Q\left(\frac{d\tilde{Q}}{dQ} \mid \mathcal{F}_r\right)\right)^{-1} \cdot E_Q\left(\frac{S_t^1}{S_t^0} \cdot E_Q\left(\frac{d\tilde{Q}}{dQ} \mid \mathcal{F}_t\right) \mid \mathcal{F}_r\right) = \left(E_Q\left(\frac{X_t^1}{X_0^1} \mid \mathcal{F}_r\right)\right)^{-1}$

$\cdot E_Q\left(\frac{S_t^1}{S_t^0} \cdot E_Q\left(\frac{X_t^1}{X_0^1} \mid \mathcal{F}_t\right) \mid \mathcal{F}_r\right) = \left(\frac{X_t^1}{X_0^1}\right)^{-1} \cdot E_Q\left(\frac{S_t^1}{S_t^0} \cdot \frac{X_t^1}{X_0^1} \mid \mathcal{F}_r\right) = \left(\frac{X_0^1}{X_0^1}\right) \cdot E_Q\left(\frac{S_t^1}{S_t^0} \cdot \left(\frac{S_0^1}{S_0^0}\right)^{-1} \cdot \left(\frac{S_t^1}{S_0^1}\right) \mid \mathcal{F}_r\right) =$
 $= \frac{S_0^1}{S_0^0} \cdot \frac{S_r^0}{S_r^1} \cdot E_Q\left(\frac{X_t^1}{X_0^1} \mid \mathcal{F}_r\right) = \frac{S_0^1}{S_0^0} \cdot \frac{S_r^0}{S_r^1} \cdot \frac{S_r^1}{S_r^0} \cdot \frac{S_0^0}{S_0^1} = \frac{S_r^0}{S_r^1} \quad \square$ (Claim 2) \square (" \supset ")

" \subset ": \leadsto using " \supset " and exchanging S^1 & S^0 we have: $\{Q \mid \frac{dQ}{d\tilde{Q}} = \frac{Y_0^0}{Y_0^1}, \tilde{Q} \in \mathcal{M}_e(S^1)\} \subset \mathcal{M}_e(S^0)$, for

$Y_t^i = \frac{S_t^i}{S_t^0}$

$\Rightarrow \forall \tilde{Q} \in \mathcal{M}_e(S^1) : \exists Q \in \mathcal{M}_e(S^0) : \frac{dQ}{d\tilde{Q}} = \frac{S_0^0}{S_0^1} \cdot \frac{S_0^1}{S_0^0} \stackrel{\text{use that } (Q \sim \tilde{Q} \text{ see } \ast)}{\Rightarrow} \frac{d\tilde{Q}}{dQ} = \frac{S_0^1}{S_0^0} \cdot \frac{S_0^0}{S_0^1} = \frac{X_0^1}{X_0^0} \Rightarrow \tilde{Q} \in \text{RHS} \quad \square$ (" \subset ")

Proof (*): we have $\frac{dQ}{d\tilde{Q}} = 1$, by Radon-Nikodym Thm, we know that for $v < \tau < \lambda$ meqs. it holds that $\frac{dY}{dX} = \frac{dY}{d\tilde{X}} \cdot \frac{d\tilde{X}}{dX}$

(Proof: $v(A) = \int_A \frac{dY}{dX} d\lambda$, by R-N, $\int_A \frac{dY}{d\tilde{X}} \cdot \frac{d\tilde{X}}{dX} d\lambda = \int_A \frac{dY}{d\tilde{X}} d\tilde{\lambda} = v(A) \leadsto$ since it holds $\forall A$ mbl $\Rightarrow \square$)

\leadsto using this we have, since $Q \sim \tilde{Q} : 1 = \frac{dQ}{d\tilde{Q}} = \frac{dQ}{dQ} \cdot \frac{d\tilde{Q}}{dQ} \Leftrightarrow \frac{dQ}{dQ} = \left(\frac{d\tilde{Q}}{dQ}\right)^{-1}$ (and also the other way)

\leadsto this is well defined since $Q \sim \tilde{Q}$ implies that $\frac{d\tilde{Q}}{dQ} > 0, \frac{dQ}{d\tilde{Q}} > 0$ P-os. \square (*)

b) Proof: \leadsto the forward Libor rate is $(L(t, T_1, T+\delta))_t = \left[\frac{1}{\delta} \left(\frac{P(t, T)}{P(t, T+\delta)} - 1\right)\right]_t$

\leadsto the forward probability $Q^{T+\delta}$ is a measure $\in \mathcal{M}_e(P(\cdot, T+\delta)) \leadsto$ by Change of numeraire Thm such a meas. exists, since $P(\cdot, T+\delta) > 0$ and can therefore be used as numeraire

\leadsto our market is given by $(B_t, P(t, T_1), \dots, P(t, T_N))_{t \in \{T_0, \dots, T_N\}}$, where $T_0 < T_1 < \dots < T_N$, $\leadsto T_1, T+\delta \in \{T_0, \dots, T_N\}$

\leadsto if we want we can assume (by change of numeraire Thm) wlog. that $B_0 = 1$ (by replacing $(B_t)_t$ by $\left(\frac{B_t}{B_0}\right)_t$ as new numeraire)

\leadsto since $Q^{T+\delta} \in \mathcal{M}_e(P(\cdot, T+\delta))$ we have that $P(\cdot, T+\delta)$ is our numeraire and $\forall \tilde{T} \in \{T_1, \dots, T_N\} : \left(\frac{P(t, \tilde{T})}{P(t, T+\delta)}\right)_t$ is a $Q^{T+\delta}$ -martingale

\Rightarrow i.e.: $\left(\frac{P(t, T)}{P(t, T+\delta)}\right)_t$ is a $Q^{T+\delta}$ -martingale \Rightarrow also $(L(t, T_1, T+\delta))_t$ is a $Q^{T+\delta}$ -martingale \checkmark

\leadsto the arbitrage free price of C_t^{caplet} , (a payoff at time $T+\delta$) is given by (using that $Q \in \mathcal{M}_e((B_t)_t)$)

$$E_Q\left(\frac{C_t^{caplet}}{B_{T+\delta}}\right) = E_{Q^{T+\delta}}\left(\frac{C_t^{caplet}}{B_{T+\delta}} \cdot \frac{dQ}{dQ^{T+\delta}}\right) = E_{Q^{T+\delta}}\left(\frac{C_t^{caplet}}{B_{T+\delta}} \cdot \frac{B_{T+\delta}}{P(T+\delta, T+\delta)} \cdot \frac{P(0, T+\delta)}{B_0}\right) = \frac{P(0, T+\delta)}{B_0} E_{Q^{T+\delta}}(C_t^{caplet}) =$$

 $= \frac{P(0, T+\delta)}{B_0} E_{Q^{T+\delta}}[\delta(L(t, T_1, T+\delta) - K)^+]$

c) Proof: \leadsto since $\forall 0 \leq k \leq N$: $P(t, T_k)$ is an asset in the market, we can extend our market s.t.

also $A(t) = \sum_{k=0}^N \alpha^k P(t, T_k)$ is an asset, without losing (NA), (since we can't generate new value processes)

\leadsto since $P(\cdot, T_k) > 0 \forall k \Rightarrow A(\cdot) > 0 \Rightarrow$ by Change of numeraire Thm, we can chose $A(t)$ as numeraire and we find a meas. $Q^A \in \mathcal{M}_e(A)$

\leadsto under Q^A we have that: $\forall k: \left(\frac{P(t, T_k)}{A(t)} \right)_t$ is a Q^A -martingale

\Rightarrow also $(S_{T_0, T_N}(t))_t = \left(\frac{P(t, T_0)}{A(t)} - \frac{P(t, T_N)}{A(t)} \right)_t$ is a Q^A -martingale, as sum of 2 Q^A -mart. \checkmark

\leadsto the arbitrage-free price of $C_{T_f}^{sw}$ (a payoff at time T_f) is given by (using again the ENN) $Q \in \mathcal{M}_e(B_t)_t$:

$$\begin{aligned} E_Q \left(\frac{C_{T_f}^{sw}}{B_{T_f}} \right) &= E_{Q^A} \left(\frac{C_{T_f}^{sw}}{B_{T_f}} \cdot \frac{dQ}{dQ^A} \right) = E_{Q^A} \left(\frac{C_{T_f}^{sw}}{B_{T_f}} \cdot \frac{B_{T_f}}{A(T_f)} \cdot \frac{A(0)}{B_0} \right) = \frac{A(0)}{B_0} \cdot E_{Q^A} \left[\frac{A(T_f)}{A(T_f)} \cdot (S_{T_0, T_N}(t) - K)^+ \right] \\ &= \frac{A(0)}{B_0} E_{Q^A} \left[(S_{T_0, T_N}(t) - K)^+ \right] \end{aligned}$$

Exercise 6.2

Derive a formula for the arbitrage free price of following contingent claims.

(a) up-and-out call option, $C_{u&o}^{call} = (S_T - K)^+ \mathbb{1}_{\max_{0 \leq t \leq T} S_t < B}$

(b) down-and-in put option, $C_{d&i}^{put} = (K - S_T)^+ \mathbb{1}_{\min_{0 \leq t \leq T} S_t \leq B}$

(c) lookback put option, $C_{max}^{put} = \max_{0 \leq t \leq T} S_t - S_T$

(d) lookback call option, $C_{min}^{call} = S_T - \min_{0 \leq t \leq T} S_t$

\leadsto we assume to have an Binomial model $((1+r)^t, S_t)_{t=0, \dots, T}$, with $S_t = S_0 \hat{b}^{z_t}$

with $z_t = \sum_{i=1}^t y_i$, where $\omega \in \Omega = [-1, 1]^T$, $\omega = (y_1, \dots, y_T)$

\leadsto let Q be the EMM with $q = Q(y_i = 1)$, $(1-q) = Q(y_i = -1)$

$\leadsto Q(Z_T = k) = \begin{cases} q^{\frac{k+T}{2}} (1-q)^{\frac{T-k}{2}} \cdot \binom{T}{\frac{k+T}{2}}, & \text{if } k \text{ is even} \\ 0, & \text{else} \end{cases}$

\leadsto let $\tilde{M}_t := \max_{0 \leq s \leq t} S_s$, $M_t := \max_{0 \leq s \leq t} Z_s$, $\tilde{N}_t := \min_{0 \leq s \leq t} S_s$, $N_t := \min_{0 \leq s \leq t} Z_s$

a) SOL: the arbitrage free price of $C_{u&o}^{call} =: C$ is given by $E_Q \left(\frac{C}{(1+r)^T} \right) = \frac{1}{(1+r)^T} E_Q(C)$

\leadsto since we have that $E_Q(C_{u&o}^{call}) + E_Q(C_{d&i}^{call}) = \frac{1}{(1+r)^T} E_Q((S_T - K)^+ (\mathbb{1}_{\max_{0 \leq t \leq T} S_t < B} + \mathbb{1}_{\max_{0 \leq t \leq T} S_t \geq B})) = \frac{1}{(1+r)^T} E_Q((S_T - K)^+)$

$$\Rightarrow E_Q(C_{u&o}^{call}) = E_Q \left(\frac{(S_T - K)^+}{(1+r)^T} \right) - E_Q(C_{d&i}^{call})$$

$\leadsto E_Q(C_{u&i}^{call})$ can be computed as done in lecture $\Rightarrow E_Q \left(\frac{C_{u&i}^{call}}{(1+r)^T} \right) = \frac{1}{(1+r)^T} \cdot (E_Q[(S_T - K)^+ \mathbb{1}_{S_T \geq B}] + \left(\frac{q}{1-q} \right)^k \left(\frac{B}{S_0} \right)^2$

$\cdot E_Q[(S_T - \tilde{K})^+ \mathbb{1}_{S_T < \tilde{B}}])$, where $B \in \{S_0 \hat{b}^j | j \in \{-T, -T+2, \dots, -T\}\}$ and k s.t. $B = S_0 \hat{b}^k$, $\tilde{K} = K \hat{b}^{-2k} = K \left(\frac{S_0}{B} \right)^2$, $\tilde{B} := \frac{S_0^2}{B}$

$\leadsto E_Q \left(\frac{(S_T - K)^+}{(1+r)^T} \right) = \frac{1}{(1+r)^T} \sum_{n=0}^T (S_0 \hat{b}^{T-2n} - K)^+ q^{T-n} (1-q)^n \binom{T}{T-n}$, same can be done for $E_Q(C_{u&i}^{call})$

b) SOL: \leadsto the arbitrage free price for $C_{d&i}^{put}$ is given by: $E_Q \left(\frac{C_{d&i}^{put}}{(1+r)^T} \right) = \frac{1}{(1+r)^T} E_Q((K - S_T)^+ \mathbb{1}_{\tilde{N}_T \leq B})$

\leadsto wlog: $B \in \{S_0 \hat{b}^j | j \in \{-T, -T+2, \dots, T\}\}$ \leadsto let $k \in \mathbb{N}$ s.t. $B = S_0 \hat{b}^k$

$\leadsto E_Q((K - S_T)^+ \mathbb{1}_{\tilde{N}_T \leq B}) = E_Q((K - S_T)^+ \mathbb{1}_{S_T \leq B}) + E_Q((K - S_T)^+ \mathbb{1}_{S_T > B} \mathbb{1}_{\tilde{N}_T \leq B})$

$\leadsto E_Q((K - S_T)^+ \mathbb{1}_{S_T \leq B}) = \sum_{j \in \{-T, -T+2, \dots, k\}} (K - S_0 \hat{b}^j)^+ \cdot Q(Z_T = j) = \sum_{j \in \{-T, -T+2, \dots, k\}} (K - S_0 \hat{b}^j)^+ \cdot q^{\frac{T+j}{2}} (1-q)^{\frac{T-j}{2}} \binom{T}{\frac{T+j}{2}}$

$\leadsto E_Q((K - S_T)^+ \mathbb{1}_{S_T > B} \mathbb{1}_{\tilde{N}_T \leq B}) = \sum_{l \geq 1} E_Q((K - S_0 \hat{b}^{k+2l})^+ \mathbb{1}_{Z_T = k+l} \mathbb{1}_{N_T \leq k}) = \sum_{l \geq 1} (K - S_0 \hat{b}^{k+2l})^+ \cdot Q(Z_T = k+l, N_T \leq k) =: (I)$

$\leadsto Q(Z_T = k+l, \min_{t \leq T} Z_t \leq k) = \left(\frac{1-q}{q} \right)^{-l} Q(Z_T = k-l) = \left(\frac{1-q}{q} \right)^{-l}$

$\cdot q^{\frac{T+k+l}{2}} (1-q)^{\frac{T-k-l}{2}} \binom{T}{\frac{T+k+l}{2}} = \left(\frac{1-q}{q} \right)^{-l} q^{\frac{T-k+l}{2}} (1-q)^{\frac{T-k-l}{2}} \cdot q^{k-l} (1-q)^{l-k} (1-q)^{-l} q^l = Q(Z_T = -k+l) \cdot \left(\frac{q}{1-q} \right)^k$

$$\Rightarrow (I) = \sum_{l \geq 1} (K - S_0 \hat{b}^{k+2l})^+ \cdot Q(Z_T = -k+l) \left(\frac{q}{1-q} \right)^k = \left(\frac{q}{1-q} \right)^k \sum_{l \geq 1} \hat{b}^{2k} (K \hat{b}^{-2k} - S_0 \hat{b}^{-k+2l})^+$$

$$\cdot Q(Z_T = -k+l) = \left(\frac{q}{1-q} \right)^k \hat{b}^{2k} \cdot E_Q((\hat{K} - S_0 \hat{b}^{2T})^+ \mathbb{1}_{Z_T > -k}) = \left(\frac{q}{1-q} \right)^k \hat{b}^{2k} \cdot E_Q((\hat{K} - S_T)^+ \mathbb{1}_{S_T > S_0 \hat{b}^{-k}}) =$$

$$= \left(\frac{q}{1-q} \right)^k \hat{b}^{2k} \sum_{j \in \{-k+2, -k+4, \dots, T\}} (K - S_0 \hat{b}^j)^+ \cdot q^{\frac{T+j}{2}} (1-q)^{\frac{T-j}{2}} \binom{T}{\frac{T+j}{2}}$$

\leadsto combining those calculations yields the formula

c), d) \leadsto see book Exa 5.53

Exercise 6.3 We consider a binomial market model with N periods on a period of time of length T . The riskless asset grows at a rate $r = \frac{RT}{N}$, where R is the (constant) instantaneous interest rate, and the risky asset's price goes up by a factor $1+u$ and down by a factor $1+d$ such that

$$\log\left(\frac{1+u}{1+r}\right) = -\log\left(\frac{1+d}{1+r}\right) = \sigma\sqrt{\frac{T}{N}},$$

for some constant σ . The starting values (at time $t=0$) of both the assets is 1 \mathbb{P}^* -a.s. The unique equivalent martingale measure \mathbb{P}^* for S^1 is such that the $(Y_i)_{i \in \{1,2,\dots,N\}}$ are i.i.d. and given by

$$\mathbb{P}^*[Y_i = 1+d] = 1 - \mathbb{P}^*[Y_i = 1+u] = \frac{u-r}{u-d} = p^*$$

We study the limiting case for $N \rightarrow \infty$.

(a) Let $(Z_n)_{n \in \mathbb{N}}$ be a sequence of random variables of the form :

$$Z_n = \sum_{i=1}^n X_i^n$$

for $n \in \mathbb{N}$, $X_i^n \in \{-\sigma\sqrt{\frac{T}{n}}, \sigma\sqrt{\frac{T}{n}}\}$ and the variables $(X_i^n)_{i \in \{1,2,\dots,n\}}$ are independent identically distributed with mean μ_n . The constants μ_n are such that $\lim_{n \rightarrow \infty} n\mu_n = \mu$.

Prove that the sequence $(Z_n)_{n \in \mathbb{N}}$ converges in law to a gaussian random variable with mean μ and variance $\sigma^2 T$.

Hint: Use the fact that point-wise convergence of the characteristic functions of a sequence of random variables (if the limiting function ϕ is continuous at 0) implies the convergence in law of this sequence of random variables to a random variable whose characteristic function is ϕ .

(b) We consider a European put option, with strike K and maturity T . Show that its value at time 0 is given by

$$V_0^{P,N} = \mathbb{E}^* \left[\left(\frac{K}{(1+r)^N} - S_0^1 \exp(Z_N) \right)^+ \right],$$

where \mathbb{E}^* denotes the expectation under \mathbb{P}^* , and Z_N is a random variable that you will define.

(c) Use part a) to prove the following asymptotic price :

$$\lim_{N \rightarrow \infty} V_0^{P,N} = Ke^{-RT} \Phi(-d_2) - S_0^1 \Phi(-d_1),$$

Notation:
 $S_t^u := S_t^1$
 $B_t := S_t^r$

where $d_1 = \frac{\log(\frac{S_0^1}{K}) + RT + \frac{\sigma^2 T}{2}}{\sigma\sqrt{T}}$, $d_2 = d_1 - \sigma\sqrt{T}$ and Φ is the cumulative distribution function of a standard normal random variable.

Hint: Use the value of p^* and of u to prove that $\lim_{N \rightarrow \infty} N \mathbb{E}^* \left[\log\left(\frac{Y_i}{1+r}\right) \right] = -\frac{\sigma^2 T}{2}$

a) Proof: \rightarrow use Lindeberg-Feller Thm (Thm 2.24 Prob. Th.)

\rightarrow define $\tilde{X}_i^n := X_i^n - \mu_n \Rightarrow \mathbb{E}(\tilde{X}_i^n) = 0$

$$\rightarrow \text{Var}(\tilde{X}_i^n) = \text{Var}(X_i^n) = \mathbb{E}((X_i^n)^2) - \mathbb{E}(X_i^n)^2 = (\sigma^2 \frac{T}{n} p + \sigma^2 \frac{T}{n} (1-p)) - \mu_n^2 = \sigma^2 \frac{T}{n} - \mu_n^2$$

$$\Rightarrow \lim_{n \rightarrow \infty} \sum_{i=1}^n \mathbb{E}(\tilde{X}_i^n^2) = \lim_{n \rightarrow \infty} n \cdot (\sigma^2 \frac{T}{n} - \mu_n^2) = \sigma^2 T - \lim_{n \rightarrow \infty} n \mu_n^2 = \sigma^2 T, \text{ (since } n\mu_n \xrightarrow{n \rightarrow \infty} \mu \Rightarrow \mu_n \xrightarrow{n \rightarrow \infty} 0 \Rightarrow n\mu_n^2 = \mu\mu_n \xrightarrow{n \rightarrow \infty} \mu \cdot 0 = 0)$$

$\rightarrow \forall \varepsilon > 0$ choose n_0 so big such that : $\sigma\sqrt{\frac{T}{n_0}} < \frac{\varepsilon}{2}, \mu_n < \frac{\varepsilon}{2} \forall n \geq n_0$

$$\Rightarrow \forall n \geq n_0 : |\tilde{X}_i^n| = \sigma\sqrt{\frac{T}{n}} + |\mu_n| < \varepsilon$$

$$\Rightarrow \lim_{n \rightarrow \infty} \sum_{i=1}^n \mathbb{E}(\tilde{X}_i^n^2 \mathbb{1}_{\{|\tilde{X}_i^n| > \varepsilon\}}) = 0$$

\rightarrow conditions for Lindeberg-Feller fulfilled \Rightarrow

$$\sum_{i=1}^n \tilde{X}_i^n \xrightarrow[n \rightarrow \infty]{d} N(0, \sigma^2 T)$$

$$\Leftrightarrow \lim_{n \rightarrow \infty} Z_n - \mu = \lim_{n \rightarrow \infty} \left(\sum_{i=1}^n X_i^n - n\mu_n \right) \stackrel{d}{=} N(0, \sigma^2 T)$$

$$\Leftrightarrow \lim_{n \rightarrow \infty} Z_n \stackrel{d}{=} N(\mu, \sigma^2 T)$$

b) Proof: $\rightarrow C^A = (K - S_N)^+$

\rightarrow its price is given (using $Q = \mathbb{P}^* \in \mathcal{U}_e(B_t)$) by

$$\mathbb{E}^* \left(\frac{C^A}{B_N} \right) = \mathbb{E}^* \left(\frac{1}{(1+r)^N} \cdot (K - S_N)^+ \right) = \mathbb{E}^* \left(\left(\frac{K}{(1+r)^N} - \frac{S_N}{(1+r)^N} \right)^+ \right) \quad (1)$$

\rightarrow Let $R_t = \begin{cases} u, & \text{if price goes up from time } t-1 \text{ to } t \\ d, & \text{else} \end{cases}$ as defined in lecture $\rightarrow S_N = S_0 \cdot \prod_{t=1}^N (1+R_t) = (1+r)^N \cdot S_0 \cdot \prod_{t=1}^N \frac{1+R_t}{1+r} = (1+r)^N \cdot S_0 \cdot \prod_{t=1}^N \exp\left(\log\left(\frac{1+R_t}{1+r}\right)\right)$ \rightarrow define $X_t := \log\left(\frac{1+R_t}{1+r}\right)$ \rightarrow let $Z_t := \sum_{i=1}^t X_i$ \rightarrow we have $1+R_t = Y_t$

$$\Rightarrow S_N = (1+r)^N \cdot S_0 \cdot \prod_{t=1}^N \exp(X_t) = (1+r)^N \cdot S_0 \cdot \exp\left(\sum_{t=1}^N X_t\right) = (1+r)^N \cdot S_0 \cdot \exp(Z_N) \quad (2)$$

$$\Rightarrow V_0^{P,N} = \mathbb{E}^* \left(\frac{C^A}{B_N} \right) \stackrel{(1),(2)}{=} \mathbb{E}^* \left[\left(\frac{K}{(1+r)^N} - S_0 \exp(Z_N) \right)^+ \right]$$

c) Proof: $\lim_{N \rightarrow \infty} V_0^{P,N} \stackrel{(b)}{=} \lim_{N \rightarrow \infty} \mathbb{E}^* \left(\left(\frac{K}{(1+r)^N} - S_0 \exp(Z_N) \right)^+ \right) \quad (6)$

\rightarrow since $S_N = S_0 \exp(Z_N) \geq 0$ (for $S_0 \geq 0$) and $\lim_{N \rightarrow \infty} (1+r)^N = \lim_{N \rightarrow \infty} (1 + \frac{RT}{N})^N = e^{RT} \in (0, \infty) \Rightarrow \left(\frac{1}{(1+r)^N} \right)_{N \in \mathbb{N}}$ is bounded (at least when starting from some $N_0 \in \mathbb{N}$ big enough)

$$\Rightarrow 0 \leq \left(\frac{K}{(1+r)^N} - S_0 \exp(Z_N) \right)^+ \leq \frac{K}{(1+r)^N} \leq \pi \quad \forall N \geq N_0, \text{ where } \pi < \infty \text{ is some const.} \Rightarrow \text{the integrand can be dominated} \quad (5)$$

\rightarrow by assumption on our market we have that $Z_N = \sum_{i=1}^N X_i^N$ with $X_i^N = \log\left(\frac{1+R_i}{1+r}\right) \in \left\{ \pm \sigma\sqrt{\frac{T}{N}} \right\}$, with $(X_i^N)_{i=1,\dots,N}$ iid (since R_i are iid) $\rightarrow \mathbb{E}(X_i^N) = -\sigma\sqrt{\frac{T}{N}} \cdot p^* + \sigma\sqrt{\frac{T}{N}} (1-p^*) = \sigma\sqrt{\frac{T}{N}} - 2p^* \sigma\sqrt{\frac{T}{N}} =: \mu_N$ \rightarrow obviously also u, d, p^* depend on N

$$\rightarrow \lim_{N \rightarrow \infty} N\mu_N = \lim_{N \rightarrow \infty} \sigma\sqrt{TN} - 2p^* \cdot \sigma\sqrt{TN} = (\mu) \quad \rightarrow p^* = \frac{uN - rN}{uN - dN}$$

$$\rightarrow \log\left(\frac{1+u}{1+r}\right) = \sigma\sqrt{\frac{T}{N}} \Leftrightarrow \frac{1+u}{1+r} = \exp(\sigma\sqrt{\frac{T}{N}}) \Leftrightarrow uN = (1+rN) \exp(\sigma\sqrt{\frac{T}{N}}) - 1 \quad \rightarrow \text{similar we get } dN = (1+rN) \exp(-\sigma\sqrt{\frac{T}{N}}) - 1$$

$$\Rightarrow p_N^* = \frac{(1+rN) \exp(\sigma\sqrt{\frac{T}{N}}) - 1 - rN}{(1+rN) (\exp(\sigma\sqrt{\frac{T}{N}}) - \exp(-\sigma\sqrt{\frac{T}{N}}))} = \frac{\exp(\sigma\sqrt{\frac{T}{N}}) - 1}{\exp(\sigma\sqrt{\frac{T}{N}}) - \exp(-\sigma\sqrt{\frac{T}{N}})} \quad \rightarrow (1-p_N^*) = \frac{1 - \exp(-\sigma\sqrt{\frac{T}{N}})}{\exp(\sigma\sqrt{\frac{T}{N}}) - \exp(-\sigma\sqrt{\frac{T}{N}})}$$

$$\rightarrow \text{with L'Hospital we have: } \lim_{N \rightarrow \infty} N\mu_N = \lim_{h \rightarrow 0} \frac{-\sigma\sqrt{TN} \exp(\sigma\sqrt{\frac{T}{N}}) + \sigma\sqrt{TN}}{\exp(\sigma\sqrt{\frac{T}{N}}) - \exp(-\sigma\sqrt{\frac{T}{N}})} + \frac{\sigma\sqrt{TN} - \sigma\sqrt{TN} \exp(-\sigma\sqrt{\frac{T}{N}})}{\exp(\sigma\sqrt{\frac{T}{N}}) - \exp(-\sigma\sqrt{\frac{T}{N}})} = \lim_{h \rightarrow 0} \frac{2\sigma\sqrt{TN} - \sigma\sqrt{TN} (\exp(\sigma\sqrt{\frac{T}{N}}) + \exp(-\sigma\sqrt{\frac{T}{N}}))}{\exp(\sigma\sqrt{\frac{T}{N}}) - \exp(-\sigma\sqrt{\frac{T}{N}})}$$

$$\stackrel{h = \frac{1}{\sqrt{N}}, \eta = \sigma\sqrt{T}}{=} \lim_{h \rightarrow 0} \frac{2\eta h^2 - \eta h^2 (\exp(\eta h) + \exp(-\eta h))}{\exp(\eta h) - \exp(-\eta h)} = \lim_{h \rightarrow 0} \frac{2\eta \cdot (1 - \cosh(\eta h))}{h \cdot 2 \sinh(\eta h)} \stackrel{\text{L'Hospital}}{=} \lim_{h \rightarrow 0} \frac{(-\eta^2 \sinh(\eta h))}{\sinh(\eta h) + h \eta \cosh(\eta h)} \stackrel{\text{L'Hospital}}{=} \lim_{h \rightarrow 0} \frac{(-\eta^3 \cosh(\eta h))}{\eta \cosh(\eta h) + \eta \cosh(\eta h) + h \eta^2 \sinh(\eta h)} =$$

$$= -\frac{\eta^2}{2} = -\frac{\sigma^2 T}{2} =: \mu \in \mathbb{R}$$

$$\Rightarrow \text{the conditions for a) are fulfilled} \Rightarrow Z_N \xrightarrow[N \rightarrow \infty]{d} Z \sim N(\mu, \sigma^2 T) \quad (3)$$

\rightarrow We know from Prop. Th. (Thm 2.7) that (3) implies the existence of R.V. \tilde{Z}_N s.t. $\tilde{Z}_N \sim Z_N$ and s.t. $\tilde{Z}_N \xrightarrow{P.O.S} Z$ (4)

$$\Rightarrow \lim_{N \rightarrow \infty} V_0^{P,N} \stackrel{(6)}{=} \lim_{N \rightarrow \infty} \mathbb{E}^* \left(\left(\frac{K}{(1+r)^N} - S_0 \exp(Z_N) \right)^+ \right) \stackrel{\substack{\mathbb{E}^*(\cdot) \text{ depends only} \\ \text{on the law of } Z_N}}{=} \lim_{N \rightarrow \infty} \mathbb{E}^* \left(\left(\frac{K}{(1+r)^N} - S_0 \exp(\tilde{Z}_N) \right)^+ \right) \stackrel{\substack{\text{with (5) \& (4) we can} \\ \text{use the dom. conv., since} \\ \text{the random } f \text{ is contin.}}}{=} \mathbb{E}^* \left(\left(\frac{K}{(1+r)^N} - S_0 \exp(\tilde{Z}_N) \right)^+ \right) = f(Z_N)$$

$$= \mathbb{E}^* \left(\left(\lim_{N \rightarrow \infty} \left(\frac{K}{(1+r)^N} - S_0 \exp(\tilde{Z}_N) \right)^+ \right) \right) = \mathbb{E}^* \left(\left(\frac{K}{e^{RT}} - S_0 \exp(Z) \right)^+ \right) = \mathbb{E}^* \left[\left(Ke^{-RT} - S_0 \exp(\mu + \sigma\sqrt{T}W) \right)^+ \right] =$$

$Z \sim N(\mu, \sigma^2 T) \Leftrightarrow \frac{Z - \mu}{\sigma\sqrt{T}} =: W \sim N(0,1) \Leftrightarrow Z = \sigma\sqrt{T}W + \mu$

$$= \int (K e^{-RT} - S_0 \exp(\mu + \sigma \sqrt{T} \omega)) \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{\omega^2}{2}\right) d\omega =: A + B =: (II)$$

$$\{\omega \in \mathbb{R} \mid S_0 \exp(\mu + \sigma \sqrt{T} \omega) \leq K e^{-RT}\} =: D$$

$$\leadsto S_0 \exp(\mu + \sigma \sqrt{T} \omega) \leq K e^{-RT} \Leftrightarrow \mu + \sigma \sqrt{T} \omega \leq \log\left(\frac{K e^{-RT}}{S_0}\right) = \log\left(\frac{K}{S_0}\right) - RT \Leftrightarrow \omega \leq \frac{\log(K/S_0) - RT - \mu}{\sigma \sqrt{T}} =: -d_2$$

$$\Rightarrow A = \int_D K e^{-RT} \cdot \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{\omega^2}{2}\right) d\omega = K e^{-RT} \int_{-\infty}^{-d_2} \frac{\exp\left(-\frac{\omega^2}{2}\right)}{\sqrt{2\pi}} d\omega = K e^{-RT} \Phi(-d_2) \quad (7)$$

$$\leadsto B = -\int_D S_0 \exp(\mu + \sigma \sqrt{T} \omega) \exp\left(-\frac{\omega^2}{2}\right) \cdot \frac{1}{\sqrt{2\pi}} d\omega = -S_0 \int_{-\infty}^{-d_2} \exp\left(-\frac{1}{2}(\omega^2 - 2\sigma \sqrt{T} \omega + \sigma^2 T)\right) \frac{1}{\sqrt{2\pi}} d\omega$$

$\mu = -\frac{\sigma^2 T}{2}$ $= (\omega - \sigma \sqrt{T})^2 = x^2$ substitution $x = \omega - \sigma \sqrt{T}$

$$= -S_0 \int_{-\infty}^{-d_2 - \sigma \sqrt{T}} \exp\left(-\frac{x^2}{2}\right) \frac{1}{\sqrt{2\pi}} dx = -S_0 \cdot \Phi(-d_1) \quad (8)$$

$$\Rightarrow \lim_{N \rightarrow \infty} V_0^P N = (II) = A + B = K e^{-RT} \Phi(-d_2) - S_0 \Phi(-d_1)$$

with $d_1 = d_2 + \sigma \sqrt{T} = \frac{\log(S_0/K) + RT - \frac{\sigma^2 T}{2}}{\sigma \sqrt{T}} + \sigma \sqrt{T} = \frac{\log(S_0/K) + RT + \frac{\sigma^2 T}{2}}{\sigma \sqrt{T}}$

$$= \frac{\log(S_0/K) + RT + \frac{\sigma^2 T}{2}}{\sigma \sqrt{T}}$$

Exercise 6.4 Python - Trinomial model Inspire yourself from binomial price array to complete the trinomial price function.

the trinomial model is given by 2 assets. 1 is the numeraire $(B_t)_t$ (e.g.: $B_t = e^{rt}$) and the other is S_t with S_0 fixed and $S_{t+1} = S_t \cdot Y_{t+1}$, with $Y_{t+1} \in \{d, 1, u\}$ a RV.

we choose an (EMM) Q s.t. all Y_t are iid and we need to have that $\frac{S_t}{B_t}$ is a martingale, i.e.:

$$\mathbb{E}_Q\left(\frac{S_t}{B_t} \mid \mathcal{F}_{t-1}\right) = \frac{S_{t-1}}{B_{t-1}} \Leftrightarrow \frac{S_t}{B_t} = \mathbb{E}_Q\left(\frac{S_{t-1} \cdot Y_t}{B_{t-1}} \mid \mathcal{F}_{t-1}\right) = \frac{1}{B_t} \cdot S_{t-1} \mathbb{E}_Q(Y_t) \Leftrightarrow e^r = d q_1 + 1 \cdot (1 - q_1 - q_2) + u \cdot q_2$$

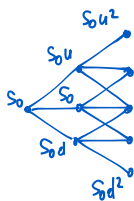
we assume that $d = \frac{1}{u}$ and $q_1, q_2, 1 - q_1 - q_2 > 0$

obviously Q is not unique

one possibility is to use (see paper on trinomial trees): $u = e^{\sigma \sqrt{2\Delta t}}$, $d = e^{-\sigma \sqrt{2\Delta t}}$, $B_n = e^{r \Delta t n}$, $q_u = \left(\frac{e^{r \Delta t/2} - e^{-\sigma \sqrt{2\Delta t}}}{e^{\sigma \sqrt{2\Delta t}} - e^{-\sigma \sqrt{2\Delta t}}}\right)^2$

$$p_d = \left(\frac{e^{-\sigma \sqrt{2\Delta t}} - e^{r \Delta t/2}}{e^{\sigma \sqrt{2\Delta t}} - e^{-\sigma \sqrt{2\Delta t}}}\right)^2, \quad p_m = 1 - p_u - p_d$$

since $d = \frac{1}{u}$ we get 2 new possible values in each step \leadsto a corresponding tree looks like:



if we have n steps, i.e. S_{n-1} is the value at maturity then we know that S_{n-1} can take $1+2(n-1)$ values

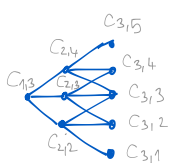
we want to find a recursion formula for the prices of an option C

we want to find arbitrage free prices wrt. our fixed (EMM) $Q \Rightarrow$ the price process $\pi_t\left(\frac{C}{B_t}\right)$ is on Q -martingale (otherwise we wouldn't get an arbitrage-free price) \leadsto writing $(C_{n,j})_j$ for the possible values of the undiscounted price process at time n , we get from the martingale property:

$$\mathbb{E}_Q\left(\frac{C_{n+1}}{B_{n+1}} \mid C_{n,j}\right) = \frac{C_{n,j}}{B_n} \Leftrightarrow \frac{B_n}{B_{n+1}} \cdot \mathbb{E}_Q(C_{n+1} \mid C_{n,j}) = C_{n,j}$$

$$= q_u C_{n+1,j+1} + q_d C_{n+1,j-1} + (1 - q_u - q_d) C_{n+1,j}$$

$$\Rightarrow \text{this yields the recursion formula: } C_{n,j} = \frac{B_n}{B_{n+1}} \cdot (q_u C_{n+1,j+1} + q_d C_{n+1,j-1} + q_m C_{n+1,j})$$



use an algorithm that computes everything (the $C_{n,j}$ values) only in one vector (not a matrix)