

Introduction to Mathematical Finance

Solution sheet 6

Solution 6.1

- (a) If $(N_t)_{t \in \{0, \dots, T\}}$ a numeraire, $\left(\frac{S_t}{B_t}\right)_t$ is a Q -EMM \iff $\left(\frac{S_t}{N_t}\right)_t$ is a Q^N -EMM

where

$$\left. \frac{dQ^N}{dQ} \right|_{\mathcal{F}_t} = \frac{N_T/N_t}{B_T/B_t}.$$

- (b) The forward Libor rate $(L(t, T, T + \delta))_t$ is a martingale under the *forward probability*.

$$\begin{aligned} L(t, T, T + \delta) &= \frac{1}{\delta} \left(\frac{1}{P(t, T, T + \delta)} - 1 \right) \\ &= \frac{1}{\delta} \left(\frac{P(t, T)}{P(t, T + \delta)} - 1 \right) \end{aligned}$$

By the numeraire change theorem, if we take as numeraire the zero-coupon bond $(N_t)_{t \in \{0, \dots, T\}} = (P(t, T + \delta))_{t \in \{0, \dots, T\}}$ then we have that

$$\begin{aligned} \left(\frac{P(t, T)}{B_t}\right)_t \text{ is a } Q\text{-EMM} &\iff \left(\frac{P(t, T)}{P(t, T + \delta)}\right)_t \text{ is a } Q^{T+\delta}\text{-EMM} \\ &\iff \left(\frac{1}{\delta} \left(\frac{P(t, T)}{P(t, T + \delta)} - 1\right)\right)_t \text{ is a } Q^{T+\delta}\text{-EMM} \\ &\iff (L(t, T, T + \delta))_t \text{ is a } Q^{T+\delta}\text{-EMM}. \end{aligned}$$

The arbitrage free price of a caplet with payoff $C^{caplet} = \delta(L(t, T, T + \delta) - K)_+$ is

$$\begin{aligned} \pi(C^{caplet}) &= E_t^Q \left(\frac{C^{caplet}}{B_{T+\delta}/B_t} \right) \\ &= B_t E_t^Q \left(\frac{\delta(L(T, T, T + \delta) - K)_+}{B_{T+\delta}} \right) \\ &= B_t E_t^{Q^N} \left(\left. \frac{dQ}{dQ^N} \right|_{\mathcal{F}_t} \frac{\delta(L(T, T, T + \delta) - K)_+}{B_{T+\delta}} \right) \\ &= B_t E_t^{Q^N} \left(\frac{B_{T+\delta}/B_t}{N_{T+\delta}/N_t} \frac{\delta(L(T, T, T + \delta) - K)_+}{B_{T+\delta}} \right) \\ &= N_t E_t^{Q^N} \left(\frac{\delta(L(T, T, T + \delta) - K)_+}{N_{T+\delta}} \right). \end{aligned}$$

The numeraire here is the *zero-coupon* $N_t = P(t, T + \delta)$.

$$\begin{aligned} \pi(C^{caplet}) &= P(t, T + \delta) E_t^{Q^{T+\delta}} \left(\frac{\delta(L(T, T, T + \delta) - K)_+}{P(T + \delta, T + \delta)} \right) \\ &= P(t, T + \delta) E_t^{Q^{T+\delta}} (\delta(L(T, T, T + \delta) - K)_+) \end{aligned}$$

(c) The swap rate $(S_{T_0, T_N}(t))_t$ is a martingale under the *forward probability*.

$$\begin{aligned} \left(\frac{P(t, T_i)}{B_t}\right)_t \text{ is a } Q\text{-EMM } (i = 0, N) &\iff \left(\frac{P(t, T_i)}{A(t)}\right)_t \text{ is a } Q^A\text{-EMM } (i = 0, N) \\ &\iff \left(\frac{P(t, T_0) - P(t, T_N)}{A(t)}\right)_t \text{ is a } Q^A\text{-EMM} \\ &\iff (S_{T_0, T_N}(t))_t \text{ is a } Q^A\text{-EMM} \end{aligned}$$

The arbitrage free price of a swaption with payoff $C^{swaption} = A(T_f)(S_{T_0, T_N}(T_f) - K)_+$ is

$$\begin{aligned} \pi(C^{swaption}) &= E_t^Q \left(\frac{C^{swaption}}{B_T/B_t} \right) \\ &= B_t E_t^Q \left(\frac{A(T_f)(S_{T_0, T_N}(T_f) - K)_+}{B_T} \right) \\ &= A(t) E_t^{Q^A} \left(\frac{A(T_f)(S_{T_0, T_N}(T_f) - K)_+}{A(T_f)} \right) \\ &= A(t) E_t^{Q^A} ((S_{T_0, T_N}(T_f) - K)_+) \end{aligned}$$

Solution 6.2

- (a) Book chapter 5.5.2
- (b) Book chapter 5.6.1
- (c) Book chapter 5.53
- (d) Book chapter 5.62

Solution 6.3

(a) Let us compute the characteristic function of Z_n :

$$\begin{aligned} \phi_{Z_n}(z) &= \mathbb{E} [e^{izZ_n}] \\ &= \mathbb{E} \left[e^{iz \sum_{i=1}^n X_i^n} \right] \\ &= \prod_{i=1}^n \mathbb{E} [e^{izX_i^n}] \\ &= \mathbb{E} [e^{izX_1^n}]^n \end{aligned}$$

where we get the third equality by independence of the X_i^n , and the fourth because they are identically distributed. A Taylor expansion gives :

$$\mathbb{E} [e^{izX_1^n}] = \left(1 + iz\mu_n - \frac{z^2\sigma^2 T}{2n} + o\left(\frac{1}{n}\right) \right).$$

Hence $\lim_{n \rightarrow \infty} \phi_{Y_n}(z) = \exp\left(iz\mu - \frac{\sigma^2 z^2 T}{2}\right)$, which is continuous at 0 and the characteristic function of a normal random variable with mean μ and variance $\sigma^2 T$.

- (b) The price at time 0 of an attainable claim is given by the expected value under an equivalent martingale measure of its discounted payoff :

$$\begin{aligned} V_0^{P,N} &= \frac{\mathbb{E}^* \left[\left(K - \tilde{S}_N^1 \right)^+ \right]}{(1+r)^N} \\ &= \frac{\mathbb{E}^* \left[\left(K - S_0^1 \prod_{i=1}^N Y_i \right)^+ \right]}{(1+r)^N} \\ &= \mathbb{E}^* \left[\left(\frac{K}{(1+r)^N} - S_0^1 \exp(Z_N) \right)^+ \right], \end{aligned}$$

where $Z_N = \sum_{i=1}^N \log \left(\frac{Y_i}{1+r} \right)$.

- (c) By assumption, $\log \left(\frac{Y_i}{1+r} \right)$ is valued in $\left\{ -\sigma \sqrt{\frac{T}{N}}, \sigma \sqrt{\frac{T}{N}} \right\}$ and the Y_i 's are i.i.d. under \mathbb{P}^* . Furthermore we have :

$$\begin{aligned} \mathbb{E}^* \left[\log \left(\frac{Y_i}{1+r} \right) \right] &= -p^* \sigma \sqrt{\frac{T}{N}} + (1-p^*) \sigma \sqrt{\frac{T}{N}} \\ &= (1-2p^*) \sigma \sqrt{\frac{T}{N}} \\ &= \left(1 - 2 \frac{u-r}{u-d} \right) \sigma \sqrt{\frac{T}{N}} \\ &= \left(1 - 2 \frac{e^{\sigma \sqrt{\frac{T}{N}}} \left(1 + \frac{RT}{N} \right) - 1 - \frac{RT}{N}}{e^{\sigma \sqrt{\frac{T}{N}}} \left(1 + \frac{RT}{N} \right) - e^{-\sigma \sqrt{\frac{T}{N}}} \left(1 + \frac{RT}{N} \right)} \right) \sigma \sqrt{\frac{T}{N}} \\ &= \frac{2 - e^{\sigma \sqrt{\frac{T}{N}}} - e^{-\sigma \sqrt{\frac{T}{N}}}}{e^{\sigma \sqrt{\frac{T}{N}}} - e^{-\sigma \sqrt{\frac{T}{N}}}} \sigma \sqrt{\frac{T}{N}} \\ &= \frac{2 - \left(1 + \sigma \sqrt{\frac{T}{N}} + \frac{\sigma^2 T}{2N} + o\left(\frac{1}{N}\right) \right) - \left(1 - \sigma \sqrt{\frac{T}{N}} + \frac{\sigma^2 T}{2N} + o\left(\frac{1}{N}\right) \right)}{\left(1 + \sigma \sqrt{\frac{T}{N}} + \frac{\sigma^2 T}{2N} + o\left(\frac{1}{N}\right) \right) - \left(1 - \sigma \sqrt{\frac{T}{N}} + \frac{\sigma^2 T}{2N} + o\left(\frac{1}{N}\right) \right)} \sigma \sqrt{\frac{T}{N}} \\ &= \frac{-\frac{\sigma^2 T}{N} + o\left(\frac{1}{N}\right)}{2\sigma \sqrt{\frac{T}{N}} + o\left(\frac{1}{N}\right)} \sigma \sqrt{\frac{T}{N}} \\ &= \frac{-\frac{\sigma^2 T}{N} + o\left(\frac{1}{N}\right)}{2 + o\left(\sqrt{\frac{1}{N}}\right)} \\ &= -\frac{\sigma^2 T}{2N} + o\left(\frac{1}{N}\right) \end{aligned}$$

and we can use the result proved in a) with $\mu = -\frac{\sigma^2 T}{2}$.

Consider the function $f(z) = (K e^{-RT} - S_0^1 e^z)^+$ and the difference :

$$\begin{aligned} \left| V_0^{P,N} - \mathbb{E}^* [f(Z_N)] \right| &= \left| \mathbb{E}^* \left[\left(\frac{K}{(1+r)^N} - S_0^1 \exp(Z_N) \right)^+ \right] - \mathbb{E}^* \left[(K e^{-RT} - S_0^1 \exp(Z_N))^+ \right] \right| \\ &\leq K \left| \left(1 + \frac{RT}{N} \right)^{-N} - e^{-RT} \right| \end{aligned}$$

Hence, we have $\lim_{N \rightarrow \infty} V_0^{P,N} = \lim_{N \rightarrow \infty} \mathbb{E}^* [f(Z_N)]$. And since f is bounded and continuous, and the sequence $(Z_N)_{N \in \mathbb{N}}$ converges in law to a normal random variable with mean $\mu = -\frac{\sigma^2 T}{2}$ and variance $\sigma^2 T$, we have :

$$\begin{aligned}
 \lim_{N \rightarrow \infty} \mathbb{E}^* [f(Z_N)] &= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} \left(K e^{-RT} - S_0 e^{\sigma\sqrt{T}y - \frac{\sigma^2 T}{2}} \right)^+ e^{-\frac{y^2}{2}} dy \\
 &= K e^{-RT} \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} \mathbb{1}_{\{K e^{-RT} \geq S_0 e^{\sigma\sqrt{T}y - \frac{\sigma^2 T}{2}}\}} e^{-\frac{y^2}{2}} dy \\
 &\quad - S_0 \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} \mathbb{1}_{\{K e^{-RT} \geq S_0 e^{\sigma\sqrt{T}y - \frac{\sigma^2 T}{2}}\}} e^{\sigma\sqrt{T}y - \frac{\sigma^2 T}{2}} e^{-\frac{y^2}{2}} dy \\
 &= K e^{-RT} \Phi(-d_2) - S_0 \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{-d_2} e^{-\frac{1}{2}(y - \sigma\sqrt{T})^2} dy \\
 &= K e^{-RT} \Phi(-d_2) - S_0 \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{-d_1} e^{-\frac{1}{2}y^2} dy \\
 &= K e^{-RT} \Phi(-d_2) - S_0 \Phi(-d_1)
 \end{aligned}$$

Which gives the result.

Solution 6.4 The solution will be given the next week.