

Probability Theory

Solution sheet 11

Solution 11.1 We define $X_0 := 0$, $X_n := \sum_{m=1}^n (1_{A_m} - P[A_m|\mathcal{F}_{m-1}])$, $n \geq 1$. Then X_n is an \mathcal{F}_n -martingale, since

$$E[X_{n+1} - X_n|\mathcal{F}_n] = E[1_{A_{n+1}} - P[A_{n+1}|\mathcal{F}_n]|\mathcal{F}_n] = 0.$$

Furthermore $|X_{n+1} - X_n| \leq 2$. We apply the result of Exercise 10.3 to obtain $P[C \cup D] = 1$. Note that:

- $\sum_{n \geq 1} 1_{A_n} = \infty \iff \sum_{n \geq 1} P[A_n|\mathcal{F}_{n-1}] = \infty$ on C .
- $\sum_{n \geq 1} 1_{A_n} = \infty$ and $\sum_{n \geq 1} P[A_n|\mathcal{F}_{n-1}] = \infty$ on D .

Using that $P[C \cup D] = 1$, we get that for an event N with $P[N] = 0$,

$$\left\{ \sum_{n \geq 1} 1_{A_n} = \infty \right\} \cap N^c = \left\{ \sum_{n \geq 1} P[A_n|\mathcal{F}_{n-1}] = \infty \right\} \cap N^c.$$

Finally, the claim follows since

$$\limsup_{n \rightarrow \infty} A_n = \left\{ \sum_{n \geq 1} 1_{A_n} = \infty \right\}.$$

Solution 11.2 Let $S_n = \sum_{m=1}^n X_m$ ($S_0 = 0$) be the simple random walk on \mathbb{Z}^d such that X_n , $n \geq 1$ are i.i.d. \mathbb{Z}^d valued random variable with $P[X_1 = e] = 1/(2d)$ for any $e \in \mathbb{R}^d$, $\|e\| = 1$ (i.e., $e = e_i$ or $e = -e_i$ for $e_i = (0, \dots, 0, 1, 0, \dots, 0)$ with 1 on the i th entry, $i = 1, \dots, d$). We can easily show that S_n is an \mathcal{F}_n martingale taking values in \mathbb{Z}^d and S_n^j , the j th coordinate of S_n , is also an \mathcal{F}_n real-valued martingale for each $j = 1, \dots, d$.

(a) Consider the martingale

$$M_n = \sum_{i=1}^d S_n^i,$$

with bounded increments

$$M_n - M_{n-1} = \sum_{i=1}^d (S_n^i - S_{n-1}^i) \in \{-1, 1\}.$$

By Exercise 10.3, it thus holds that

$$P[\{\lim M_n \in \mathbb{Z} \text{ exists}\} \cup \{\liminf M_n = -\infty \text{ and } \limsup M_n = \infty\}] = 1.$$

However, since $|M_n - M_{n-1}| = 1$, M_n cannot converge, from which it follows that $\limsup M_n = \infty$ P -a.s. and $\liminf M_n = -\infty$ P -a.s. This, in turn, implies that S_n must exit the finite set A , since $\max_{(a_1, \dots, a_d) \in A} \sum_{i=1}^d |a_i| < \infty$.

(b) If $x \notin A$, then $T_x = 0 \Rightarrow f(x + S_{T_x}) = f(x)$.

If $x \in A$, then

$$\begin{aligned} E[f(x + S_{T_x})] &= \sum_{\|e\|=1} E[f(x + S_{T_x}) \mathbf{1}_{S_1=e}] \\ &= \frac{1}{2d} \sum_{\|e\|=1} E[f(x + e + S_{T_{x+e}})]. \end{aligned}$$

To see the last equality, we note that if $x + e \notin A$, then in view of the definition of T_{x+e} (resp. T_x) we have $T_{x+e} = 0$ (resp. $T_x = 1$) and in this case it holds that $E[f(x + S_{T_x}) \mathbf{1}_{S_1=e}] = f(x + e)P[S_1 = e] = \frac{f(x+e)}{2d} = E[f(x + e + S_{T_{x+e}})]$. On the other hand, if $x + e \in A$, then we must have $T_x \geq 2$ and $T_{x+e} \geq 1$. In this case we have

$$\begin{aligned} E[f(x + S_{T_x}) \mathbf{1}_{S_1=e}] &= \sum_{n \geq 2} E[f(x + e + S_n - S_1) \mathbf{1}_{\{S_1=e\} \cap \{T_x=n\}}] \\ &= \sum_{n \geq 2} E[f(x + e + S_n - S_1) \mathbf{1}_{\{S_1=e\} \cap B_n}], \end{aligned}$$

where $B_n = \{\forall 2 \leq m < n, x + e + (S_m - S_1) \in A; x + e + S_n - S_1 \notin A\}$. But since $S_m - S_1 = \sum_{j=2}^m X_j$ is independent of $S_1 = X_1$ for all $m \geq 2$, the event B_n is independent of $\{S_1 = e\}$. Moreover, it is easy to see that $x + e + S_n - S_1 = x + e + \sum_{j=2}^n X_j$, $m \geq 1$ is the simple random walk started from $x + e$, which implies that $B_n = \{T_{x+e} = n - 1\}$ for all $n \geq 2$ and $x + e + S_n - S_1 = x + e + S_{T_{x+e}}$ on B_n . Hence, we can deduce that

$$\begin{aligned} \sum_{n \geq 2} E[f(x + e + S_n - S_1) \mathbf{1}_{\{S_1=e\} \cap B_n}] &= \sum_{n \geq 2} E[f(x + e + S_n - S_1) \mathbf{1}_{B_n}] P[S_1 = e] \\ &= \frac{1}{2d} E[f(x + e + S_{T_{x+e}}) \mathbf{1}_{\{T_{x+e} \geq 1\}}]. \end{aligned}$$

Now we complete our proof by writing $y = x + e$ for some e with $\|e\| = 1$.

(c) In case $x \notin A$ the statement is trivial.

If $x \in A$, then $S_{T_x \wedge 1} = S_1$, whence it follows that

$$E[f(x + S_{T_x}) \mathbf{1}_{S_1=e}] = \frac{1}{2d} E[f(x + e + S_{T_{x+e}})] = \frac{1}{2d} g(x + e)$$

and

$$E[g(x + S_1) \mathbf{1}_{S_1=e}] = \frac{1}{2d} E[g(x + e)] = \frac{1}{2d} g(x + e).$$

This concludes the proof.

Solution 11.3

(a) Since $(M_n)_{n \geq 0}$ is adapted to the filtration $\mathcal{F}_n = \sigma(Y_0, \dots, Y_n)_{n \in \mathbb{N}}$ and, for $n \in \mathbb{N}$, $E[|M_n|] = E[Y_1] \cdots E[Y_n] = 1$, it holds that $M_n \in L^1$. Moreover,

$$\begin{aligned} E[M_{n+1} | \mathcal{F}_n] &= E[M_n Y_{n+1} | \mathcal{F}_n] \\ (M_n \mathcal{F}_n\text{-meas.}) &= M_n E[Y_{n+1} | \mathcal{F}_n] \\ (Y_{n+1}, \mathcal{F}_n \text{ indep.}) &= M_n E[Y_{n+1}] = M_n. \end{aligned}$$

Hence, $(M_n)_{n \geq 0}$ is a \mathcal{F}_n -martingale with $\sup_{n \in \mathbb{N}} E[M_n] = 1 < \infty$ and, by the martingale convergence theorem ((3.4.23) in the lecture notes), M_n converges to an integrable random variable M_∞ P -a.s.

(b) From the Cauchy-Schwarz (or Hölder) inequality it follows that,

$$a_n = E[\sqrt{Y_n}] \leq E[Y_n]^{1/2} E[1]^{1/2} = 1.$$

Thus, since $Y_n \geq 0$ P -a.s. and therefore $\sqrt{Y_n} \geq 0$ P -a.s. and $E[Y_n] = 1$, it must hold that $a_n > 0$.

(c) We define $\hat{Y}_n = \sqrt{Y_n}/a_n$ and $\hat{M}_n = \hat{Y}_1 \hat{Y}_2 \dots \hat{Y}_n$, $\hat{M}_0 = 1$. Then, from (a), the process $(\hat{M}_n)_{n \in \mathbb{N}}$ is also a non-negative martingale, converging to an integrable random variable \hat{M}_∞ . Furthermore $M_n \leq \hat{M}_n^2$ for $n \in \mathbb{N}$.

Suppose that $\prod_n a_n > 0$. Then we get,

$$E[\hat{M}_n^2] = (a_1 a_2 \dots a_n)^{-2} \leq \left(\prod_n a_n \right)^{-2} < \infty.$$

Therefore, $\hat{M}_n \in L^2$, and by Doob's inequality (see Corollary 3.34),

$$E \left[\sup_{n \in \mathbb{N}} M_n \right] \leq E \left[\sup_{n \in \mathbb{N}} \hat{M}_n^2 \right] \leq 4 \sup_{n \in \mathbb{N}} E[\hat{M}_n^2] < \infty.$$

From Lebesgue convergence theorem ($|M_n - M_\infty| \leq 2 \sup_{n \in \mathbb{N}} M_n \in L^1$) we obtain convergence also in L^1 .

(d) Suppose that $\prod_n a_n = 0$. Then, from the P -a.s. convergence of $\hat{M}_n = \sqrt{M_n}/(\prod_{i=1}^n a_i)$ to an a.s. finite random variable $\hat{M}_\infty \in L^1$, it follows that $M_\infty = 0$ a.s.