

Probability and Statistics

Exercise sheet 8

Exercise 8.1

- (a) Let $X \sim U([0, 1])$. Compute $E(X^n)$, $E(X^{\frac{1}{n}})$ ($n \geq 1$), and $\Psi_X(t) = E[e^{tX}]$ whenever it is defined.
- (b) Let $X \sim \text{Beta}(\alpha, \beta)$, $\alpha > 0$ and $\beta > 0$. Compute $E(X)$ and $\text{var}(X)$.
Hint: Use the “trick” that any density function f has to integrate to 1.
- (c) Let $X \sim \text{Exp}(\lambda)$, for $\lambda > 0$. Compute the cdf of X and $E(X^n)$ for $n \geq 1$.
Remark: Watch out for the parametrisation - different sources may use different parametrisations. In the lectures we consider the density function of an $\text{Exp}(\lambda)$ -distributed random variable to be $f(x) = \lambda e^{-\lambda x}$ for $x \geq 0$.

Exercise 8.2 (Waiting time.)

An auto towing company services a 50 mile stretch of a highway. The company is located 20 miles from one end of the stretch. Breakdowns occur uniformly along the highway and the towing trucks travel at a constant speed of 50mph. Find the mean and variance of the time elapsed between the instant the company is called and a towing truck arrives.

Exercise 8.3 (Uniforms, uniforms...)

- (a) Consider a random variable $X \sim U\left(\left[-\frac{\pi}{2}, \frac{\pi}{2}\right]\right)$. Find $E(\sin(X))$ and $\text{var}(\sin(X))$.
- (b) The lengths of the sides of a triangle are X , $2X$ and $2.5X$ with $X \sim U([0, \alpha])$ for some $\alpha > 0$.

- Find the mean and variance of its area.

Hint: Recall that if

$$s = \frac{a + b + c}{2}$$

with a, b, c the lengths, then the area of the triangle is

$$|\Delta| = \sqrt{s(s-a)(s-b)(s-c)}$$

(Heron's formula).

- How should we choose α so that the mean area is ≥ 1 ?

- (c) Take X_1, \dots, X_n to be $\stackrel{\text{iid}}{\sim} U([0, 1])$. Let $M_n = \max(X_1, \dots, X_n)$. Find the cdf and pdf of M_n . Can you recognise this distribution? What are $E(M_n)$ and $\text{var}(M_n)$?

Exercise 8.4 (Quantile transformation.) Recall that for a given cdf F , the quantile t_α of order $\alpha \in (0, 1)$ is defined as

$$t_\alpha = \inf\{t : F(t) \geq \alpha\} =: F^{-1}(\alpha).$$

F^{-1} denotes the generalised inverse of F . When the latter is bijective (at least in the neighbourhood of t_α), then F^{-1} is the inverse of F in the classical sense.

(a) Consider $U \sim U([0, 1])$. Show that $1 - U \sim U([0, 1])$.

Hint: Compute, for example, the cdf or the pdf of $1 - U$.

(b) Consider

$$X := -\frac{1}{\lambda} \log(U)$$

(for $\omega : U(\omega) = 0$, take $X(\omega) = 0$, say).

Find the cdf of X . Can you recognise this distribution?

(c) Now, consider the following problem: take a cdf F which is bijective when viewed as a map $F : (a, b) \rightarrow (0, 1)$, for some $-\infty \leq a < b \leq +\infty$.

Define $X = F^{-1}(U)$, with $U \sim U([0, 1])$.

- Compute the cdf of X .

Hint: You may need to consider the cases $a = -\infty$, $a > -\infty$, $b = +\infty$, $b < +\infty$.

- Compute the pdf of X , assuming that F is C^1 on (a, b) with $F'(x) > 0 \forall x \in (a, b)$.
- Can you make the link to (b)?

(d) Suppose you are given a numerical algorithm which enables you to generate a random number from $[0, 1]$. You would like to generate a random number X which follows the Cauchy distribution, i.e.

$$f_X(x) = \frac{1}{\pi(1+x^2)}, \quad x \in \mathbb{R}.$$

Can you propose a way to do that, based on your previous findings?

Exercise 8.5 (Optional.) Let X_1, \dots, X_n (for some $n \geq 2$) be random variables defined on the same probability space (Ω, \mathcal{A}, P) . A necessary and sufficient condition for (mutual) independence of X_1, \dots, X_n is that

$$P(X_1 \leq x_1, \dots, X_n \leq x_n) = \prod_{i=1}^n P(X_i \leq x_i) \quad (1)$$

for any $(x_1, \dots, x_n) \in \mathbb{R}^n$.

The goal of this exercise is to show that when X_1, \dots, X_n are discrete, (1) is equivalent to

$$P(X_{i_1} = x_1, \dots, X_{i_m} = x_m) = \prod_{j=1}^m P(X_{i_j} = x_j) \quad (2)$$

for any $1 \leq m \leq n$ and $i_1 < \dots < i_m$, $(x_1, \dots, x_m) \in \mathbb{R}^m$.

(a) Take $m = n$ (in which case $i_1 = 1, i_2 = 2, \dots, i_m = n$).

- Focus on X_1 and show that (2) implies that for any $(x_1, \dots, x_n) \in \mathbb{R}^n$,

$$P(X_1 \leq x_1, X_2 = x_2, \dots, X_n = x_n) = P(X_1 \leq x_1)P(X_2 = x_2) \dots P(X_n = x_n).$$

- Repeating this argument inductively, show that (1) holds.

(b) Now, we want to show that (1) implies (2). Fix $\{i_1, \dots, i_m\} = \mathcal{J} \subset \{1, \dots, n\}$, with strict inclusion.

- Show that (1) implies that

$$P(X_{i_1} \leq x_1, \dots, X_{i_m} \leq x_m) = \prod_{j=1}^m P(X_{i_j} \leq x_j).$$

Hint: You may want to take limits in (1) as $x_i \rightarrow +\infty$ for $i \notin \mathcal{J}$.

- Focus on i_1 (hold the other events depending on i_2, \dots, i_m fixed). Show that we have

$$P(X_{i_1} < x_1, X_{i_2} \leq x_2, \dots, X_{i_m} \leq x_m) = P(X_{i_1} < x_1)P(X_{i_2} \leq x_2) \dots P(X_{i_m} \leq x_m)$$

for any $(x_1, \dots, x_m) \in \mathbb{R}^m$.

Hint: Consider $A_k := \{X_{i_1} \leq x_1 - \frac{1}{k}, X_{i_2} \leq x_2, \dots, X_{i_m} \leq x_m\}$ and use the monotone convergence theorem.

- Conclude that

$$P(X_{i_1} = x_1, X_{i_2} \leq x_2, \dots, X_{i_m} \leq x_m) = P(X_{i_1} = x_1)P(X_{i_2} \leq x_2) \dots P(X_{i_m} \leq x_m).$$

- Repeat your argument inductively to conclude.