## Probability and Statistics

## Exercise sheet 1

Exercise 1.1 Each of 3 people toss a fair coin. What is the probability of someone being the "odd man out"? This means that 2 of the players obtain the same outcome, while the third gets a different one.

## Solution 1.1

$$
\begin{aligned}
p & :=P(\text { one of the players is the "odd man out" }) \\
& =1-P(\text { all } 3 \text { players get the same outcome }) \\
& =1-P(\mathrm{HHH} \text { or } \mathrm{TTT}) \\
& =1-P(\mathrm{HHH})-P(\mathrm{TTT}) \\
& =1-\frac{1}{2^{3}} \times 2 \\
& =1-\frac{1}{4} \\
& =\frac{3}{4} .
\end{aligned}
$$

Here, $P(\mathrm{HHH})=P(\mathrm{TTT})=\frac{1}{\operatorname{card}(\Omega)}$, where $\operatorname{card}(\Omega)=2 \times 2 \times 2=2^{3}=8$, the number of all possible outcomes for the 3 tosses.

Exercise 1.2 An urn contains 5 red, 5 black and 5 white balls. 3 balls are chosen without replacement at random. What is the probability that they are of exactly two different colors?

Solution 1.2 There are at least two methods to solve this question.
Method 1: Let $p:=P(A)$ be the probability of the event

$$
A:=\{\text { the } 3 \text { balls are of exactly } 2 \text { different colors }\} .
$$

Then

$$
\begin{aligned}
p & =1-P(\text { the } 3 \text { balls have the same color or all have different colors) } \\
& =1-P(S)-P(D)
\end{aligned}
$$

where $S=\{$ same color $\}$ and $D=\{$ all different colors $\}$. Using the notation $\binom{n}{k}\left(=C_{n}^{k}\right)$ for binomial coefficients, we compute

$$
\begin{aligned}
P(S) & =P(\text { all red })+P(\text { all black })+P(\text { all white }) \\
& =\frac{\binom{5}{3}}{\binom{15}{3}}+\frac{\binom{5}{3}}{\binom{15}{3}}+\frac{\binom{5}{3}}{\binom{15}{3}} \\
& =\frac{3 \times \frac{5 \times 4 \times 3!}{3!2}}{\frac{15 \times 14 \times 13}{3 \times 2}} \\
& =\frac{3 \times 6 \times 10}{15 \times 14 \times 13} \\
& =\frac{180}{2730} \\
& =\frac{6}{91} \\
P(D) & =P(\{\text { white, red, black }\}) \\
& =\frac{\binom{5}{1} \times\binom{ 5}{1} \times\binom{ 5}{1}}{\binom{15}{3}} \\
& =\frac{125 \times 6}{15 \times 14 \times 13} \\
& =\frac{750}{2730} \\
& =\frac{25}{91}
\end{aligned}
$$

In conclusion,

$$
p=1-\frac{6+25}{91}=\frac{60}{91} \approx 0.66
$$

Method 2: More directly, we want to compute the probability $p$ of the event

$$
\begin{aligned}
A= & \{\text { red, red, white }\},\{\text { red, red, black }\},\{\text { white }, \text { white }, \text { red }\} \\
& \{\text { white }, \text { white, black }\},\{\text { black, black, white }\},\{\text { black, black, red }\}\}
\end{aligned}
$$

Since $\operatorname{card}(A)=6 \times\binom{ 5}{2} \times\binom{ 5}{1}$,

$$
p=\frac{6 \times\binom{ 5}{2} \times\binom{ 5}{1}}{\binom{15}{3}}=\frac{6 \times \frac{5 \times 4}{2} \times 5}{\frac{15 \times 14 \times 13}{6}}=\frac{60}{91} \approx 0.66
$$

Exercise 1.33 families with three members each organize themselves randomly into a line, to take a random picture. What is the probability that members of each family appear together in the picture? (i.e., not separated by members from other families.)

Solution 1.3 Write

$$
\begin{aligned}
& F_{1}=\left\{M_{1}, M_{2}, M_{3}\right\} \\
& F_{2}=\left\{N_{1}, N_{2}, N_{3}\right\} \\
& F_{3}=\left\{J_{1}, J_{2}, J_{3}\right\}
\end{aligned}
$$

for the members of families $1,2,3$ respectively.

To count the possible positionings where the members of each family are together, first note that there are 3 ! possibilities for the positions of the families $\left(F_{1} F_{2} F_{3}, F_{1} F_{3} F_{2}, \ldots, F_{3} F_{2} F_{1}\right)$, and that for each of these possibilities, the members within each family can be organized in (3! $)^{3}$ ways.

Hence, the cardinality of the event of interest is

$$
3!\times(3!)^{3}=(3!)^{4}=6^{4}
$$

Furthermore, the number of all possibilities for a random picture is 9 !, and hence

$$
p=\frac{6^{4}}{9!} \approx 0.0036
$$

Exercise 1.4 In a building with 6 floors (plus the ground floor), an elevator starts with 4 people at the ground floor. What is the probability that these people get off at exactly 2 floors?

Solution 1.4 If $P_{i}(i=1, \ldots, 4)$ are the 4 people, each has 6 possibilities of where to get off. Thus, $\operatorname{card}(\Omega)=6^{4}$. Write

$$
A=\left\{P_{1}, P_{2}, P_{3}, P_{4} \text { get off at exactly } 2 \text { floors }\right\}
$$

To compute $\operatorname{card}(A)$, there are $\binom{6}{2}$ ways of selecting those two floors. Then, given the 2 floors, there are $\binom{4}{1}+\binom{4}{2}+\binom{4}{3}$ ways of distributing people between those floors. Therefore,

$$
\begin{aligned}
P(A) & =\frac{\left(\binom{4}{1}+\binom{4}{2}+\binom{4}{3}\right)\binom{6}{2}}{6^{4}} \\
& =\frac{(4+6+4) \frac{6 \times 5}{2}}{1296} \\
& =\frac{14 \times 15}{1296} \\
& =\frac{35}{216} \\
& \approx 0.162
\end{aligned}
$$

Exercise 1.5 Consider an arbitrary sample space $\Omega$ and probability measure $P$. Using induction, prove that for any events $A_{1}, \ldots, A_{n} \subseteq \Omega$,

$$
P\left(\bigcup_{i=1}^{n} A_{i}\right)=\sum_{k=1}^{n}(-1)^{k+1} \sum_{1 \leq i_{1}<\ldots<i_{k} \leq n} P\left(\bigcap_{j=1}^{k} A_{i_{j}}\right)
$$

Solution 1.5 For $n=1$ this is trivial, and for $n=2$, we have

$$
A_{1} \cup A_{2}=A_{1} \cup\left(A_{2} \backslash A_{1} \cap A_{2}\right)
$$

where $A_{1}$ and $A_{2} \backslash A_{1} \cap A_{2}$ are disjoint.
Thus $P\left(A_{1} \cup A_{2}\right)=P\left(A_{1}\right)+P\left(A_{2} \backslash A_{1} \cap A_{2}\right)$. Also, $P\left(A_{1} \cap A_{2}\right)+P\left(A_{2} \backslash A_{1} \cap A_{2}\right)=P\left(A_{2}\right)$. Thus,

$$
P\left(A_{1} \cup A_{2}\right)=P\left(A_{1}\right)+P\left(A_{2}\right)-P\left(A_{1} \cap A_{2}\right)
$$

which clearly matches the claimed formula for $n=2$.

For the inductive step, suppose the formula holds for $n$ and let us show that it continues to hold for $n+1$.

$$
\begin{aligned}
P\left(\bigcup_{i=1}^{n+1} A_{i}\right) & =P\left(A_{1} \cup \ldots \cup A_{n+1}\right) \\
& =P\left(A_{1} \cup \ldots \cup A_{n}\right)+P\left(A_{n+1}\right)-P\left(\left(A_{1} \cup \ldots \cup A_{n}\right) \cap A_{n+1}\right) \\
\text { (induction) } & =\sum_{k=1}^{n}(-1)^{k+1} \sum_{1 \leq i_{1}<\ldots<i_{k} \leq n} P\left(\bigcap_{j=1}^{k} A_{i_{j}}\right)+P\left(A_{n+1}\right)-\sum_{k=1}^{n}(-1)^{k+1} \sum_{1 \leq i_{1}<\ldots<i_{k} \leq n} P\left(\bigcap_{j=1}^{k} A_{i_{j}} \cap A_{n+1}\right) \\
& =\sum_{i=1}^{n+1} P\left(A_{i}\right)+\sum_{k=2}^{n}(-1)^{k+1} \sum_{1 \leq i_{1}<\ldots<i_{k} \leq n} P\left(\bigcap_{j=1}^{k} A_{i_{j}}\right)-\sum_{k^{\prime}=2}^{n+1}(-1)^{k^{\prime}} \sum_{1 \leq i_{1}<\ldots<i_{k^{\prime}}=n+1} P\left(\bigcap_{j=1}^{k^{\prime}} A_{i_{j}}\right) .
\end{aligned}
$$

In the last step, we separated the first term from the first sum and relabelled $k^{\prime}=k+1$. Since the name of the dummy variable is irrelevant, we can rewrite the last sum:

$$
\begin{aligned}
-\sum_{k^{\prime}=2}^{n+1}(-1)^{k^{\prime}} \sum_{1 \leq i_{1}<\ldots<i_{k^{\prime}}=n+1} P\left(\bigcap_{j=1}^{k^{\prime}} A_{i_{j}}\right) & =\sum_{k^{\prime}=2}^{n+1}(-1)^{k^{\prime}+1} \sum_{1 \leq i_{1}<\ldots<i_{k^{\prime}}=n+1} P\left(\bigcap_{j=1}^{k^{\prime}} A_{i_{j}}\right) \\
& =\sum_{k=2}^{n+1}(-1)^{k+1} \sum_{1 \leq i_{1}<\ldots<i_{k}=n+1} P\left(\bigcap_{j=1}^{k} A_{i_{j}}\right)
\end{aligned}
$$

To conclude,

$$
\begin{aligned}
P\left(\bigcup_{i=1}^{n+1} A_{i}\right) & =\sum_{i=1}^{n+1} P\left(A_{i}\right)+\sum_{k=2}^{n}(-1)^{k+1} \sum_{1 \leq i_{1}<\ldots<i_{k} \leq n} P\left(\bigcap_{j=1}^{k} A_{i_{j}}\right)+\sum_{k=2}^{n+1}(-1)^{k+1} \sum_{1 \leq i_{1}<\ldots<i_{k}=n+1} P\left(\bigcap_{j=1}^{k} A_{i_{j}}\right) \\
& =\sum_{i=1}^{n+1} P\left(A_{i}\right)+\sum_{k=2}^{n+1}(-1)^{k+1} \sum_{1 \leq i_{1}<\ldots<i_{k} \leq n+1} P\left(\bigcap_{j=1}^{k} A_{i_{j}}\right) \\
& =\sum_{k=1}^{n+1}(-1)^{k+1} \sum_{1 \leq i_{1}<\ldots<i_{k} \leq n+1} P\left(\bigcap_{j=1}^{k} A_{i_{j}}\right)
\end{aligned}
$$

Thus the formula holds by induction. This is a probabilistic version of the "inclusion-exclusion" principle.

