

# Non-Life Insurance: Mathematics and Statistics

## Exercise sheet 1

### Exercise 1.1 Discrete Distribution

Suppose that  $N$  follows a geometric distribution with parameter  $p \in (0, 1)$ , i.e.

$$\mathbb{P}[N = k] = \begin{cases} (1-p)^{k-1}p, & \text{if } k \in \mathbb{N}_{>0}, \\ 0, & \text{else.} \end{cases}$$

- Show that the geometric distribution indeed defines a probability distribution on  $\mathbb{R}$ .
- Let  $n \in \mathbb{N}_{>0}$ . Calculate  $\mathbb{P}[N \geq n]$ .
- Calculate  $\mathbb{E}[N]$ .
- Let  $r < -\log(1-p)$ . Calculate the moment generating function  $M_N(r) = \mathbb{E}[\exp\{rN\}]$  of  $N$ .
- Calculate  $\frac{d}{dr}M_N(r)|_{r=0}$ . What do you observe?

### Exercise 1.2 Absolutely Continuous Distribution

Suppose that  $Y$  follows an exponential distribution with parameter  $\lambda > 0$ , i.e. the density  $f_Y$  of  $Y$  is given by

$$f_Y(x) = \begin{cases} \lambda \exp\{-\lambda x\}, & \text{if } x \geq 0, \\ 0, & \text{else.} \end{cases}$$

- Show that the exponential distribution indeed defines a probability distribution on  $\mathbb{R}$ .
- Let  $0 < y_1 < y_2$ . Calculate  $\mathbb{P}[y_1 \leq Y \leq y_2]$ .
- Calculate  $\mathbb{E}[Y]$  and  $\text{Var}(Y)$ .
- Let  $r < \lambda$ . Calculate the cumulant generating function  $\log M_Y(r) = \log \mathbb{E}[\exp\{rY\}]$  of  $Y$ .
- Calculate  $\frac{d^2}{dr^2} \log M_Y(r)|_{r=0}$ . What do you observe?

### Exercise 1.3 Gaussian Distribution

For a random variable  $X$  we write  $X \sim \mathcal{N}(\mu, \sigma^2)$  if  $X$  follows a Gaussian distribution with mean  $\mu \in \mathbb{R}$  and variance  $\sigma^2 > 0$ . The density  $f_X$  of  $X \sim \mathcal{N}(\mu, \sigma^2)$  is given by

$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{1}{2} \frac{(x-\mu)^2}{\sigma^2}\right\}, \quad \text{for all } x \in \mathbb{R}.$$

- Show that the moment generating function  $M_X$  of  $X \sim \mathcal{N}(\mu, \sigma^2)$  is given by

$$M_X(r) = \exp\left\{r\mu + \frac{r^2\sigma^2}{2}\right\}, \quad \text{for all } r \in \mathbb{R}.$$

(b) Let  $X \sim \mathcal{N}(\mu, \sigma^2)$  and  $a, b \in \mathbb{R}$ . Show that

$$a + bX \sim \mathcal{N}(a + b\mu, b^2\sigma^2).$$

(c) Let  $X_1, \dots, X_n$  be independent with  $X_i \sim \mathcal{N}(\mu_i, \sigma_i^2)$  for all  $i \in \{1, \dots, n\}$ . Show that

$$\sum_{i=1}^n X_i \sim \mathcal{N}\left(\sum_{i=1}^n \mu_i, \sum_{i=1}^n \sigma_i^2\right).$$

#### Exercise 1.4 $\chi^2$ -Distribution

For all  $k \in \mathbb{N}_{>0}$  we assume that  $X_k$  has a  $\chi^2$ -distribution with  $k$  degrees of freedom, i.e.  $X_k$  has density

$$f_{X_k}(x) = \begin{cases} \frac{1}{2^{k/2}\Gamma(k/2)} x^{k/2-1} \exp\{-x/2\}, & \text{if } x \geq 0, \\ 0, & \text{else.} \end{cases}$$

(a) Let  $M_{X_k}$  be the moment generating function of  $X_k$ . Show that

$$M_{X_k}(r) = \frac{1}{(1-2r)^{k/2}}, \quad \text{for } r < 1/2.$$

(b) Let  $Z \sim \mathcal{N}(0, 1)$ . Show that  $Z^2 \stackrel{(d)}{=} X_1$ .

(c) Let  $Z_1, \dots, Z_k \stackrel{\text{i.i.d.}}{\sim} \mathcal{N}(0, 1)$ . Show that  $\sum_{i=1}^k Z_i^2 \stackrel{(d)}{=} X_k$ .