

# Introduction to Mathematical Finance

## Solution sheet 11

**Solution 11.1** The dynamic programming principle (Proposition III.2.5) indicates that we should solve

$$\sigma_k(v_k) = \text{ess sup}_{\pi_k, \gamma_k} \left( U_c(\gamma_k v_k) + E \left[ \sigma_{k+1} (v_k(1 + \pi_k \cdot (R_{k+1} - 1) - \gamma_k)) \mid \mathcal{F}_k \right] \right)$$

by backwards induction to get candidate optimal solution.

- First note that  $\sigma_T(v_T) = U_w(v_T) = \log(v_T)$
- Consider  $k = T - 1$ :

$$\begin{aligned} \sigma_{T-1}(v_{T-1}) &= \text{ess sup}_{\pi_{T-1}, \gamma_{T-1}} \left( \log(\gamma_{T-1} v_{T-1}) + E \left[ \sigma_T (v_{T-1}(1 + \pi_{T-1} \cdot (R_T - 1) - \gamma_{T-1})) \mid \mathcal{F}_{T-1} \right] \right) \\ &= \text{ess sup}_{\pi_{T-1}, \gamma_{T-1}} \left( \log(\gamma_{T-1}) + \log(v_{T-1}) + \right. \\ &\quad \left. E \left[ \log(v_{T-1}) + \log(1 + \pi_{T-1} \cdot (R_T - 1) - \gamma_{T-1}) \mid \mathcal{F}_{T-1} \right] \right) \\ &= 2 \log(v_{T-1}) + \text{ess sup}_{\pi_{T-1}, \gamma_{T-1}} \left( \log(\gamma_{T-1}) + E \left[ \log((1 + \pi_{T-1} \cdot (R_T - 1) - \gamma_{T-1}) \mid \mathcal{F}_{T-1}) \right] \right) \\ &= 2 \log(v_{T-1}) + \\ &\quad \text{ess sup}_{\pi_{T-1}, \gamma_{T-1}} \left( \log(\gamma_{T-1}) + E \left[ \log((1 + \pi \cdot (R_T - 1) - \gamma)) \mid_{\pi=\pi_{T-1}, \gamma=\gamma_{T-1}} \right] \right) \\ &= 2 \log(v_{T-1}) + C_{T-1} \end{aligned}$$

where

$$C_{T-1} := \sup_{\pi, \gamma} \left( \log(\gamma) + E \left[ \log(1 + \pi \cdot (R - 1) - \gamma) \right] \right)$$

In the fourth equality above, we have used that  $R_T$  is independent of  $\mathcal{F}_{T-1}$  and  $\pi_{T-1}$  and  $\gamma_{T-1}$  are  $\mathcal{F}_{T-1}$  measurable. The random variable  $R$  appearing in the definition of  $C_{T-1}$  is distributed identically as the i.i.d random variables  $R_k$ .

- For a general  $0 \leq k \leq T - 1$ , if  $\sigma_{k+1}(v_{k+1}) = (T - (k + 1) + 1) \log(v_{k+1}) + \sum_{j=k+1}^{T-1} C_j$  for some constants  $C_j$ , then:

$$\begin{aligned}
\sigma_k(v_k) &= \text{ess sup}_{\pi_k, \gamma_k} \left( \log(\gamma_k v_k) + E \left[ \sigma_{k+1} (v_k (1 + \pi_k \cdot (R_{k+1} - 1) - \gamma_k)) \mid \mathcal{F}_k \right] \right) \\
&= \text{ess sup}_{\pi_k, \gamma_k} \left( \log(\gamma_k) + \log(v_k) + \right. \\
&\quad \left. E \left[ (T - (k + 1) + 1) \{ \log(v_k) + \log(1 + \pi_k \cdot (R_{k+1} - 1) - \gamma_k) \} + \sum_{j=k+1}^{T-1} C_j \mid \mathcal{F}_k \right] \right) \\
&= (T - k + 1) \log(v_k) + \sum_{j=k+1}^{T-1} C_j \\
&\quad + \text{ess sup}_{\pi_k, \gamma_k} \left( \log(\gamma_k) + (T - k) E \left[ \log((1 + \pi_k \cdot (R_{k+1} - 1) - \gamma_k)) \mid \mathcal{F}_k \right] \right) \\
&= (T - k + 1) \log(v_k) + \sum_{j=k+1}^{T-1} C_j \\
&\quad \text{ess sup}_{\pi_k, \gamma_k} \left( \log(\gamma_k) + (T - k) E \left[ \log((1 + \pi \cdot (R_{k+1} - 1) - \gamma)) \right]_{\pi=\pi_k, \gamma=\gamma_k} \right) \\
&= (T - k + 1) \log(v_k) + \sum_{j=k}^{T-1} C_j
\end{aligned}$$

where

$$C_k := \sup_{\pi, \gamma} \left( \log(\gamma) + (T - k) E \left[ \log(1 + \pi \cdot (R - 1) - \gamma) \right] \right)$$

This proves by induction that for all  $0 \leq k \leq T - 1$ ,

$$\sigma_k(v_k) = (T - k + 1) \log(v_k) + \sum_{j=k}^{T-1} C_j$$

Assume the constants  $C_k$  are finite and that there exist maximisers  $(\pi_k^*, \gamma_k^*)$  (constrained to satisfy  $\gamma_k^* \geq 0$  and  $1^T \pi_k^* = 1$  for all  $k$ ). We then expect by the dynamic programming principle  $(\pi_k^*, \gamma_k^*)$  to be optimal for the original static utility optimization problem (??). Indeed the dynamic programming principle (Proposition III.2.5) tells us that for an optimal solution of (??),  $\sigma_k(v_k)$  computed above is equal to the optimal remaining utility at time  $k$ . In particular,

$$E[\sigma_0(v_0)] = \mathcal{U}^*$$

is the optimal value of (??). To find the optimal strategy achieving this value  $\mathcal{U}^*$ , we first calculate backwards  $\sigma_k(v_k)$  and the corresponding optimizers  $\pi_k^*(v_k)$  and  $\gamma_k^*(v_k)$ . In general this backward step gives the optimal strategy *as a function of the current wealth*  $v_k$ . To get the optimal strategy we then need to iterate forward starting from the initial wealth  $v_0$ :  $W_0 = v_0$  and if  $W_k$  is known, set

$$\begin{aligned}
\pi_k^* &= \pi_k^*(W_k) \\
\gamma_k^* &= \gamma_k^*(W_k) \\
W_{k+1} &= W_k (1 + \pi_k^* \cdot (R_{k+1} - 1) - \gamma_k^*)
\end{aligned}$$

However, calculations simplify considerably in our setup, since the optimal  $\gamma_k^*$  and  $\pi_k^*$  do NOT depend on the current wealth  $v_k$  (see Remark below).

### Remarks

1. Note that the dynamic programming principle only gives a *candidate* solution to (??). To confirm the optimality, one should check the sufficient condition of the Martingale Optimality Principle, i.e. show that  $J(\vartheta^*, c^*)$  is a true martingale (where recall that  $c_k^* = \gamma_k^* W_k^{v_0, \pi^*, \gamma^*}$  and  $\vartheta_{k+1}^{*,l} = \frac{\pi_k^{*,l} W_k^{v_0, \pi^*, \gamma^*}}{X_k^l}$ ).
2. The (candidate) optimal proportions  $\gamma_k^*, \pi_k^*$  are deterministic. But there is still trading and stochastic consumption involved because

$$c_k^* = \gamma_k^* W_k^{v_0, \pi^*, \gamma^*}$$

and

$$\vartheta_{k+1}^{*,l} = \frac{\pi_k^{*,l} W_k^{v_0, \pi^*, \gamma^*}}{X_k^l}$$

still depend on the randomly evolving wealth and asset values.

### Solution 11.2

- (a) First note by induction that  $X_n$  is  $\sigma(X_0, \epsilon_0, \dots, \epsilon_{n-1})$  measurable. Indeed it is clearly true for  $n = 0$  and by induction, if  $X_n$  is  $\sigma(X_0, \epsilon_0, \dots, \epsilon_{n-1})$  measurable, then  $X_n = f_n(X_0, \epsilon_0, \dots, \epsilon_{n-1})$  for a certain measurable function  $f_n$  and so

$$X_{n+1} = \phi_n(X_n, \epsilon_n) = \phi_n(f_n(X_0, \epsilon_0, \dots, \epsilon_{n-1}), \epsilon_n)$$

is  $\sigma(X_0, \epsilon_0, \dots, \epsilon_n)$  measurable.

Since by assumption,  $\epsilon_n$  is independent of  $\sigma(X_0, \epsilon_0, \dots, \epsilon_{n-1})$ , we get by the above result that  $\epsilon_n$  is independent of  $\sigma(X_0, \dots, X_n)$ . Hence

$$\begin{aligned} E[\mathbb{1}_{\{X_{n+1} \in A\}} | \sigma(X_0, \dots, X_n)] &= E[\mathbb{1}_{\{\phi_n(X_n, \epsilon_n) \in A\}} | \sigma(X_0, \dots, X_n)] \\ &= P_n(X_n, A) \end{aligned}$$

where in the last equality we used the following proposition: if  $X, Y$  are random variables such that  $X$  is  $\mathcal{G}$  measurable with respect to some sigma algebra  $\mathcal{G}$  and  $Y$  is independent of  $\mathcal{G}$ , then

$$E[f(X, Y) | \mathcal{G}] = E[f(z, Y)] \Big|_{z=X}$$

- (b) The equality holds for  $f = \mathbb{1}_A$  for any  $A \in \mathcal{E}$  by the previous question. By linearity it also holds for simple functions, i.e functions for the form  $f = \sum_{k=1}^n a_k \mathbb{1}_{A_k}$ . Finally it also holds for positive measurable functions since every such functions is the increasing limit of a sequence of simple functions.

**Solution 11.3**

- (a) The proof is very similar to the solution of the question 11.1.a). Again by induction one can see that  $X_n$  and  $U_n$  are  $\sigma(X_0, \epsilon_0, \dots, \epsilon_{n-1})$  measurable. Indeed it is again clearly true for  $n = 0$  and by induction, if  $X_n$  and  $U_n$  are  $\sigma(X_0, \epsilon_0, \dots, \epsilon_{n-1})$  measurable, then there exists some measurable functions  $f_n$  and  $g_n$  such that  $X_n = f_n(X_0, \epsilon_0, \dots, \epsilon_{n-1})$  and  $U_n = g_n(X_0, \epsilon_0, \dots, \epsilon_{n-1})$

$$X_{n+1} = \phi_n(X_n, U_n, \epsilon_n) = \phi_n(f_n(X_0, \epsilon_0, \dots, \epsilon_{n-1}), g_n(X_0, \epsilon_0, \dots, \epsilon_{n-1}), \epsilon_n)$$

is  $\sigma(X_0, \epsilon_0, \dots, \epsilon_n)$  measurable and thus so is  $U_{n+1}$  since  $U$  is adapted to the natural filtration of  $X$ .

Since by assumption,  $\epsilon_n$  is independent of  $\sigma(X_0, \epsilon_0, \dots, \epsilon_{n-1})$ , we get by the above result that  $\epsilon_n$  is independent of  $\sigma(X_0, \dots, X_n)$ . Hence we have (using the same proposition as in question 11.1.a)),

$$\begin{aligned} E \left[ \mathbb{1}_{\{X_{n+1} \in A\}} | \sigma(X_0, \dots, X_n) \right] &= E \left[ \mathbb{1}_{\{\phi_n(X_n, U_n, \epsilon_n) \in A\}} | \sigma(X_0, \dots, X_n) \right] \\ &= P_n^{(U_n)}(X_n, A) \end{aligned}$$

- (b) The solution follows the same ideas as question 11.1.b). The equality holds for  $f = \mathbb{1}_A$  for any  $A \in \mathcal{E}$  by the previous question. By linearity it also holds for simple functions, i.e functions for the form  $f = \sum_{k=1}^n a_k \mathbb{1}_{A_k}$ . Finally it also holds for positive measurable function since every such functions is the increasing limit of a sequence of simple functions.

**Solution 11.4** As suggested by the hint, we use bakward induction for both questions.

- (a) It is clearly true for  $n = N$ :

$$O_N(X_N) = \gamma(X_N) = E_{x_0}^{\eta^{min}} [\gamma(X_N) | \sigma(X_0, \dots, X_N)]$$

Now we suppose that the desired property holds for  $n + 1$  and show that it

still holds for  $n$

$$\begin{aligned}
O_n(X_n) &= c_n(X_n, U_n^{(min)}(X_n)) + (P_n^{(U_n^{(min)}(X_n))} O_{n+1})(X_n) \quad (\text{by definition}) \\
&= c_n(X_n, U_n^{(min)}(X_n)) + E_{x_0}^{\eta^{min}} [O_{n+1}(X_{n+1}) | \sigma(X_0, \dots, X_n)] \\
&\quad (\text{using the result from question 2}) \\
&= c_n(X_n, U_n^{(min)}(X_n)) \\
&\quad + E_{x_0}^{\eta^{min}} \left[ E_{x_0}^{\eta^{min}} \left[ \sum_{k=n+1}^{N-1} c_k(X_k, U_k) + \gamma(X_N) | \sigma(X_0, \dots, X_{n+1}) \right] | \sigma(X_0, \dots, X_n) \right] \\
&\quad (\text{by induction hypotheses}) \\
&= c_n(X_n, U_n^{(min)}(X_n)) \\
&\quad + E_{x_0}^{\eta^{min}} \left[ \sum_{k=n+1}^{N-1} c_k(X_k, U_k) + \gamma(X_N) | \sigma(X_0, \dots, X_n) \right] \\
&\quad (\text{by tower property}) \\
&= E_{x_0}^{\eta^{min}} \left[ \sum_{k=n}^{N-1} c_k(X_k, U_k) + \gamma(X_N) | \sigma(X_0, \dots, X_n) \right]
\end{aligned}$$

- (b) Fix a strategy  $\eta$  corresponding to controls  $U$ . The claimed result is clearly true for  $n = N$ :

$$O_N(X_N) = \gamma(X_N) = E_{x_0}^{\eta} [\gamma(X_N) | \sigma(X_0, \dots, X_N)]$$

Now we suppose that the desired property holds for  $n + 1$  and show that it

still holds for  $n$

$$\begin{aligned}
O_n(X_n) &= c_n(X_n, U_n^{(min)}(X_n)) + (P_n^{(U_n^{(min)}(X_n))} O_{n+1})(X_n) \quad (\text{by definition}) \\
&\leq c_n(X_n, U_n(X_n)) + (P_n^{(U_n(X_n))} O_{n+1})(X_n) \quad (\text{by definition of optimal strategy}) \\
&= c_n(X_n, U_n(X_n)) + E_{x_0}^\eta [O_{n+1}(X_{n+1}) | \sigma(X_0, \dots, X_n)] \\
&\quad (\text{using the result from question 3}) \\
&\leq c_n(X_n, U_n(X_n)) \\
&\quad + E_{x_0}^\eta \left[ E_{x_0}^\eta \left[ \sum_{k=n+1}^{N-1} c_k(X_k, U_k) + \gamma(X_N) | \sigma(X_0, \dots, X_{n+1}) \right] | \sigma(X_0, \dots, X_n) \right] \\
&\quad (\text{by induction hypotheses + result from question (a)}) \\
&= c_n(X_n, U_n(X_n)) \\
&\quad + E_{x_0}^\eta \left[ \sum_{k=n+1}^{N-1} c_k(X_k, U_k) + \gamma(X_N) | \sigma(X_0, \dots, X_n) \right] \\
&\quad (\text{by tower property}) \\
&= E_{x_0}^\eta \left[ \sum_{k=n}^{N-1} c_k(X_k, U_k) + \gamma(X_N) | \sigma(X_0, \dots, X_n) \right]
\end{aligned}$$