

# Non-Life Insurance: Mathematics and Statistics

## Exercise sheet 2

### Exercise 2.1 Maximum Likelihood and Hypothesis Test

Let  $Y_1, \dots, Y_n$  be claim amounts in CHF that an insurance company has to pay. We assume that  $Y_1, \dots, Y_n$  are independent and identically distributed (i.i.d.) random variables following a log-normal distribution with unknown parameters  $\mu \in \mathbb{R}$  and  $\sigma^2 > 0$ . Then, by definition,  $\log Y_1, \dots, \log Y_n$  are i.i.d. Gaussian random variables with mean  $\mu \in \mathbb{R}$  and variance  $\sigma^2 > 0$ . Let  $n = 8$  and suppose that we have the following observations  $x_1, \dots, x_8$  for  $\log Y_1, \dots, \log Y_8$ :

$$x_1 = 9, \quad x_2 = 4, \quad x_3 = 6, \quad x_4 = 7, \quad x_5 = 3, \quad x_6 = 11, \quad x_7 = 6, \quad x_8 = 10.$$

- Write down the joint density  $f_{\mu, \sigma^2}(x_1, \dots, x_8)$  of  $\log Y_1, \dots, \log Y_8$ .
- Calculate  $\log f_{\mu, \sigma^2}(x_1, \dots, x_8)$ .
- Calculate the maximum likelihood estimates (MLEs)

$$(\hat{\mu}, \hat{\sigma}^2) = \arg \max_{(\mu, \sigma^2) \in \mathbb{R} \times \mathbb{R}_{>0}} \log f_{\mu, \sigma^2}(x_1, \dots, x_8).$$

- Now suppose that we are interested in the mean  $\mu$  of the logarithms of the claim amounts. An expert claims that  $\mu = 6$ . Perform a statistical test to test the null hypothesis  $H_0: \mu = 6$  against the (two-sided) alternative hypothesis  $H_1: \mu \neq 6$ .

### Exercise 2.2 Chebychev's Inequality and Law of Large Numbers

Suppose that an insurance company provides insurance against bike theft. In our model a bike gets stolen with a probability of 0.1, and in case of a theft the insurance company has to pay 1'000 CHF. We assume that we have  $n$  i.i.d. risks  $X_1, \dots, X_n$  with

$$X_i = \begin{cases} 1'000, & \text{with probability } 0.1, \\ 0, & \text{with probability } 0.9, \end{cases}$$

for all  $i = 1, \dots, n$ . In this exercise we are interested in the probability

$$p(n) \stackrel{\text{def}}{=} \mathbb{P} \left[ \left| \frac{1}{n} \sum_{i=1}^n X_i - \mu \right| \geq 0.1\mu \right]$$

of a deviation of the sample mean  $\frac{1}{n} \sum_{i=1}^n X_i$  to the mean claim size  $\mu = \mathbb{E}[X_1]$  of at least 10%, and how diversification effects this probability.

- Calculate  $\mu$ .
- Suppose that  $n = 1$ . Calculate  $p(1)$ .
- Suppose that  $n = 1'000$ . Calculate  $p(1'000)$ .
- Apply Chebychev's inequality to derive a minimum number  $n$  of risks such that  $p(n) < 0.01$ .
- What can you say about  $\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n X_i$ ?

### Exercise 2.3 Central Limit Theorem

Let  $n$  be the number of claims and  $Y_1, \dots, Y_n$  the corresponding claim sizes, where we assume that  $Y_1, \dots, Y_n$  are i.i.d. random variables with expectation  $\mathbb{E}[Y_1] = \mu$  and coefficient of variation  $\text{Vco}(Y_1) = 4$ .

- (a) Use the Central Limit Theorem to determine an approximate minimum number of claims  $n^{\text{CLT}}$  such that

$$\mathbb{P} \left[ \left| \frac{1}{n^{\text{CLT}}} \sum_{i=1}^{n^{\text{CLT}}} Y_i - \mu \right| < 0.01\mu \right] \geq 0.95,$$

i.e. with probability of at least 95% the deviation of the sample mean  $\frac{1}{n} \sum_{i=1}^n Y_i$  from the mean claim size  $\mu$  is less than 1%.

- (b) Compare the resulting minimum number of claims  $n^{\text{CLT}}$  to the corresponding minimum number of claims  $n^{\text{Che}}$  when using Chebychev's inequality instead of the Central Limit Theorem in part (a). What do you observe?

### Exercise 2.4 Conditional Distribution and Variance Decomposition

Suppose that  $\Theta$  follows an exponential distribution with parameter  $\lambda > 0$ . We assume that, conditionally given  $\Theta$ , the number of claims  $N$  in a particular line of business of an insurance company is modeled by a Poisson distribution with frequency parameter  $\Theta v$ , where  $v > 0$  denotes the volume, i.e. we have

$$\mathbb{P}[N = k|\Theta] = \begin{cases} e^{-\Theta v} \frac{(\Theta v)^k}{k!}, & \text{if } k \in \mathbb{N}, \\ 0, & \text{else.} \end{cases}$$

We remark that the expectation and the variance of a Poisson distribution are equal to its frequency parameter, i.e. here we have  $\mathbb{E}[N|\Theta] = \text{Var}(N|\Theta) = \Theta v$ .

- (a) Calculate  $\mathbb{P}[N = 0]$ .
- (b) Calculate  $\mathbb{E}[N]$ .
- (c) Show that  $\text{Var}(N) = \mathbb{E}[\text{Var}(N|\Theta)] + \text{Var}(\mathbb{E}[N|\Theta])$  and use this result to calculate  $\text{Var}(N)$ .