

# Brownian Motion and Stochastic Calculus

## Exercise sheet 4

**Exercise 4.1** Let  $W$  be a Brownian motion on  $[0, \infty)$  and  $S_0 > 0$ ,  $\sigma > 0$ ,  $\mu \in \mathbb{R}$  constants. The stochastic process  $S = (S_t)_{t \geq 0}$  given by

$$S_t = S_0 \exp(\sigma W_t + (\mu - \sigma^2/2)t)$$

is called *geometric Brownian motion*.

- (a) Prove that for  $\mu \neq \sigma^2/2$ , we have

$$\lim_{t \rightarrow \infty} S_t = \infty \quad P\text{-a.s.} \quad \text{or} \quad \lim_{t \rightarrow \infty} S_t = 0 \quad P\text{-a.s.}$$

When do the respective cases arise?

- (b) Discuss the behaviour of  $(S_t)$  as  $t \rightarrow \infty$  in the case  $\mu = \sigma^2/2$ .  
(c) Henceforth, suppose that  $\mu = 0$ . Show that  $S$  is a martingale, but not uniformly integrable.  
(d) Let  $\tau$  be a finite stopping time independent of  $W$ . Show that  $E[S_\tau] = S_0$ .  
(e) Fix  $S_0 = 1$ , let  $a \in (0, 1)$  and let  $\tau_a = \inf\{t : S_t \leq a\}$  be its hitting time. Show that  $\tau_a < \infty$  almost surely and that  $S_{\tau_a} = a < 1$ . In particular,  $E[S_{\tau_a}] = a < 1 = S_0$ .

### Solution 4.1

- (a) Noting that a.s.  $W_t/t \rightarrow 0$ , we have

- If  $(\mu - \sigma^2/2) > 0$ , then  $\sigma W_t + (\mu - \sigma^2/2)t \rightarrow \infty$  a.s., thus  $\lim_{t \rightarrow \infty} S_t = \infty$ .
- If  $(\mu - \sigma^2/2) < 0$ , then  $\sigma W_t + (\mu - \sigma^2/2)t \rightarrow -\infty$  a.s., thus  $\lim_{t \rightarrow \infty} S_t = 0$ .

- (b) The fact that a.s.  $\liminf_{t \rightarrow \infty} B_t = -\infty$  and  $\limsup_{t \rightarrow \infty} B_t = \infty$  implies that when  $\mu = \sigma^2/2$ ,  $\liminf_{t \rightarrow \infty} S_t = 0$  and  $\limsup_{t \rightarrow \infty} S_t = \infty$ . In particular,  $(S_t)$  almost surely does not converge as  $t \rightarrow \infty$ .

- (c) Note that if  $s \leq t$ , we have that  $W_t - W_s$  is independent of  $\mathcal{F}_s$  and follows the law of a centred normal with variance  $t - s$ , so that

$$\begin{aligned} E[S_t | \mathcal{F}_s] &= S_0 E[\exp(\sigma(W_t - W_s) + \sigma W_s - \sigma^2 t/2) | \mathcal{F}_s] \\ &= S_0 \exp(\sigma W_s - \sigma^2 s/2) E[\exp(\sigma(W_t - W_s) - \sigma^2(t - s)/2)] = S_s. \end{aligned}$$

Since  $S \geq 0$ , the same calculation with  $s = 0$  shows that  $S_t$  is integrable for all  $t \in [0, \infty)$ . Thus,  $(S_t)$  is a martingale that converges to 0 a.s. due to (a). By contradiction, suppose that it is uniformly integrable. We should have then  $S_0 = E[S_\infty] = 0$ , which does not hold.

- (d) Since  $\tau$  is independent of  $W$  and hence also of  $S$ , we can condition on  $\tau$  to find that

$$E[S_\tau] = E[E[S_\tau | \tau]] = E[E[S_t]_{t=\tau}] = S_0$$

because  $S$  is a martingale.

- (e) Since  $\lim_{t \rightarrow \infty} S_t = 0$  a.s., it follows that  $P[\exists t \geq 0 : S_t \leq a] = 1$ , and therefore  $P[\tau_a < \infty] = 1$ . As  $S_0 = 1$ , we have that  $P[\tau_a > 0] = 1$ . Note  $S$  is  $P$ -a.s. continuous; thus for some  $\tilde{\Omega} \subseteq \Omega$  with  $P[\tilde{\Omega}] = 1$ , we have for all  $\omega \in \tilde{\Omega}$  that  $S_{\tau_a(\omega)}(\omega) = \lim_{t \nearrow \tau_a(\omega)} S_t(\omega) \geq a$ , since  $S_t(\omega) > a$  for  $0 \leq t < \tau_a(\omega)$ , and  $S_{\tau_a(\omega)}(\omega) = \lim_{t \searrow \tau_a(\omega)} S_t(\omega) \leq a$ , since for any  $\varepsilon > 0$ , there exists  $t \in [\tau_a(\omega), \tau_a(\omega) + \varepsilon]$  such that  $S_t(\omega) \leq a$ . Therefore we must have that  $S_{\tau_a(\omega)}(\omega) = a$  for  $\omega \in \tilde{\Omega}$ , i.e.  $S_{\tau_a} = a$   $P$ -a.s.

**Exercise 4.2** Consider two stopping times  $\sigma, \tau$  on a filtered probability space  $(\Omega, \mathcal{F}, (\mathcal{F}_t), P)$ . The goal of this exercise, together with exercise **3.1**, is to show that

$$E[E[\cdot | \mathcal{F}_\sigma] | \mathcal{F}_\tau] = E[\cdot | \mathcal{F}_{\sigma \wedge \tau}] = E[E[\cdot | \mathcal{F}_\tau] | \mathcal{F}_\sigma] \quad P\text{-a.s.}, \quad (\star)$$

i.e., the operators  $E[\cdot | \mathcal{F}_\tau]$  and  $E[\cdot | \mathcal{F}_\sigma]$  commute and their composition equals  $E[\cdot | \mathcal{F}_{\sigma \wedge \tau}]$ .

*Remark:* For arbitrary sub- $\sigma$ -algebras  $\mathcal{G}, \mathcal{G}' \subseteq \mathcal{F}$ , the conditional expectations  $E[E[\cdot | \mathcal{G}] | \mathcal{G}']$ ,  $E[E[\cdot | \mathcal{G}'] | \mathcal{G}]$  and  $E[\cdot | \mathcal{G} \cap \mathcal{G}']$  do **not** coincide in general.

- (a) Show that if  $Y$  is  $\mathcal{F}_\sigma$ -measurable, then  $Y \mathbb{1}_{\{\sigma \leq \tau\}}$  and  $Y \mathbb{1}_{\{\sigma < \tau\}}$  are  $\mathcal{F}_{\sigma \wedge \tau}$ -measurable.
- (b) Show that  $E[Y | \mathcal{F}_\tau]$  is  $\mathcal{F}_{\sigma \wedge \tau}$ -measurable if  $Y$  is an integrable  $\mathcal{F}_\sigma$ -measurable random variable. Conclude  $(\star)$ .
- (c) Let  $M = (M_t)_{t \geq 0}$  be a martingale with all trajectories right-continuous. Show that the stopped process  $M^\tau = (M_{\tau \wedge t})_{t \geq 0}$  is again a martingale.  
*Hint:* Use  $(\star)$  and the stopping theorem.

**Solution 4.2**

- (a) Since  $\{\sigma \leq \tau\}, \{\sigma < \tau\} \in \mathcal{F}_\sigma$  by exercise **3.1(a)**, we have that  $Y \mathbb{1}_{\{\sigma \leq \tau\}}, Y \mathbb{1}_{\{\sigma < \tau\}}$  are both  $\mathcal{F}_\sigma$ -measurable. Now let us prove that they are  $\mathcal{F}_\tau$ -measurable. This holds if  $Y$  takes finitely many values. Indeed, let  $Y^n = \sum_{i=1}^n \lambda_i \mathbb{1}_{A_i}$  for some  $A_1, \dots, A_n \in \mathcal{F}_\sigma$  and  $\lambda_1, \dots, \lambda_n \in \mathbb{R}$ . Then  $Y^n \mathbb{1}_{\{\sigma \leq \tau\}}$  is  $\mathcal{F}_\tau$ -measurable if  $A_i \cap \mathbb{1}_{\{\sigma \leq \tau\}}$  is  $\mathcal{F}_\tau$ -measurable for each  $i$ , which holds by exercise **3.1(b)**. The argument for  $Y^n \mathbb{1}_{\{\sigma < \tau\}}$  is analogous.

For general  $Y$ , we can construct simple random variables  $Y^n$  of the above form such that  $Y^n(\omega) \rightarrow Y(\omega)$  for all  $\omega \in \Omega$ , and thus  $Y^n \mathbb{1}_{\{\sigma \leq \tau\}} \rightarrow Y \mathbb{1}_{\{\sigma \leq \tau\}}$ , which is therefore  $\mathcal{F}_\tau$ -measurable, and likewise for  $Y \mathbb{1}_{\{\sigma < \tau\}}$ . By exercise **3.1(a)**, we conclude that  $Y \mathbb{1}_{\{\sigma \leq \tau\}}$  and  $Y \mathbb{1}_{\{\sigma < \tau\}}$  are  $\mathcal{F}_{\sigma \wedge \tau}$ -measurable.

- (b) We note that

$$E[Y | \mathcal{F}_\tau] = E[Y \mathbb{1}_{\{\tau < \sigma\}} | \mathcal{F}_\tau] + E[Y \mathbb{1}_{\{\sigma \leq \tau\}} | \mathcal{F}_\tau] = E[Y | \mathcal{F}_\tau] \mathbb{1}_{\tau < \sigma} + Y \mathbb{1}_{\sigma \leq \tau},$$

where each term is  $\mathcal{F}_{\sigma \wedge \tau}$ -measurable by (a) so that  $E[Y | \mathcal{F}_\tau]$  is  $\mathcal{F}_{\sigma \wedge \tau}$ -measurable.

To show  $(\star)$  it is enough to note that if  $Z$  is integrable, then  $E[Z | \mathcal{F}_\sigma]$  is  $\mathcal{F}_\sigma$ -measurable and  $E[E[Z | \mathcal{F}_\sigma] | \mathcal{F}_\tau]$  is  $\mathcal{F}_{\sigma \wedge \tau}$ -measurable. Therefore

$$E[E[Z | \mathcal{F}_\sigma] | \mathcal{F}_\tau] = E[E[E[Z | \mathcal{F}_\sigma] | \mathcal{F}_\tau] | \mathcal{F}_{\sigma \wedge \tau}] = E[E[Z | \mathcal{F}_\sigma] | \mathcal{F}_{\sigma \wedge \tau}] = E[Z | \mathcal{F}_{\sigma \wedge \tau}],$$

by the tower property. The other direction follows by symmetry.

- (c) Take  $s \leq t$  and note that  $\tau \wedge s \leq \tau \wedge t$  are bounded stopping times. By the stopping theorem,

$$E[M_{\tau \wedge t} | \mathcal{F}_s] = E[E[M_t | \mathcal{F}_{\tau \wedge t}] | \mathcal{F}_s] = E[M_t | \mathcal{F}_{\tau \wedge s}] = M_{\tau \wedge s}$$

where in the second equality we used  $(\star)$ .

**Exercise 4.3** Let  $(S, \mathcal{S})$  be a measurable space, let  $Y = (Y_t)_{t \geq 0}$  be the canonical process on  $(\mathcal{S}^{[0, \infty)}, \mathcal{S}^{[0, \infty)})$ , i.e.,  $Y_t(y) = y(t)$  for  $y \in \mathcal{S}^{[0, \infty)}$ ,  $t \geq 0$ , and let  $(K_t)_{t \geq 0}$  be a transition semigroup on  $(S, \mathcal{S})$ . Moreover, for each  $x \in S$ , assume that there exists a unique probability measure  $\mathbb{P}_x$  on  $(\mathcal{S}^{[0, \infty)}, \mathcal{S}^{[0, \infty)})$  under which  $Y$  is a Markov process with transition semigroup  $(K_t)_{t \geq 0}$  and initial distribution  $\nu = \delta_{\{x\}}$ .

Suppose  $Z \geq 0$  is an  $\mathcal{S}^{[0, \infty)}$ -measurable random variable on  $\mathcal{S}^{[0, \infty)}$ . Use the monotone class theorem to prove that the map  $x \mapsto \mathbb{E}_x[Z]$ ,  $x \in S$ , is  $\mathcal{S}$ -measurable.

**Solution 4.3** Let  $\mathcal{H}$  denote the real vector space of all bounded,  $\mathcal{S}^{[0, \infty)}$ -measurable functions  $Z : \mathcal{S}^{[0, \infty)} \rightarrow \mathbb{R}$  such that the map  $x \mapsto \mathbb{E}_x[Z]$ ,  $x \in S$ , is  $\mathcal{S}$ -measurable. Since pointwise limits of measurable functions are measurable,  $\mathcal{H}$  is closed under monotone bounded convergence. The family

$$\mathcal{M} = \left\{ \prod_{k=0}^n f_k(Y_{t_k}) : n \in \mathbb{N}, 0 = t_0 < t_1 < \dots < t_n, f_k : S \rightarrow \mathbb{R} \text{ } \mathcal{S}\text{-measurable and bounded} \right\}$$

is closed under multiplication and  $\sigma(\mathcal{M}) = \mathcal{S}^{[0, \infty)}$ . It remains to show that  $\mathcal{M} \subseteq \mathcal{H}$  (note that  $1 \in \mathcal{M}$ ). Indeed, for an element  $Z = \prod_{k=0}^n f_k(Y_{t_k})$  in  $\mathcal{M}$ , we have for all  $x \in S$  that

$$\begin{aligned} \mathbb{E}_x[Z] &= \int_S \delta_{\{x\}}(dx_0) f_0(x_0) \int_S K_{t_1-t_0}(x_0, dx_1) f_1(x_1) \cdots \int_S K_{t_n-t_{n-1}}(x_{n-1}, dx_n) f_n(x_n) \\ &= f_0(x) \int_S K_{t_1-t_0}(x, dx_1) f_1(x_1) \cdots \int_S K_{t_n-t_{n-1}}(x_{n-1}, dx_n) f_n(x_n). \end{aligned} \quad (1)$$

Using measure-theoretic induction, it is easy to see that  $x \mapsto \int_S f(y) K(x, dy)$ ,  $x \in S$ , is  $\mathcal{S}$ -measurable for any bounded,  $\mathcal{S}$ -measurable function  $f : S \rightarrow \mathbb{R}$  and any stochastic kernel  $K$  on  $(S, \mathcal{S})$  (in fact, a more general result follows from the proof of Fubini's theorem for measures of the form  $\nu \otimes K$ , see for instance lecture notes "Wahrscheinlichkeitstheorie" (Föllmer/Schweizer), proof of Theorem II.1.4). Using this fact inductively in (1), we conclude that  $x \mapsto \mathbb{E}_x[Z]$  is  $\mathcal{S}$ -measurable for  $Z \in \mathcal{M}$ .

The monotone class theorem implies that  $\mathcal{H}$  contains all bounded,  $\mathcal{S}^{[0, \infty)}$ -measurable  $Z$ . For a general  $\mathcal{S}^{[0, \infty)}$ -measurable  $Z \geq 0$ , we have for each  $x \in S$  that  $\mathbb{E}_x[Z] = \lim_{n \rightarrow \infty} \mathbb{E}_x[Z \wedge n]$  by monotone convergence. Thus, as a pointwise limit of  $\mathcal{S}$ -measurable functions,  $x \mapsto \mathbb{E}_x[Z]$  is  $\mathcal{S}$ -measurable.

**Exercise 4.4** Part (a) of this exercise is optional, but the results are needed in (b) and (c).

(a) Let  $L \in \mathbb{N}$  and consider a matrix  $Q \in \mathbb{R}^{L \times L}$ . For  $t \in [0, +\infty)$ , define  $\exp(tQ) := \sum_{n=0}^{\infty} \frac{t^n Q^n}{n!}$ .

1. Show that  $\exp(tQ)$  is well-defined for any  $t \in [0, +\infty)$  and  $\exp(0Q) = \mathbb{I}$ , the identity matrix.
2. Show that  $Q^n \exp(tQ) = \exp(tQ)Q^n$  and that  $\exp(sQ) \exp(tQ) = \exp(tQ) \exp(sQ) = \exp((t+s)Q)$  for any  $n \in \mathbb{N}$  and  $s, t \geq 0$ .
3. Show that

$$\lim_{h \searrow 0} \frac{\exp((t+h)Q) - \exp(tQ)}{h} = Q \exp(tQ).$$

4. Show that

$$\lim_{M \rightarrow \infty} \left(1 + \frac{tQ}{M}\right)^M = \exp(tQ).$$

(b) Consider  $S = \{x_1, \dots, x_L\} \subseteq \mathbb{R}$ . Define the operators  $(K_t)_{t \geq 0}$  by

$$K_t(x_i, A) = \sum_{x_j \in A} (\exp(tQ))_{ij}, \quad \text{for } A \subseteq S.$$

Show that  $K_{s+t}(x_i, \{x_\ell\}) = \sum_{j=1}^L K_s(x_i, \{x_j\}) K_t(x_j, \{x_\ell\})$  for  $s, t \geq 0$ .

(c) Suppose that there exists a Markov process  $X$  taking values in  $S$  such that

$$\mathbb{P}_{x_i}[X_t = x_j] = K_t(x_i, \{x_j\}) = (\exp(tQ))_{ij}. \quad (2)$$

Noting the fact that for all  $t \geq 0$  and  $x_i \in S$ , the map  $A \mapsto K_t(x_i, A)$  must then be a probability measure, what does this imply about  $Q$ ?

**Solution 4.4**

(a) 1. Let  $\|\cdot\|$  be the operator norm for a matrix, defined as

$$\|A\| = \sup_{v \neq 0} \frac{|Av|}{|v|}.$$

The set of matrices  $\mathbb{R}^{L \times L}$  is a Banach space with respect to this norm. For any  $n \geq 0$ , we have the inequality  $\|A^n\| \leq \|A\|^n$ , and therefore

$$\sum_{n=0}^{\infty} \left\| \frac{t^n Q^n}{n!} \right\| \leq \sum_{n=0}^{\infty} \frac{t^n \|Q\|^n}{n!} = \exp(t\|Q\|) < \infty.$$

Therefore,  $\exp(tQ)$  is well-defined as an absolutely convergent series in  $\mathbb{R}^{L \times L}$ .

2. We have that

$$Q^n \exp(tQ) = Q^n \sum_{m=0}^{\infty} \frac{t^m Q^m}{m!} = \sum_{m=0}^{\infty} \frac{t^m Q^{m+n}}{m!} = \exp(tQ) Q^n$$

and

$$\exp(sQ) \exp(tQ) = \sum_{m=0}^{\infty} \sum_{n=0}^{\infty} \frac{s^m t^n Q^{m+n}}{m!n!} = \sum_{l=0}^{\infty} \frac{Q^l}{l!} \sum_{m+n=l} \binom{l}{m} s^m t^n = \sum_{l=0}^{\infty} \frac{(s+t)^l Q^l}{l!} = \exp((s+t)Q).$$

3. We have that

$$\exp((t+h)Q) = \exp(tQ) \exp(hQ) = \exp(tQ) (\mathbb{I} + hQ + O(h^2))$$

and therefore

$$\frac{d}{dt} \exp(tQ) = \exp(tQ)Q = Q \exp(tQ).$$

4. For each  $k \in \mathbb{N}$ , we have that

$$\lim_{M \rightarrow \infty} \binom{M}{k} \frac{(tQ)^k}{M^k} = \lim_{M \rightarrow \infty} \frac{M(M-1) \cdots (M-k+1)}{M^k} \frac{(tQ)^k}{k!} = \frac{(tQ)^k}{k!}.$$

Moreover, we have the uniform bound  $\left\| \binom{M}{k} \frac{(tQ)^k}{M^k} \right\| \leq \frac{(t\|Q\|)^k}{k!}$ , which is summable. Therefore, by the dominated convergence theorem (with respect to the counting measure),

$$\lim_{M \rightarrow \infty} \left( \mathbb{I} + \frac{tQ}{M} \right)^M = \lim_{M \rightarrow \infty} \sum_{k=0}^M \binom{M}{k} \frac{(tQ)^k}{M^k} = \sum_{k=0}^{\infty} \frac{(tQ)^k}{k!} = \exp(tQ).$$

(b) We calculate

$$\begin{aligned} \sum_{j=1}^L K_s(x_i, \{x_j\}) K_t(x_j, \{x_\ell\}) &= \sum_{j=1}^L (\exp(sQ))_{ij} \exp(tQ)_{j\ell} \\ &= (\exp(sQ) \exp(tQ))_{i\ell} \\ &= \exp((s+t)Q)_{i\ell} \\ &= K_{s+t}(x_i, \{x_\ell\}). \end{aligned}$$

(c) Suppose that such a Markov process exists. Note that for each  $i, j \in \{1, \dots, L\}$ , we have that  $(\exp(0Q))_{ij} = \mathbb{I}_{ij} = \mathbb{1}_{\{i=j\}}$ , by (a1). From (a3), we find that for  $i \neq j$ ,

$$\lim_{h \searrow 0} \frac{\mathbb{P}_{x_i}[X_h = x_j]}{h} = \lim_{h \searrow 0} \frac{\exp(hQ)_{ij} - \exp(0Q)_{ij}}{h} = Q_{ij}$$

and

$$\lim_{h \searrow 0} \frac{\mathbb{P}_{x_i}(X_h = x_i) - 1}{h} = \lim_{h \searrow 0} \frac{\exp(hQ)_{ii} - \exp(0Q)_{ii}}{h} = Q_{ii}.$$

for  $i \in \{1, \dots, L\}$ .

Since  $X$  exists, the operator  $K_t$  must be a transition kernel, and therefore  $A \mapsto K_t(x_i, A)$  must be a probability measure for each  $t \geq 0$  and  $x_i \in S$ . In particular, note that

$$K_t(x_i, S) = 1 = \sum_{j=1}^L (\exp(tQ))_{ij}$$

and therefore  $\sum_{j=1}^L (\exp(tQ))_{ij} = 1$  must hold for all  $t \geq 0$ . Taking the derivative at  $t = 0$ , we find that  $\sum_{j=1}^L Q_{ij} = 0$  for each  $i$ .

Moreover, we must have that  $K_t(x_i, \{x_j\}) = (\exp(tQ))_{ij} \geq 0$  for all  $i, j$ . Since  $(\exp(tQ))_{ij} = 0$  for  $t = 0$  and  $i \neq j$ , we can take the derivative to find that  $Q_{ij} \geq 0$  for all  $i \neq j$ .

We conclude that the conditions  $Q_{ij} \geq 0$  for  $i \neq j$  and  $Q_{ii} = -\sum_{i \neq j} Q_{ij}$  are necessary. One can show that they are also sufficient.