## Brownian Motion and Stochastic Calculus Exercise sheet 9

**Exercise 9.1** Let W be a Brownian motion with respect to its natural filtration. Show that

$$M_t^{(1)} = e^{t/2} \cos W_t, \quad M_t^{(2)} = tW_t - \int_0^t W_u du, \quad M_t^{(3)} = W_t^3 - 3tW_t$$

are martingales.

*Remark:* Prove this with a different method than the one used in Exercise 3.3.

*Hint:* Recall that for each  $t \ge 0$ , the running maximum  $W_t^* := \sup_{0 \le s \le t} |W_s|$  has the same distribution as  $|W_t|$ .

**Solution 9.1** We can express  $M^{(1)}, M^{(2)}, M^{(3)}$  in the form

$$M_t^{(1)} = f^{(1)}(t, W_t), \quad M_t^{(2)} = f^{(2)}\left(t, W_t, \int_0^t W_u du\right), \quad M_t^{(3)} = f^{(3)}(t, W_t),$$

where

$$f^{(1)}(t,w) = e^{t/2}\cos w, \quad f^{(2)}(t,w,x) = tw - x, \quad f^{(3)}(t,w) = w^3 - 3tw$$

are  $C^2$  functions. We note that the processes I and X defined by  $I_t = t$  and  $X_t = \int_0^t W_u du$ , respectively, are continuous and have finite variation, while  $\langle W \rangle_t = t$ . Therefore, by Itô's formula,

$$\begin{split} M_t^{(1)} &= M_0^{(1)} + \int_0^t \frac{\partial f^{(1)}}{\partial t} (s, W_s) ds + \int_0^t \frac{\partial f^{(1)}}{\partial w} (s, W_s) dW_s + \frac{1}{2} \int_0^t \frac{\partial^2 f^{(1)}}{\partial w^2} (s, W_s) ds \\ &= 1 + \frac{1}{2} \int_0^t e^{s/2} \cos W_s ds - \int_0^t e^{s/2} \sin W_s dW_s - \frac{1}{2} \int_0^t e^{s/2} \cos W_s ds \\ &= 1 - \int_0^t e^{s/2} \sin W_s dW_s, \\ M_t^{(2)} &= M_0^{(2)} + \int_0^t \frac{\partial f^{(2)}}{\partial t} (s, W_s, X_s) ds + \int_0^t \frac{\partial f^{(2)}}{\partial w} (s, W_s, X_s) dW_s + \int_0^t \frac{\partial f^{(2)}}{\partial x} (s, W_s, X_s) dX_s \\ &+ \frac{1}{2} \int_0^t \frac{\partial^2 f^{(2)}}{\partial w^2} (s, W_s, X_s) ds \\ &= \int_0^t W_s ds + \int_0^t s dW_s - X_t \\ &= \int_0^t s dW_s, \\ M_t^{(3)} &= M_0^{(3)} + \int_0^t \frac{\partial f^{(3)}}{\partial t} (s, W_s) ds + \int_0^t \frac{\partial f^{(3)}}{\partial w} (s, W_s) dW_s + \frac{1}{2} \int_0^t \frac{\partial^2 f^{(3)}}{\partial w^2} (s, W_s) ds \\ &= -3 \int_0^t W_s ds + \int_0^t (3W_s^2 - 3s) dW_s + 3 \int_0^t W_s ds \\ &= \int_0^t (3W_s^2 - 3s) dW_s. \end{split}$$

Since the integrands are continuous, hence locally bounded, we immediately get that  $M^{(1)}, M^{(2)}, M^{(3)}$  are local martingales. To show that they are martingales, note that  $W_t^* \stackrel{(d)}{=} |W_t|$ . Since all moments

of a Gaussian distribution are finite, the same must then be true of  $W^{\ast}_t.$  Therefore,

$$E\left[\int_0^T e^s (\sin W_s)^2 d\langle W \rangle_s\right] = E\left[\int_0^T e^s (\sin W_s)^2 ds\right] \le Te^T < \infty,$$
$$E\left[\int_0^T s^2 d\langle W \rangle_s\right] = E\left[\int_0^T s^2 ds\right] = T^3/3 < \infty,$$
$$E\left[\int_0^T (3W_s^2 - 3s)^2 d\langle W \rangle_s\right] = E\left[\int_0^T (3W_s^2 - 3s)^2 ds\right] \le TE[(3(W_T^*)^2 + 3T)^2] < \infty,$$

which shows that  $(M^{(1)})^T, (M^{(2)})^T, (M^{(3)})^T \in \mathcal{H}^{2,c}$  for any T > 0. In particular,  $M^{(1)}, M^{(2)}$  and  $M^{(3)}$  are martingales.

**Exercise 9.2** Let  $W = (W_t)_{t>0}$  be a 1-dimensional Brownian motion.

- (a) Let  $p : \mathbb{R} \to \mathbb{R}$  be a polynomial. Show that the process H = p(W) is in  $L^2_{loc}(W)$ , and therefore the stochastic integral  $\int p(W)dW$  is well defined. Moreover, show that  $\int p(W)dW$  is also a martingale.
- (b) For what polynomials  $p : \mathbb{R}^2 \to \mathbb{R}$  is the process  $X_t = p(W_t, t)$  a martingale? Given a fixed  $\lambda \in \mathbb{R}$ , for what polynomials  $p : \mathbb{R}^2 \to \mathbb{R}$  is the process  $Y_t = e^{-\lambda t} p(W_t, t)$  a martingale?
- (c) Let W' be another Brownian motion independent of W and  $\rho$  a predictable process satisfying  $|\rho| \leq 1$ . Prove that the process  $B = (B_t)_{t>0}$  given by

$$B_t = \int_0^t \rho_s \, dW_s + \int_0^t \sqrt{1 - \rho_s^2} \, dW'_s$$

is a Brownian motion. Moreover, compute  $\langle B, W \rangle$ .

*Remark:* The pair (W, B) is sometimes called *correlated Brownian motion* with *instantaneous* correlation  $\rho$ .

(d) Define  $Y_t = (\cos W_t, \sin W_t)^{\top}$  and  $Z_t = (-\sin W_t, \cos W_t)^{\top}$ . Show that Y is not a martingale, but  $Z \cdot Y$  is a martingale.

## Solution 9.2

(a) By linearity, it suffices to check the claim for monomials of the form  $p(x) = x^m, m \in \mathbb{N}$ . Since H = p(W) is continuous and adapted, it is predictable and locally bounded, so  $H \in L^2_{loc}(W)$ . Alternatively, set  $\tau_n := n$  for all  $n \in \mathbb{N}$ . Note that  $E[W_s^{2m}] = E[W_1^{2m}] s^m = Cs^m < \infty$  for  $s \geq 0$ , so by Fubini's theorem,

$$E\left[\int_0^{\tau_n} H_s^2 d\langle W \rangle_s\right] = E\left[\int_0^n W_s^{2m} ds\right] = \int_0^n E[W_s^{2m}] ds < \infty.$$

Therefore,  $\int p(W)dW$  is well defined and a local martingale. The same computations also give that, for all  $T \ge 0$ ,

$$E\left[\left\langle \int p(W)dW\right\rangle_T\right] = E\left[\int_0^T W_s^{2m}d\langle W\rangle_s\right] = E\left[\int_0^T W_s^{2m}ds\right] < \infty.$$

This proves that  $(\int p(W)dW)^T \in \mathcal{H}_0^{2,c}$  for all  $T \ge 0$  by Exercise 8.3(a), implying that  $\int p(W)dW$  is a true martingale.

(b) Note that the process  $I_t := t$  has finite variation, and therefore  $\langle I \rangle \equiv \langle W, I \rangle \equiv 0$ . Moreover,  $\langle W \rangle_t = t$  for  $t \ge 0$ .

Since any polynomial is  $C^2$ , by Itô's formula

$$X_t = X_0 + \int_0^t \partial_x p(W_s, s) dW_s + \int_0^t \partial_t p(W_s, s) ds + \frac{1}{2} \int_0^t \partial_{xx} p(W_s, s) ds.$$

Note that  $\partial_x p$  is a polynomial, therefore by (a),  $\int_0^{\cdot} \partial_x p(W_s, s) dW_s$  is a true martingale for any p. It follows that X is a martingale if  $\partial_t p = -\frac{1}{2} \partial_{xx} p$ .

Conversely, suppose that X is a martingale. Then, by Proposition 4.1.4 from the lecture notes,

$$\int_0^t \left( \partial_t p(W_s, s) + \frac{1}{2} \partial_{xx} p(W_s, s) \right) ds = 0, \quad t \ge 0,$$

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since the integral is a continuous martingale with finite variation. By continuity of the integrand in t, it follows that  $\partial_t p(W_t, t) + \frac{1}{2} \partial_{xx} p(W_t, t) = 0$  a.s. for all  $t \ge 0$ . In particular, it holds that

$$E\left[\left(\partial_t p(W_t, t) + \frac{1}{2}\partial_{xx} p(W_t, t)\right)^2\right] = 0, \quad \text{for all } t \ge 0.$$

Letting  $g_t$  be the Gaussian density function of  $W_t$ , we have that

$$\int_{-\infty}^{\infty} \left( \partial_t p(x,t) + \frac{1}{2} \partial_{xx} p(x,t) \right)^2 g_t(x) dx = 0, \quad \text{for all } t \ge 0.$$

Since the integrand is nonnegative,  $\partial_t p(x,t) + \frac{1}{2} \partial_{xx} p(x,t) = 0$  for  $\lambda$ -almost all  $x \in \mathbb{R}$  and all  $t \geq 0$ . By continuity,  $\partial_t p(x,t) = -\frac{1}{2} \partial_{xx} p(x,t)$  must hold for all  $x \in \mathbb{R}, t \geq 0$ . By properties of polynomials (e.g., since polynomials are analytic), we obtain that  $\partial_t p \equiv -\frac{1}{2} \partial_{xx} p$ .

Consider now the function  $f(x,t) = e^{-\lambda t} p(x,t)$ , which is  $C^2$ . By Itô's formula,

$$Y_t = X_0 + \int_0^t \partial_x f(W_s, s) dW_s + \int_0^t \left( e^{-\lambda s} \partial_t p(W_s, s) - \lambda f(W_s, s) \right) ds + \frac{1}{2} \int_0^t \partial_{xx} f(W_s, s) ds.$$

Since  $|\partial_x f(W_s, s)| = |e^{-\lambda s} \partial_x p(W_s, s)| \le (1 \vee e^{-\lambda t}) |\partial_x p(W_s, s)|$ , we find that  $\int_0^t \partial_x f(W_s, s) dW_s$  is a true martingale. By a similar argument as before, Y is a martingale if and only if

$$e^{-\lambda t}\partial_t p(x,t) = \lambda f(x,t) - \frac{1}{2}\partial_{xx}f(x,t)$$
  

$$\Leftrightarrow \partial_t p(x,t) = \lambda p(x,t) - \frac{1}{2}\partial_{xx}p(x,t) \quad \text{for all } x,t \in \mathbb{R}$$

(c) Since  $\rho$  is predictable and  $|\rho| \leq 1$ , we have that  $\rho \in L^2_{loc}(W)$  and  $\sqrt{1-\rho^2} \in L^2_{loc}(W')$ . It follows that B is a local martingale. Moreover, for each  $t \geq 0$ , using bilinearity of  $\langle \cdot, \cdot \rangle$  and the fact that  $\langle W, W' \rangle = 0$  due to independence of W and W',

$$\langle B \rangle_t = \left\langle \int \rho \, dW \right\rangle_t + \left\langle \int \sqrt{1 - \rho^2} \, dW' \right\rangle_t = \int_0^t \rho_s^2 \, ds + \int_0^t (1 - \rho_s^2) \, ds = t,$$

and so Lévy's characterisation of Brownian motion yields that B is a Brownian motion. Finally,

$$\langle B, W \rangle_t = \int_0^t \rho_s \, d\langle W, W \rangle_s + \int_0^t \sqrt{1 - \rho_s^2} \, d\langle W', W \rangle_s = \int_0^t \rho_s \, ds$$

(d) By Itô's formula, we have that

$$\cos W_{t} = 1 - \int_{0}^{t} \sin W_{s} dW_{s} - \frac{1}{2} \int_{0}^{t} \cos W_{s} ds$$
$$\sin W_{t} = \int_{0}^{t} \cos W_{s} dW_{s} - \frac{1}{2} \int_{0}^{t} \sin W_{s} ds.$$

In vector form, we obtain that

$$Y_{t} = Y_{0} + \int_{0}^{t} Z_{s} dW_{s} - \frac{1}{2} \int_{0}^{t} Y_{s} ds$$

Since the integrand is bounded, it is clear that  $\int_0^t Z_s dW_s$  is a martingale. But the term  $\frac{1}{2} \int_0^t Y_s ds$  is nonzero, continuous and has finite variation, and therefore Y is not a martingale. By associativity of stochastic integrals, we have that

$$(Z \bullet Y)_t = \int_0^t Z_s^\top Z_s dW_s - \frac{1}{2} \int_0^t Z_s^\top Y_s ds = W_t,$$

since  $Z_s^{\top} Z_s \equiv 1$  and  $Z_s^{\top} Y_s \equiv 0$ . In particular,  $Z \bullet Y$  is a martingale.

**Exercise 9.3** Let  $M \in \mathcal{H}_0^{2,c}$ . Show that  $b\mathcal{E}$  is dense in  $L^2(M)$ .

*Hint:* Let  $\overline{\Omega} = \Omega \times [0, \infty)$  be equipped with the predictable  $\sigma$ -algebra  $\mathcal{P}$ . Let  $C = E[M_{\infty}^2]$  and consider the probability measure  $P_M = C^{-1}P \otimes [M]$  on  $(\overline{\Omega}, \mathcal{P})$ . Let  $(\Pi_n)_{n \in \mathbb{N}}$  be an increasing sequence of partitions of  $[0, \infty)$  with  $\lim_{n \to \infty} |\Pi_n| = 0$ . Use the martingale convergence theorem on  $(\overline{\Omega}, \mathcal{P}, P_M)$  with respect to the discrete filtration  $(\mathcal{P}_n)_{n \in \mathbb{N}}$  defined by

$$\mathcal{P}_n = \sigma(\{A_i \times (t_i, t_{i+1}] : t_i \in \Pi_n, A_i \in \mathcal{F}_{t_i}\}).$$

**Solution 9.3** We first note that  $L^2(M) = L^2_{P_M}$ , since both are equal to the set of (equivalence classes of) predictable processes  $\tilde{H}$  such that

$$\|\tilde{H}\|_{L^{2}(M)}^{2} = E\left[\int_{0}^{\infty} \tilde{H}_{s}^{2} d\langle M \rangle_{s}\right] = C E_{M}[\tilde{H}^{2}] = C \|\tilde{H}\|_{L^{2}_{P_{M}}}^{2} < \infty$$

Let  $H \in L^2(M)$ . We want to approximate H in  $L^2(M)$  by elements of  $b\mathcal{E}$ . Since  $H\mathbb{1}_{\{|H| \leq n\}} \to H$ in  $L^2(M)$ , we only need to approximate each  $H\mathbb{1}_{\{|H| \leq n\}}$ . Thus, we assume w.l.o.g. that H is bounded.

Define a  $P_M$ -martingale  $(H_n)$  adapted to  $(\mathcal{P}_n)$  by  $H^n := E_M[H \mid \mathcal{P}_n]$ . Since  $H \in L^2_{P_M} = L^2(M)$ , we have that  $(H^n)$  is an  $L^2_{P_M}$ -bounded martingale. Let  $\mathcal{P}_{\infty} = \sigma(\bigcup_{n \in \mathbb{N}} \mathcal{P}_n)$  and  $H^{\infty} := E[H \mid \mathcal{P}_{\infty}]$ . By the martingale convergence theorem, we have that  $H^n \to H^{\infty} P_M$ -a.s. and in  $L^2_{P_M}$ .

We claim that  $\mathcal{P}_{\infty} = \mathcal{P}$  and each  $H^n \in b\mathcal{E}, n \in \mathbb{N}$ . If this holds, then we can approximate  $H = H^{\infty}$  in  $L^2_{P_M} = L^2(M)$  as the limit of  $(H^n)$ , where each  $H^n \in b\mathcal{E}$ . Thus, the two claims imply the result.

To show that  $\mathcal{P}_{\infty} = \mathcal{P}$ , we first note that each  $\mathcal{P}_n \subseteq \mathcal{P}$ . Indeed, let  $H = \mathbb{1}_{A_i} \mathbb{1}_{(t_i, t_{i+1}]}$  for some  $t_i \in \Pi_n$  and  $A_i \in \mathcal{F}_{t_i}$ . As  $\tilde{H}$  is adapted and left-continuous, it is predictable, i.e.,  $\mathcal{P}$ -measurable. Therefore, each  $\mathcal{P}_n \subseteq \mathcal{P}$ . Taking the union,  $\mathcal{P}_{\infty} \subseteq \mathcal{P}$ .

Conversely, to show that  $\mathcal{P} \subseteq \mathcal{P}_{\infty}$ , we show that any left-continuous adapted process  $\tilde{H}$  is  $\mathcal{P}_{\infty}$ -measurable. Define

$$\tilde{H}^n := \sum_{t_i \in \Pi_n} \mathbb{1}_{(t_i, t_{i+1}]} \tilde{H}_{t_i},$$

which is  $\mathcal{P}_n$ -measurable, hence also  $\mathcal{P}_{\infty}$ -measurable for each n. For each  $t \in [0, \infty)$  and  $n \in \mathbb{N}$ , we have that  $\tilde{H}_t^n(\omega) = \tilde{H}_{t(n)}(\omega)$ , where  $t(n) = \max\{t_i \in \Pi_n : t_i < t\}$ . We have that t(n) is increasing in n, since  $(\Pi_n)$  is an increasing sequence. Moreover,  $t(n) \nearrow t$  since  $|\Pi_n| \searrow 0$ . As  $\tilde{H}$  is left-continuous, we conclude that  $\tilde{H}_t^n(\omega) \to \tilde{H}_t(\omega)$  for all  $t \in \mathbb{R}_+$  and  $\omega \in \Omega$ . Therefore, as each  $\tilde{H}^n$ is  $\mathcal{P}_{\infty}$ -measurable, so is  $\tilde{H}$ . This shows that  $\mathcal{P}_{\infty} = \mathcal{P}$ , as we wanted.

Finally, we need to prove that  $H^n \in b\mathcal{E}$  for each n. We give two proofs for this.

**Proof 1:** Note that  $H^n = E[H | \mathcal{P}_n]$  is bounded, as H is. Since  $H^n$  is  $\mathcal{P}_n$ -measurable, the result follows if we show that every bounded  $\mathcal{P}_n$ -measurable process belongs to  $b\mathcal{E}$ . We show this using the monotone class theorem. Let

$$\mathcal{M} = \{\mathbb{1}_{A_i} \mathbb{1}_{(t_i, t_{i+1}]} : t_i \in \Pi_n, A_i \in \mathcal{F}_{t_i}\}$$

and

$$\mathcal{H} = \left\{ \tilde{H} = \sum_{t_i \in \Pi_n} Z_i \mathbb{1}_{(t_i, t_{i+1}]} : Z_i \text{ bounded and } \mathcal{F}_{t_i}\text{-measurable} \right\}$$

It is clear that  $\mathcal{M}$  is closed under products, generates  $\mathcal{P}_n$  and is contained in  $\mathcal{H}$ . Moreover,  $\mathcal{H}$  is a vector space and contains 1. To see that  $\mathcal{H}$  is closed under bounded monotone convergence, let  $\mathcal{H} \ni \tilde{H}^m \nearrow \tilde{H}$ . Then, it must be the case that  $Z_i^m = \tilde{H}_{t_{i+1}}^m \nearrow \tilde{H}_{t_{i+1}} =: Z_i$ , where  $Z_i$  is bounded and  $\mathcal{F}_{t_i}$ -measurable. Moreover, since

$$\tilde{H}_t = \lim_{m \to \infty} \tilde{H}_t^m = \lim_{m \to \infty} \tilde{H}_{t(n)}^m = \tilde{H}_{t(n)}, \quad t \ge 0,$$

we see that

$$\tilde{H} = \sum_{t_i \in \Pi_n} Z_i \mathbb{1}_{(t_i, t_{i+1}]} \in \mathcal{H}.$$

Therefore, by the monotone class theorem,  $\mathcal{H}$  contains all bounded  $\mathcal{P}_n$ -measurable processes. Since  $\mathcal{H} \subseteq b\mathcal{E}$ , we have shown that every bounded  $\mathcal{P}_n$ -measurable process belongs to  $b\mathcal{E}$ , as we wanted.

**Proof 2**: We claim that  $H_t^n = \sum_{t_i \in \Pi_n} \mathbb{1}_{\{t \in (t_i, t_{i+1}]\}} Z_i$  for each  $n \in \mathbb{N}$ , where

$$Z_{i} = \frac{E\left[\int_{t_{i}}^{t_{i+1}} H_{u} d\langle M \rangle_{u} \mid \mathcal{F}_{t_{i}}\right]}{E[\langle M \rangle_{t_{i+1}} - \langle M \rangle_{t_{i}} \mid \mathcal{F}_{t_{i}}]}.$$

If this holds, then  $H^n \in b\mathcal{E}$ , as  $Z_i$  is bounded and  $\mathcal{F}_{t_i}$ -measurable.

To show that  $H^n$  is the conditional expectation, note that  $H^n$  is  $\mathcal{P}_n$ -measurable. We also have that  $\{A_i \times (t_i, t_{i+1}] : t_i \in \Pi_n, A_i \in \mathcal{F}_{t_i}\}$  is a  $\pi$ -system generating  $\mathcal{P}_n$ . Therefore, it suffices to check that  $E_M[\mathbb{1}_{A_i \times (t_i, t_{i+1}]}H^n] = E_M[\mathbb{1}_{A_i \times (t_i, t_{i+1}]}H]$  for any  $t_i \in \Pi_n$  and  $A_i \in \mathcal{F}_{t_i}$ . Indeed,

$$\begin{split} E_M \left[ \mathbbm{1}_{A_i \times (t_i, t_{i+1}]} H \right] &= C^{-1} E \left[ \mathbbm{1}_{A_i} \int_{t_i}^{t_{i+1}} H_u d\langle M \rangle_u \right] \\ &= C^{-1} E \left[ \mathbbm{1}_{A_i} E \left[ \int_{t_i}^{t_{i+1}} H_u d\langle M \rangle_u \middle| \mathcal{F}_{t_i} \right] \right] \\ &= C^{-1} E \left[ \int_0^\infty \mathbbm{1}_{A_i} \frac{E \left[ \int_{t_i}^{t_{i+1}} H_u d\langle M \rangle_u \middle| \mathcal{F}_{t_i} \right]}{E[\langle M \rangle_{t_{i+1}} - \langle M \rangle_{t_i} \mid \mathcal{F}_{t_i}]} \mathbbm{1}_{\{s \in (t_i, t_{i+1}]\}} d\langle M \rangle_s \right] \\ &= E_M \left[ \mathbbm{1}_{A_i \times (t_i, t_{i+1}]} H^n \right], \end{split}$$

as we wanted.

**Exercise 9.4** Let  $d \ge 2$ ,  $\Omega = C([0,\infty); \mathbb{R}^d)$  and  $Y = (Y_t)_{t\ge 0}$  denote the coordinate process. For each  $x \in \mathbb{R}^d$ , let  $\mathbb{P}_x$  be the unique probability measure on  $(\Omega, \mathcal{Y}^0_{\infty})$  under which Y is a d-dimensional Brownian motion started at x.

Let  $x \in \mathbb{R}^d \setminus \{0\}$  and a, b such that 0 < a < |x| < b. Consider the stopping times

$$\tau_a := \inf \{ t \ge 0 : |Y_t| \le a \}, \quad \tau_b := \inf \{ t \ge 0 : |Y_t| \ge b \}.$$

- (a) Suppose that  $d \ge 3$ . Show that  $(X_t)_{t\ge 0}$  defined by  $X_t := |Y_{\tau_a \wedge t}|^{2-d}$  is a bounded martingale under  $\mathbb{P}_x$ .
- (b) Suppose that d = 2. Show that  $(X_t)_{t \ge 0}$  defined by  $X_t := -\log |Y_{\tau_a \wedge \tau_b \wedge t}|$  is a bounded martingale under  $\mathbb{P}_x$ .
- (c) Let  $d \ge 2$ . Show that for any  $x \in \mathbb{R}^d \setminus \{0\}$ ,

$$\mathbb{P}_x[Y_t \neq 0 \text{ for all } t \ge 0] = 1.$$

(d) Let  $d \geq 3$ . Show that for any  $x \in \mathbb{R}^d$ , we have

$$\mathbb{P}_x\Big[\lim_{t\to\infty}|Y_t|=\infty\Big]=1.$$

*Remark:* The result in (d) is known as *transience of Brownian motion in*  $\mathbb{R}^d$ , for  $d \geq 3$ .

## Solution 9.4

(a) For any  $C^2$  function  $g: (0,\infty) \to \mathbb{R}$ , we can define the radial function  $f: \mathbb{R}^d \setminus \{0\} \to \mathbb{R}$  by

f(x) := g(|x|).

By straightforward calculations, we have for every  $x \neq 0$  that

$$\Delta f(x) = g''(r) + \frac{d-1}{r}g'(r), \quad \text{with } r = |x|.$$

We consider  $g(r) := r^{2-d}$ , which is  $C^2$  on  $(0, \infty)$ , and let f(x) := g(|x|). For any  $x \neq 0$ ,

$$\Delta f(x) = g''(r) + \frac{d-1}{r}g'(r) = (2-d)(1-d)r^{-d} + (2-d)(d-1)r^{-d} = 0,$$

which means that  $f(x) = |x|^{2-d}$  is harmonic on  $\mathbb{R}^d \setminus \{0\}$ . By applying Itô's formula, as  $\Delta f = 0$  so that the second order term vanishes, we see that  $\mathbb{P}_x$ -a.s. for all  $t \ge 0$ 

$$X_t = \left|Y_{\tau_a \wedge t}\right|^{2-d} = f(Y_{\tau_a \wedge t}) = f(x) + \int_0^{\tau_a \wedge t} \nabla f(Y_s) \, dY_s,$$

which already shows that  $(X_t)_{t\geq 0}$  is a local martingale, since f is  $C^2$  and Y is continuous, so that  $\nabla f(Y_s)$  is locally bounded. Moreover, as  $d\geq 3$ , we have that

$$0 \le X_t = \left| Y_{\tau_a \wedge t} \right|^{2-d} = \frac{1}{|Y_{\tau_a \wedge t}|^{d-2}} \le \frac{1}{a^{d-2}}.$$

Thus, since X is uniformly bounded, it is a true martingale.

(b) Similarly to (a), set  $g(r) = -\log r$  and f(x) = g(|x|). We have that

$$\Delta f(x) = g''(r) + \frac{d-1}{r}g'(r) = \frac{1}{r^2} - \frac{1}{r^2} = 0,$$

therefore

$$X_t = -\log \left| Y_{\tau_a \wedge \tau_b \wedge t} \right| = f(Y_{\tau_a \wedge \tau_b \wedge t}) = f(x) + \int_0^{\tau_a \wedge \tau_b \wedge t} \nabla f(Y_s) \, dY_s.$$

By the same argument as in (a), X is a local martingale. Since  $a \leq |W_{\tau_a \wedge \tau_b \wedge t}| \leq b$ , we obtain that X is bounded by  $-\log b \leq X_t \leq -\log a$  and therefore it is a true martingale.

(c) For the cases  $d \ge 3$  and d = 2, let g and f be as defined in (a) and (b) respectively. Consider the martingale  $M_t = f(Y_{\tau_a \land \tau_b \land t})$  (which is the same as  $X^{\tau_b}$  in the case  $d \ge 3$ ). Note that  $\limsup_{t\to\infty} |Y_t| = \infty$ ; thus  $\tau_b < \infty$  and also  $\tau_a \land \tau_b < \infty$  a.s. Therefore, we have that  $\lim_{t\to\infty} M_t = f(Y_{\tau_a \land \tau_b})$  where  $|Y_{\tau_a \land \tau_b}| \in \{a, b\}$ . Since M is bounded, hence a uniformly integrable martingale, we obtain that

$$g(|x|) = M_0 = \mathbb{E}_x \left[ \lim_{t \to \infty} M_t \right] = g(a) \mathbb{P}_x[\tau_a < \tau_b] + g(b) \mathbb{P}_x[\tau_b < \tau_a].$$

Rearranging, we get

$$\mathbb{P}_x[\tau_a < \tau_b] = \frac{g(|x|) - g(b)\mathbb{P}_x[\tau_b < \tau_a]}{g(a)}$$

Note that in both cases d = 2 and  $d \ge 3$ , it holds that  $g(a) \to \infty$  as  $a \searrow 0$ . Therefore, taking the limit (by dominated convergence), we obtain that

$$\mathbb{P}_x[\tau_0 < \tau_b] = 0, \quad b > |x|.$$

Taking the limit again as  $b \to \infty$ , we obtain that

$$\mathbb{P}_x\left[\tau_0 < \lim_{b \to \infty} \tau_b\right] = 0.$$

However, since Y is locally bounded we have that  $\lim_{b\to\infty} \tau_b = \infty$ . Therefore,  $\tau_0 = \infty$  a.s. and

$$\mathbb{P}_x | X_t \neq 0 \text{ for all } t \geq 0 | = 1$$

as we wanted.

(d) By translation invariance of Brownian motion, we can assume that  $x \neq 0$  without loss of generality. For  $d \geq 3$ , define

$$M_t = g(|Y_t|) = |Y_t|^{2-d}.$$

This is  $\mathbb{P}_x$ -a.s. well defined for all  $t \ge 0$ , since  $\tau_0 = \infty$  a.s. by (c).

As in (a), M is a local martingale. In this case, while M is not bounded, it is a nonnegative local martingale, hence a supermartingale by Exercise 8.2. Since M is a nonnegative supermartingale, it also follows by the supermartingale convergence theorem that  $M_t \to M_\infty$ a.s. for some  $\mathcal{F}_\infty$ -measurable random variable  $M_\infty$ . Noting that

$$\limsup_{t \to \infty} |Y_t| = \infty \quad \mathbb{P}_x\text{-a.s.},$$

it follows that

$$M_{\infty} = \lim_{t \to \infty} M_t = \liminf_{t \to \infty} |Y_t|^{2-d} = 0,$$

and therefore

$$|Y_t| = M_t^{\frac{1}{2-d}} \to \infty \quad \mathbb{P}_x\text{-a.s.}$$

as we wanted.