

Mathematical Foundations for Finance

Exercise Sheet 12

Please hand in your solutions by 12:00 on Wednesday, December 14 via the course homepage.

Exercise 12.1 (*Stochastic product rule*) Let X and Y be two continuous semi-martingales. Show that for all $t \geq 0$,

$$X_t Y_t - X_0 Y_0 = \int_0^t X_s dY_s + \int_0^t Y_s dX_s + [X, Y]_t.$$

How does this compare to the "classical" product rule from calculus?

Solution 12.1 Let $f : \mathbb{R}^2 \rightarrow \mathbb{R}$ be given by $f(x, y) = xy$. Then $f \in C^2(\mathbb{R}^2; \mathbb{R})$, and $\frac{\partial f}{\partial x}(x, y) = y$, $\frac{\partial f}{\partial y}(x, y) = x$, $\frac{\partial^2 f}{\partial x^2}(x, y) = 0$, $\frac{\partial^2 f}{\partial x \partial y}(x, y) = 1$, and $\frac{\partial^2 f}{\partial y^2}(x, y) = 0$. Noting that $f(X_t, Y_t) = X_t Y_t$, Itô's formula then gives

$$\begin{aligned} X_t Y_t - X_0 Y_0 &= \int_0^t X_s dY_s + \int_0^t Y_s dX_s + \int_0^t d[X, Y]_s \\ &= \int_0^t X_s dY_s + \int_0^t Y_s dX_s + [X, Y]_t, \end{aligned}$$

as required.

The above equality has the extra " $[X, Y]_t$ " term, which vanishes in classical calculus.

Exercise 12.2 (*Stochastic exponential*) Let $X = (X_t)_{t \geq 0}$ be a continuous semi-martingale. Define the process $\mathcal{E}(X)$ by

$$\mathcal{E}(X)_t := \exp\left(X_t - \frac{1}{2}[X]_t\right).$$

(a) Show that $\mathcal{E}(X)$ is a solution to the SDE

$$Z_t = e^{X_0} + \int_0^t Z_s dX_s, \quad \forall t \geq 0. \tag{1}$$

(b) Prove that $\mathcal{E}(X)$ is the unique solution of (1).

Hint: For a solution Z of (1), consider the process $\frac{Z}{\mathcal{E}(X)}$.

(c) Let $Y = (Y_t)_{t \geq 0}$ be another continuous semimartingale. Prove *Yor's formula*

$$\mathcal{E}(X)\mathcal{E}(Y) = \mathcal{E}(X + Y + [X, Y]).$$

Solution 12.2

(a) Let $f(x, y) = \exp(x - \frac{1}{2}y)$. Then f is C^2 , and we have $\frac{\partial f}{\partial x}(x, y) = f(x, y)$, $\frac{\partial f}{\partial y}(x, y) = -\frac{1}{2}f(x, y)$, and $\frac{\partial^2 f}{\partial x^2}(x, y) = f(x, y)$. Noting that $\mathcal{E}(X)_t = f(X_t, [X]_t)$ (and $[[X]]_t = [X, [X]]_t = [[X], X]_t = 0$ because X is continuous), we apply Itô's formula to get

$$\begin{aligned} \mathcal{E}(X)_t &= \mathcal{E}(X)_0 + \int_0^t \mathcal{E}(X)_s dX_s + \int_0^t \left(-\frac{1}{2}\mathcal{E}(X)_s\right) d[X]_s + \frac{1}{2} \int_0^t \mathcal{E}(X)_s d[X]_s \\ &= e^{X_0} + \int_0^t \mathcal{E}(X)_s dX_s. \end{aligned}$$

It follows immediately that $\mathcal{E}(X)$ is a solution of (1).

(b) Suppose that Z is another solution to (1), and define $Y_t := \frac{Z_t}{\mathcal{E}(X)_t}$. Consider the function $f : (0, \infty) \times (0, \infty) \rightarrow \mathbb{R}$ defined by $f(x, y) = \frac{x}{y}$. Then f is C^2 , and $\frac{\partial f}{\partial x} = \frac{1}{y}$, $\frac{\partial f}{\partial y} = -\frac{x}{y^2}$, $\frac{\partial^2 f}{\partial x^2} = 0$, $\frac{\partial^2 f}{\partial x \partial y} = -\frac{1}{y^2}$, and $\frac{\partial^2 f}{\partial y^2} = \frac{2x}{y^3}$. Noting that $Y_t = f(Z_t, \mathcal{E}(X)_t)$, Itô's formula then gives

$$\begin{aligned} Y_t &= Y_0 + \int_0^t \frac{1}{\mathcal{E}(X)_s} dZ_s + \int_0^t \left(-\frac{Z_s}{\mathcal{E}(X)_s^2}\right) d\mathcal{E}(X)_t \\ &\quad - \int_0^t \frac{1}{\mathcal{E}(X)_s^2} d[Z, \mathcal{E}(X)]_s + \frac{1}{2} \int_0^t \frac{2Z_s}{\mathcal{E}(X)_s^3} d[Z]_s. \end{aligned}$$

Since $\mathcal{E}(X)$ and Z are solutions of (1), $d\mathcal{E}(X)_t = \mathcal{E}(X)_t dX_s$ and $dZ_t = Z_t dX_s$, and so $d[Z, \mathcal{E}(X)]_t = Z_t \mathcal{E}(X)_t d[X]_s$ and $d[Z]_t = Z_t^2 d[X]_s$. The above can now be rewritten as

$$\begin{aligned} Y_t &= Y_0 + \int_0^t \frac{Z_s}{\mathcal{E}(X)_s} dX_s + \int_0^t \left(-\frac{Z_s}{\mathcal{E}(X)_s^2} \mathcal{E}(X)_s\right) dX_s \\ &\quad - \int_0^t \frac{1}{\mathcal{E}(X)_s^2} Z_s \mathcal{E}(X)_s d[X]_s + \frac{1}{2} \int_0^t \frac{2Z_s}{\mathcal{E}(X)_s^3} \mathcal{E}(X)_s^2 d[X]_s \\ &= Y_0. \end{aligned}$$

Since $\mathcal{E}(X)_0 = e^{X_0} = Z_0$, then $Y_0 = 1$, and hence we have

$$Y \equiv 1.$$

It follows immediately that $Z = \mathcal{E}(X)$, and thus $\mathcal{E}(X)$ is the unique solution of (1).

(c) The stochastic product rule gives

$$\mathcal{E}(X)_t \mathcal{E}(Y)_t = e^{X_0} e^{Y_0} + \int_0^t \mathcal{E}(X)_s d\mathcal{E}(Y)_s + \int_0^t \mathcal{E}(Y)_s d\mathcal{E}(X)_s + [\mathcal{E}(X), \mathcal{E}(Y)]_t.$$

Using part (a), we have $d\mathcal{E}(X)_t = \mathcal{E}(X)_t dX_t$ and $d\mathcal{E}(Y)_t = \mathcal{E}(Y)_t dY_t$, and so $d[\mathcal{E}(X), \mathcal{E}(Y)]_t = \mathcal{E}(X)_t \mathcal{E}(Y)_t d[X, Y]_t$. Thus, we can rewrite the above as

$$\begin{aligned} \mathcal{E}(X)_t \mathcal{E}(Y)_t &= e^{X_0} e^{Y_0} + \int_0^t \mathcal{E}(X)_s \mathcal{E}(Y)_s dY_s + \int_0^t \mathcal{E}(Y)_s \mathcal{E}(X)_s dX_s \\ &\quad + \int_0^t \mathcal{E}(X)_s \mathcal{E}(Y)_s d[X, Y]_s \\ &= e^{X_0} e^{Y_0} + \int_0^t \mathcal{E}(X)_s \mathcal{E}(Y)_s d(X + Y + [X, Y])_s. \end{aligned}$$

We have thus shown that $\mathcal{E}(X)\mathcal{E}(Y)$ is a solution to the SDE

$$Z_t = e^{(X+Y+[X,Y])_0} + \int_0^t Z_s d(X + Y + [X, Y])_s, \quad \forall t \geq 0.$$

By part (b), it follows immediately that

$$\mathcal{E}(X)\mathcal{E}(Y) = \mathcal{E}(X + Y + [X, Y]),$$

as required.

Exercise 12.3 (*Itô process*) Let W be a Brownian motion with respect to P and \mathbb{F} . An *Itô process* is a stochastic process of the form

$$X_t = X_0 + \int_0^t \mu_s ds + \int_0^t \sigma_s dW_s, \quad t \geq 0,$$

where μ and σ are predictable processes (satisfying appropriate integrability conditions). Show that for any C^2 function f , the process $f(X)$ is again an Itô process, and give its decomposition.

Solution 12.3 By Itô's formula, we have

$$f(X_t) = f(X_0) + \int_0^t f'(X_s) dX_s + \frac{1}{2} \int_0^t f''(X_s) d[X]_s.$$

We know that $dX_t = \mu_t dt + \sigma_t dW_t$, and hence $d[X]_t = \sigma_t^2 dt$. The above equality can thus be rewritten as

$$\begin{aligned} f(X_t) &= f(X_0) + \int_0^t f'(X_s) \mu_s ds + \int_0^t f'(X_s) \sigma_s dW_s + \frac{1}{2} \int_0^t f''(X_s) \sigma_s^2 ds \\ &= f(X_0) + \int_0^t \left(\mu_s f'(X_s) + \frac{1}{2} \sigma_s^2 f''(X_s) \right) ds + \int_0^t \sigma_s f'(X_s) dW_s. \end{aligned}$$

Thus, $f(X)$ is an Itô process with the decomposition

$$\tilde{\mu}_s = \mu_s f'(X_s) + \frac{1}{2} \sigma_s^2 f''(X_s) \quad \text{and} \quad \tilde{\sigma}_s = \sigma_s f'(X_s).$$

Exercise 12.4 (*Distribution of stochastic integral*) Let $g : [0, \infty) \rightarrow \mathbb{R}$ be a continuous function. Show that for each $t \geq 0$, the random variable

$$X_t := \int_0^t g(s) dW_s$$

is normally distributed, and find its mean and variance.

Hint: For each fixed $\eta \in \mathbb{R}$, show that the stochastic process $Z = (Z_t)_{t \geq 0}$, given by $Z_t := e^{-\frac{\eta^2}{2} \int_0^t g^2(s) ds + \eta \int_0^t g(s) dW_s}$, is a local martingale (you may use without proof that Z is actually a true martingale), and then compute $E[e^{\eta \int_0^t g(s) dW_s}]$.

Solution 12.4 Following the hint, we first show that Z is a martingale. Letting $Y_t := \int_0^t g^2(s) ds$ (which is well-defined because g is bounded on $[0, t]$, since g is continuous), we can write

$$Z_t = e^{-\frac{\eta^2}{2} Y_t + \eta X_t}.$$

Let $f(x, y) = e^{-\frac{\eta^2}{2} y + \eta x}$. Then $f \in C^2(\mathbb{R}^2; \mathbb{R})$, and $\frac{\partial f}{\partial x} = \eta f$, $\frac{\partial f}{\partial y} = -\frac{\eta^2}{2} f$, $\frac{\partial^2 f}{\partial x^2} = \eta^2 f$, $\frac{\partial^2 f}{\partial x \partial y} = -\frac{\eta^3}{2} f$, and $\frac{\partial^2 f}{\partial y^2} = \frac{\eta^4}{4} f$. Noting that $Z_t = f(X_t, Y_t)$, we apply Itô's formula to get

$$\begin{aligned} Z_t &= Z_0 + \int_0^t \eta Z_s dX_s + \int_0^t \left(-\frac{\eta^2}{2} Z_s \right) dY_s \\ &\quad + \frac{1}{2} \int_0^t \eta^2 Z_s d[X]_s + \int_0^t \left(-\frac{\eta^3}{2} Z_s \right) d[X, Y]_s + \frac{1}{2} \int_0^t \frac{\eta^4}{4} Z_s d[Y]_s. \end{aligned}$$

The equalities $X_t = \int_0^t g(s) dW_s$ and $Y_t = \int_0^t g^2(s) ds$ can be equivalently written as $dX_t = g(t) dW_t$ and $dY_t = g^2(t) dt$. We then also have $d[X]_t = g^2(t) dt$, $d[X, Y]_t = 0$, and $d[Y]_t = 0$. We can then rewrite the above as

$$\begin{aligned} Z_t &= Z_0 + \int_0^t \eta Z_s g(s) dW_s + \int_0^t \left(-\frac{\eta^2}{2} Z_s g^2(s) \right) ds + \frac{1}{2} \int_0^t \eta^2 Z_s g^2(s) ds \\ &= Z_0 + \int_0^t \eta Z_s g(s) dW_s. \end{aligned}$$

It follows that Z is a local martingale, and in fact a true martingale, by the hint. In particular, $E[Z_t] = E[Z_0]$. Noting that $Z_0 = 1$, we get

$$1 = E[e^{-\frac{\eta^2}{2} Y_t + \eta X_t}] = e^{-\frac{\eta^2}{2} Y_t} E[e^{\eta X_t}],$$

and thus

$$E[e^{\eta X_t}] = e^{\frac{\eta^2}{2} Y_t}.$$

Thus, the moment generating function of the random variable X_t is given by $e^{\frac{\eta^2}{2} Y_t}$. Since the moment generating function determines the distribution, it follows that X_t is normally distributed with mean zero and variance $Y_t = \int_0^t g^2(s) ds$. This completes the proof.