# Sample PT 2023 ETHZ exam [Total number of points: 50]

At any point you can use results proved in the lecture without proof, unless explicitly asked for a proof. If you use a result from the lecture, please reference it appropriately.

Please pay attention to the quality, the precision and the presentation of your mathematical writing. Intermediate steps may be marked.

*Exercise 1.* [17 points] Let  $(X_n)_{n \ge 1}$  be a sequence of i.i.d. random variables following the exponential distribution of parameter 1.

- (1) Let  $(A_n)_{n\geq 1}$  be a sequence of events. Give the definition of the event  $\limsup A_n$ . State and prove the Borel-Cantelli Lemmas.
- (2) [2 points] Fix c > 1. Show that

 $\mathbb{P}(X_n > c \ln(n) \text{ for infinitely many } n) = 0.$ 

(3) [2 points] Fix  $c \in (0, 1]$  Show that

 $\mathbb{P}(X_n > c \ln(n) \text{ for infinitely many } n) = 1.$ 

(4) [3 points] Fix c > 0. Compute, with justification, the quantity

 $\mathbb{P}(X_n \leq c \ln(n) \text{ for infinitely many } n).$ 

#### Solution:

(1) Let  $(A_n)_{n\geq 1}$  be a sequence of events. We write

$$\limsup A_n = \bigcap_{n \ge 0} \bigcup_{k \ge n} A_k.$$

**[1 points] giving the definition of**  $\limsup A_n$ **Borel-Cantelli 1.** If  $\sum_{n=1}^{\infty} \mathbb{P}(A_n) < \infty$ , then  $\mathbb{P}(\limsup A_n) = 0$ . **[1 point] correct statement** 

*Proof.* For  $n \ge 1$ ,  $\limsup A_n \subset \bigcup_{k \ge n} A_k$ , so by monotonicity

$$\mathbb{P}(\limsup A_n) \le \mathbb{P}\left(\bigcup_{k=n} A_k\right) \le \sum_{k=n}^{\infty} \mathbb{P}(A_k) \quad \underset{n \to \infty}{\longrightarrow} \quad \text{o}$$

[1 point] justification of the first inequalityas the remainder of a convergent series.[1 point] justification of why  $\sum_{k=n}^{\infty} \mathbb{P}(A_k) \rightarrow 0$ 

**Borel-Cantelli 2.** If  $\sum_{n=1}^{\infty} \mathbb{P}(A_n) = \infty$  and if  $(A_n)_{n \ge 1}$  are independent, then  $\mathbb{P}(\limsup A_n) = 1$ . [1 point] correct statement] *Proof.* Fix  $n \ge \ell \ge 1$  and write

$$\mathbb{P}\left(\bigcap_{k=\ell}^{n} A_{k}^{c}\right) = \prod_{k=\ell}^{n} \mathbb{P}\left(A_{k}^{c}\right) = \prod_{k=\ell}^{n} (1 - \mathbb{P}(A_{k}))$$

## [1 point] for these equalities with justification

where we have used independence for the first equality. Using the inequality  $\ln(1-x) \le -x$  valid for  $0 \le x \le 1$ , we get

$$\prod_{k=\ell}^{n} (\mathbf{1} - \mathbb{P}(A_k)) = \exp\left(\sum_{k=\ell}^{n} \ln(\mathbf{1} - \mathbb{P}(A_k))\right) \le \exp\left(-\sum_{k=\ell}^{n} \mathbb{P}(A_k)\right).$$

[1 point] the last inequality with justification But  $\sum_{k=l}^{n} \mathbb{P}(A_k) \to \infty$  as  $n \to \infty$ , so

$$\prod_{k=\ell}^{n} (\mathbf{1} - \mathbb{P}(A_k)) \xrightarrow[n \to \infty]{} \mathbf{0}$$

[1 point] this convergence  $\rightarrow 0$  with justification But

$$\mathbb{P}\left(\bigcap_{k=\ell}^{n} A_{k}^{c}\right) \xrightarrow[n \to \infty]{} \mathbb{P}\left(\bigcap_{k=\ell}^{\infty} A_{k}^{c}\right)$$

as a decreasing sequence of events. Thus

$$\mathbb{P}\left(\bigcap_{k=\ell}^{\infty}A_{k}^{c}\right) = \mathbf{o}.$$

Thus

$$\mathbb{P}\left(\bigcup_{\ell=1}^{\infty}\bigcap_{k=\ell}^{\infty}A_{k}^{c}\right)=\mathrm{o}$$

[1 point] this probability = 0 with justification

$$\mathbb{P}(\limsup A_n) = \mathbb{P}\left(\bigcap_{\ell=1}^{\infty} \bigcup_{k=\ell}^{\infty} A_k\right) = 1 - \mathbb{P}\left(\bigcup_{\ell=1}^{\infty} \bigcap_{k=\ell}^{\infty} A_k^c\right) = 1.$$

[1 point] conclusion by complementation

(2) We have  $\mathbb{P}(X_n \ge a) = e^{-a}$  for every  $a \ge 0$ [1 point] for this expression , so  $\mathbb{P}(X_n > c \ln(n)) = e^{-c \ln(n)} = \frac{1}{n^c}.$ Thus, for c > 1, we have  $\sum_{n=1}^{\infty} \frac{1}{n^c} < \infty$ . [1 point] for the convergence of this sum The conclusion follows from the first Borel-Cantelli Lemma, since  $\limsup A_n$  is the event " $A_n$ occurs for infinitely many n". (3) For  $c \in (0, 1]$ , we have  $\sum_{n=1}^{\infty} \frac{1}{n^c} = \infty$ . [1 point] for the divergence of this sum Since the events  $\{X_n > c \ln(n)\}$  are independent, the conclusion follows from the second Borel-Cantelli Lemma. [1 point] for stating the independence hypothesis (4) We have  $\mathbb{P}(X_n \le c \ln(n)) = 1 - 1/n^c \to 1$ . So  $\sum_{n=1}^{\infty} \mathbb{P}(X_n \le c \ln(n)) = \infty$ . [1 point] for the divergence of this sum Since the events  $\{X_n > c \ln(n)\}$  are independent, by the second Borel-Cantelli Lemma we get  $\mathbb{P}(X_n \le c \ln(n) \text{ for infinitely many } n) = 1.$ [1 point] for the independence hypothesis

[1 point] for using the Borel-Cantelli Lemma

*Exercise 2.* [14 points] Let  $(X_n)_{n \ge 1}$  be a sequence of independent random variables which follow the uniform distribution on [0, 1]. Set  $Y_n = (X_n)^n$ .

- (1) [4 points] State and prove the transfer theorem.
- (2) [1 point] Compute (with justification)  $\mathbb{E}[X_1]$ .
- (3) **[2 points]** Let  $F : \mathbb{R} \to \mathbb{R}_+$  be measurable. Using the transfer theorem, write  $\mathbb{E}[F(Y_n)]$  as an integral on [0, 1] with respect to the Lebesgue measure. Please write explicitly with what function and what random variable you apply the transfer theorem with.
- (4) **[2 points]** Using the dummy function method, deduce that  $Y_n$  is a random variable with a density, and give an expression of this density.
- (5) [1 **point**] Show that  $Y_n$  converges in probability to 0 as  $n \to \infty$ .
- (6) [1 **point**] Show that  $Y_n$  converges in  $L^1$  as  $n \to \infty$ .
- (7) [3 points] Does  $Y_n$  converge almost surely as  $n \to \infty$ ? Justify your answer.

### Solution:

(1) **Transfer theorem.** Let  $X : \Omega \to E$  be a random variable and  $f : E \to \mathbb{R}_+$  a measurable function. Then  $\mathbb{E}[f(X)] = \int_E f(x) \mathbb{P}_X(dx)$  where  $\mathbb{P}_X$  is the law of *X*.

[1 points] correct statement  $Proof. \text{ Step 1. Take } f = \mathbb{1}_A \text{ with } A \in \mathcal{E}. \text{ Then } \mathbb{E}[\mathbb{1}_A(X)] = \mathbb{E}[\mathbb{1}_{X \in A}] = \mathbb{P}(X \in A) \text{ and } \int_E \mathbb{1}_A(x) \mathbb{P}_X(dx) = \mathbb{P}_X(A) = \mathbb{P}(X \in A).$ 

[1 point] checking for indicators

Step 2. By linearity, the result is true for any nonnegative simple function.

[1 point] getting the result for nonnegative simple functions We then take a sequence  $(f_n)$  of simple functions such that  $0 \le f_n \le f$  and  $f_n \uparrow f$ . By step 1:

$$\mathbb{E}[f_n(X)] = \int_E f_n(x) \mathbb{P}_X(\mathrm{d}x)$$

and by monotone convergence (twice)

$$\mathbb{E}[f_n(X)] = \int_{\Omega} f_n(X(\omega)) \mathbb{P}(\mathrm{d}\omega) \quad \xrightarrow[n \to \infty]{} \quad \int_{\Omega} f(X(\omega)) \mathbb{P}(\mathrm{d}\omega) = \mathbb{E}[f(X))]$$

and

$$\int_E f_n(x) \mathbb{P}_X(\mathrm{d} x) \quad \xrightarrow[n \to \infty]{} \quad \int_E f(x) \mathbb{P}_X(\mathrm{d} x).$$

[1 point] Conclusion by monotone convergence

(2) Using the transfer theorem with the function f(x) = x for  $x \in [0, 1]$ , we have

$$\mathbb{E}[X_1] = \int_0^1 x \mathrm{d}x = \frac{1}{2}.$$

[1 point] obtaining the result with justification using the transfer theorem

(3) We apply the transfer theorem with the random variable  $X_n$  and the function  $f(x) = F(x^n)$  to get

$$\mathbb{E}[F(Y_n)] = \mathbb{E}[F(X_n^n)] = \int_0^1 F(x^n) dx.$$

[1 point] obtaining the result with justification using the transfer theorem

(4) We perfom the change of variables  $y = x^n$ , which gives  $x = y^{1/n}$ , so  $dx = \frac{1}{n}y^{1/n-1}dy$ :

$$\mathbb{E}\left[F(Y_n)\right] = \int_0^1 F(x^n) \mathrm{d}x = \int_0^1 F(y) \frac{1}{n} y^{1/n-1} \mathrm{d}y.$$

We conclude that  $Y_n$  has density  $\frac{1}{n}y^{1/n-1}$  on [0, 1]. [1 point] attempting a change of variables

[1 point] correct result

- (5) For  $\varepsilon \in (0, 1)$ ,  $\mathbb{P}(|Y_n| \ge \varepsilon) = 1 \varepsilon^{1/n} \to 0$  when  $n \to \infty$ , which gives the result. **[1 point]**
- (6) using the transfer theorem, since  $|Y_n| = Y_n$  we get

$$\mathbb{E}[Y_n] = \int_0^1 x \frac{x^{1/n-1}}{n} dx = \frac{1}{n} \int_0^1 x^{1/n} dx = \frac{1}{n} \frac{1}{1+1/n} \quad x \to \infty \quad 0.$$

Thus  $Y_n$  converges in  $L^1$  to o. [1 point] for the correct computation (7) Fix any  $\varepsilon \in ]0, 1[$ . Using the Taylor expansion  $\exp(x) = 1 + x + o(x)$ , which implies that  $1 - \exp(x) \sim -x$  as  $x \to 0$ , we have

$$\mathbb{P}(Y_n \ge \varepsilon) = \mathbf{1} - \varepsilon^{1/n} = \mathbf{1} - \exp\left(\frac{1}{n}\ln(\varepsilon)\right) \quad \underset{n \to \infty}{\sim} \quad -\frac{1}{n}\ln(\varepsilon),$$

so that

$$\sum_{n\geq 1} \mathbb{P}\left(Y_n \geq \varepsilon\right) = \infty$$

Since the events  $\{Y_n \ge \varepsilon\}$  are independent, by the second Borel-Cantelli lemma almost surely  $Y_n \ge \varepsilon$  infinitely often.

[1 points] for application of Borel–Cantelli 2 with  $\{Y_n \ge \varepsilon\}$ 

Similarly,  $\mathbb{P}(Y_n < \varepsilon) = \exp(\frac{1}{n}\ln(\varepsilon)) \rightarrow 1$ , so

$$\sum_{n\geq 1} \mathbb{P}\left(Y_n \leq \varepsilon\right) = \infty$$

Since the events  $\{Y_n \le \varepsilon\}$  are independent, by the second Borel-Cantelli lemma almost surely  $Y_n \le \varepsilon$  infinitely often.

[1 point] for application of Borel–Cantelli 2 with  $\{Y_n < \varepsilon\}$ 

Thus, almost surely  $Y_n \ge \varepsilon$  infinitely often and  $Y_n \le \varepsilon$  infinitely often. This shows that a.s.  $(Y_n)$  diverges.

[1 point] for the conclusion

*Exercise 3.* [9 points] Let  $\Omega$  be a set and let  $\mathcal{A}$  be a  $\sigma$ -field on  $\Omega$ . Let H be a set of functions from  $\Omega$  to  $\mathbb{R}$  which satisfies the following two properties:

- *H* contains all constant functions and is stable under increasing limits (that is if  $f : E \to \mathbb{R}$  is a function with  $f = \lim f_n$  with  $(f_n)_{n \ge 1}$  is a sequence of elements of *H* such that  $f_n(\omega) \le f_{n+1}(\omega)$  for every  $\omega \in \Omega$  and  $n \ge 0$ , then  $f \in H$ ).
- *H* is a vector space (that is, if  $a, b \in \mathbb{R}$  and  $f, g \in H$  then  $af + bg \in H$ ).
- (1) [2 points] State the Dynkin Lemma.
- (2) [3 points] Show that  $\mathcal{B} = \{A \in \mathcal{A} : \mathbb{1}_A \in H\}$  is a Dynkin system.
- (3) **[4 points]** Let  $C \subset A$  be a generating  $\pi$ -system of A. Assume that  $\mathbb{1}_A \in H$  for every  $A \in C$ . Show that H contains all A-measurable real-valued functions.

Hint. First use the Dynkin Lemma to show that *H* contains all functions of the form  $\mathbb{1}_A$  with  $A \in \mathcal{A}$ .

### Solution:

(1) **Dynkin Lemma.** Let  $\Omega$  be a set and let  $C \subset \mathcal{P}(\Omega)$  be a collection of subsets of  $\Omega$ . Assume that C is stable by finite intersections. Then the Dynkin system generated by C is equal to the  $\sigma$ -field generated by C.

[2 points] for the correct statement.

- (2) We check the three properties defining a Dynkin system:
  - (a)  $\Omega \in \mathcal{B}$  because the constant function equal to 1 is in *H*.
  - (b) If  $A \in \mathcal{B}$ , then  $\mathbb{1}_{A^c} = 1 \mathbb{1}_A$  because 1 and  $\mathbb{1}_A$  are in *H* and *H* is a vector space.
  - (c) If  $(A_n)_{n \ge 1}$  is a pairwise disjoint sequence of elements of  $\mathcal{B}$ , set  $A = \bigcup_{n \ge 1} A_n$  and then observe that

$$\mathbb{1}_A = \lim_{n \to \infty} \sum_{k=1}^n \mathbb{1}_{A_k},$$

where the limit is increasing. In addition  $\mathbb{1}_{A_k} \in H$  since H is a vector space. Since H is stable under increasing limits, we conclude that  $\mathbb{1}_A$  is in H

[4 points] 1 point for (a), 1 point for (b), 2 points for (c)

(3) *Step 1.* By question (2),  $\mathcal{B}$  is a Dynkin system containing  $\mathcal{C}$ . Therefor  $\mathcal{B}$  contains the Dynkin system generated by  $\mathcal{C}$ . Since  $\mathcal{C}$  is a  $\pi$ -system, by the Dynkin Lemma the Dynkin system generated by  $\mathcal{C}$  is  $\sigma(\mathcal{C})$ , which is  $\mathcal{A}$ . We conclude that  $\mathcal{B} = \mathcal{A}$ , so that H contains all functions of the form  $\mathbb{1}_A$ 

with  $A \in \mathcal{A}$ . [1 point] for step 1

Step 2. Since *H* is a vector space, by linearity *H* contains all simple A-measurable functions. [1 point] for step 2

Step 3. H contains all nonnegative A-measurable functions, since every nonnegative A-measurable function is an increasing limit of A-measurable simple functions and H is stable under increasing limits. [1 point] for step 3

Step 4. *H* contains all A-measurable functions, since every A-measurable function can be written as a difference of two nonngative A-measurable functions and *H* is a vector space.

[1 point] for step 4

*Exercise 4.* **10 points** Fix an integer  $n \ge 2$  and let  $(U_k)_{1 \le k \le n}$  be independent random variables, all following the uniform distribution on [0, 1]. Define

$$M_n = \max\left(\frac{1}{\sqrt{U_1}}, \dots, \frac{1}{\sqrt{U_n}}\right).$$

- (1) [3 points] Compute the cumulative distribution function of  $M_n$ .
- (2) **[2 points]** Show that  $x^2 \mathbb{P}(M_n \ge x) \to n \text{ as } x \to \infty$ .
- (3) **[2 points]** Let X be a nonnegative real-valued random variable. Show that  $\mathbb{E}[X] = \int_{0}^{\infty} \mathbb{P}(X \ge u) du$ . Note. This question is independent of questions (1) and (2).
- (4) **[3 points]** For what values of p > 0 do we have  $\mathbb{E}[M_n^p] < \infty$ ?

Note. You may use the results of the previous questions even if you didn't manage to solve them.

#### Solution:

(1) For  $u \in [0, 1]$ , we have  $\frac{1}{\sqrt{u}} \ge 1$ . As a consequence for x < 1 we have  $\mathbb{P}(M_n \le x) = 0$ . **[1 point for the case** x < 1]

For  $x \ge 1$  we have:

$$\mathbb{P}(M_n \le x) = \mathbb{P}\left(\frac{1}{\sqrt{U_1}} \le x, \cdots, \frac{1}{\sqrt{U_n}} \le x\right) = \mathbb{P}\left(U_1 \ge \frac{1}{x^2}\right) \cdots \mathbb{P}\left(U_n \ge \frac{1}{x^2}\right) = \mathbb{P}\left(U_1 \ge \frac{1}{x^2}\right)^n = \left(1 - \frac{1}{x^2}\right)^n.$$

The second equality follows from independence and the third one from the fact that  $(U_i)_{1 \le i \le n}$  have same law.

[1 point] for the second equality

[1 point] for the the final result

(2) By (1), for  $x \ge 1$  we have

$$\mathbb{P}(M_n \ge x) = 1 - \mathbb{P}(M_n < x) = 1 - \left(1 - \frac{1}{x^2}\right)^n.$$

By the Binomial formula,

$$\left(1 - \frac{1}{x^2}\right)^n = \sum_{k=0}^n \binom{n}{k} \frac{(-1)^k}{x^{2k}},$$

so

$$x^{2}\mathbb{P}(M_{n} \ge x) = x^{2}\left(1 - \left(1 - \frac{1}{x^{2}}\right)^{n}\right) = \sum_{k=1}^{n} \binom{n}{k} \frac{(-1)^{k+1}}{x^{2k-2}} = n + \sum_{k=2}^{n} \binom{n}{k} \frac{(-1)^{k+1}}{x^{2k-2}} \xrightarrow{x \to \infty} n$$

