

# Chapter 4: Part 1, Conditional expectation

- Outline:
- 1) The discrete setting
  - 2) Definition and first properties
  - 3) Non-negative r.v.
  - 4) Convergence theorems
  - 5) Some useful properties
  - 6) Conditional density functions

Intuitively speaking, the goal is to see how the knowledge of information (i.e. a  $\sigma$ -field) modifies probability measures. Here we shall define the conditional expectation of random variables given a  $\sigma$ -field.

## 1) The discrete setting

We first condition with respect to an event.

Let  $(\Omega, \mathcal{F}, \mathbb{P})$  be a probability space and  $B \in \mathcal{F}$  with  $\mathbb{P}(B) > 0$ . We can define  $\mathbb{P}(\cdot|B)$  the so-called conditional probability given  $B$  by  $\mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)}$  for every  $A \in \mathcal{F}$ .

Similarly, for  $X \in \mathcal{L}^1$ , we define  $\mathbb{E}[X|B] = \frac{\mathbb{E}[X \mathbb{1}_B]}{\mathbb{P}(B)}$ .

Interpretation: it is the average value of  $X$  when  $B$  occurs.

The notation  $\mathbb{E}[X|B]$  comes from the following fact:

Lemma Let  $X: (\Omega, \mathcal{F}) \rightarrow \mathbb{R}_+$  be a random variable. Its expectation with respect to  $\mathbb{P}(\cdot|B)$  is  $\frac{\mathbb{E}[X \mathbb{1}_B]}{\mathbb{P}(B)}$ .

If we set  $\tilde{\mathbb{P}}(A) = \mathbb{P}(A|B)$  for  $A \in \mathcal{F}$ , the expectation of  $X$  with respect to  $\mathbb{P}(\cdot|B)$  is by definition the integral of  $X$  with respect to  $\tilde{\mathbb{P}}$ , that is the quantity  $\int_{\Omega} X(\omega) \tilde{\mathbb{P}}(d\omega)$ .

Proof We show that  $\int_{\Omega} X(\omega) \tilde{\mathbb{P}}(d\omega) \stackrel{(*)}{=} \frac{\mathbb{E}[X \mathbb{1}_B]}{\mathbb{P}(B)}$

Step 1 for  $X = \mathbb{1}_A$ ,  $\int_{\Omega} \mathbb{1}_A(\omega) \tilde{\mathbb{P}}(d\omega) = \tilde{\mathbb{P}}(A) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)}$  which is  $\frac{\mathbb{E}[\mathbb{1}_A \mathbb{1}_B]}{\mathbb{P}(B)}$

Step 2 By linearity  $(*)$  holds for every simple random variable (a r.v. taking a finite number of values)

Step 3 Let  $0 \leq X_n \uparrow X$  with  $X_n$  simple. Then

$$\int_{\Omega} X(\omega) \tilde{\mathbb{P}}(d\omega) \xleftarrow{n \rightarrow \infty} \int_{\Omega} X_n(\omega) \tilde{\mathbb{P}}(d\omega) = \frac{\mathbb{E}[X_n \mathbb{1}_B]}{\mathbb{P}(B)} \xrightarrow{n \rightarrow \infty} \frac{\mathbb{E}[X \mathbb{1}_B]}{\mathbb{P}(B)}$$

by monotone convergence (twice)

**(N.B. Proof not done during the lecture)**

We shall now condition with respect to a discrete random variable

Now let  $Y: \Omega \rightarrow E$  be a r.v. with  $E$  countable. We want to define the conditional expectation of  $X$  given  $Y$ . From before, we have  $\mathbb{E}[X | Y=y] = \frac{\mathbb{E}[X \mathbb{1}_{Y=y}]}{\mathbb{P}(Y=y)}$  for every  $y$  such that  $\mathbb{P}(Y=y) > 0$ .

Thus, we naturally set:  $\mathbb{E}[X | Y] = \varphi(Y)$  where  $\varphi: E \rightarrow \mathbb{R}$  is defined by  $\varphi(y) = \begin{cases} \mathbb{E}[X | Y=y] & \text{if } \mathbb{P}(Y=y) > 0 \\ 0 & \text{otherwise} \end{cases}$

In other words  $\mathbb{E}[X | Y]$  is a random variable defined by

$$\mathbb{E}[X | Y](\omega) = \varphi(Y(\omega)).$$

**⚠** the notation  $\mathbb{E}[X(\omega) | Y(\omega)]$  makes no sense.

The value of  $\mathbb{P}(Y=y)$  when  $\mathbb{P}(Y=y)=0$  is arbitrary: it influences the definition of  $\mathbb{E}[X|Y]$  only on a  $\mathbb{P}$ -probability set, the set  $\{\omega \in \Omega: Y(\omega) \in E'\}$  with  $E' = \{y \in E: \mathbb{P}(Y=y)=0\}$ .

More generally, conditional expectations will always be defined up to  $\mathbb{P}$ -probability sets

Observe that  $\mathbb{E}[X|Y]$  is a r.v. which is  $\sigma(Y)$ -measurable, since it is a function of  $Y$

## END OF LECTURE 14

Example (dice toss) Take  $\Omega = \{1, \dots, 6\}$  with  $\mathbb{P}(\{\omega\}) = \frac{1}{6}$  for  $\omega \in \Omega$ . Set  $X(\omega) = \omega$  and  $Y(\omega) = \begin{cases} 1 & \text{if } \omega \text{ odd} \\ 0 & \text{if } \omega \text{ even} \end{cases}$ .  
Then  $\mathbb{E}[X|Y](\omega) = \begin{cases} 3 & \text{if } \omega \in \{1, 3, 5\} \\ 4 & \text{if } \omega \in \{2, 4, 6\} \end{cases}$

Lemma We have:

- $\mathbb{E}[|\mathbb{E}[X|Y]|] \leq \mathbb{E}[|X|]$ , so  $\mathbb{E}[X|Y] \in L^1$
- $\forall$  r.v.  $Z$  which is  $\sigma(Y)$ -measurable and bounded,  $\mathbb{E}[ZX] = \mathbb{E}[Z \mathbb{E}[X|Y]]$   
(we take  $Z$  bounded to ensure integrability)

Proof 1) We have  $\mathbb{E}[|\mathbb{E}[X|Y]|] = \sum_{y \in E} \mathbb{P}(Y=y) |\mathbb{E}[X|Y=y]| = \sum_{\substack{y \in E \\ \mathbb{P}(Y=y) > 0}} \mathbb{P}(Y=y) \frac{|\mathbb{E}[X \mathbb{1}_{Y=y}]|}{\mathbb{P}(Y=y)} \leq \sum_{\substack{y \in E \\ \mathbb{P}(Y=y) > 0}} \mathbb{E}[|X| \mathbb{1}_{Y=y}] = \mathbb{E}[|X|]$

2) By the Doob-Dynkin lemma we can write  $Z = F(Y)$ . Then similarly,  
 $\mathbb{E}[F(Y) \mathbb{E}[X|Y]] = \sum_{\substack{y \in E \\ \mathbb{P}(Y=y) > 0}} \mathbb{P}(Y=y) F(y) \mathbb{E}[X|Y=y] = \sum_{\substack{y \in E \\ \mathbb{P}(Y=y) > 0}} \mathbb{P}(Y=y) F(y) \frac{\mathbb{E}[X \mathbb{1}_{Y=y}]}{\mathbb{P}(Y=y)}$

Which is  $\mathbb{E}[X \sum_{\substack{y \in E \\ \mathbb{P}(Y=y) > 0}} F(y) \mathbb{1}_{Y=y}] = \mathbb{E}[X F(Y)]$  because a.s.  $\sum_{\substack{y \in E \\ \mathbb{P}(Y=y) > 0}} F(y) \mathbb{1}_{Y=y} = F(Y)$ .

It turns out that this property can be used to define conditional expectations in a general setting

## 2) Definition and first properties

Let  $(\Omega, \mathcal{F}, \mathbb{P})$  be a probability space. If  $\mathcal{A} \subset \mathcal{F}$  is a  $\sigma$ -field, we write  $X \in L^1(\Omega, \mathcal{A}, \mathbb{P})$  if

- $X: (\Omega, \mathcal{A}) \rightarrow (\mathbb{R}, \mathcal{B}(\mathbb{R}))$  is measurable
- $\mathbb{E}[|X|] < \infty$

Theorem Fix  $X \in L^1(\Omega, \mathcal{F}, \mathbb{P})$ . Let  $\mathcal{A} \subset \mathcal{F}$  be a  $\sigma$ -field. Then there exists a random variable  $X'$  with

①  $X' \in L^1(\Omega, \mathcal{A}, \mathbb{P})$

② For every r.v.  $Z$ ,  $\mathcal{A}$ -measurable and bounded,  $\mathbb{E}[ZX] = \mathbb{E}[ZX']$

Moreover, if  $X''$  is another such random variable, then  $X' = X''$  a.s.

We denote by  $\mathbb{E}[X|\mathcal{A}]$  any such random variable, called a version of the conditional expectation of  $X$  given  $\mathcal{A}$ .

Property ② is called the "characteristic property of conditional expectation"

We make some observations before the proof (assume for the moment that the theorem is true)

Remarks.  $\mathbb{E}[X|\mathcal{A}]$  is an  $\mathcal{A}$ -measurable random, uniquely defined up to zero probability events. In practice, this is transparent because we consider only expectations of  $\mathbb{E}[X|\mathcal{A}]$  or its almost sure properties. For this reason, we often say that " $\mathbb{E}[X|\mathcal{A}]$  is the conditional expectation of  $X$  given  $\mathcal{A}$ "

• By interpreting  $\mathbb{E}[Z(X-X')] = 0$  by  $\langle Z, X - \mathbb{E}[X|\mathcal{A}] \rangle = 0$ , we can interpret  $\mathbb{E}[X|\mathcal{A}]$  as the "projection" of  $X$  onto  $\mathcal{A}$ -measurable functions. This will be made rigorous when  $X \in L^2$

Notation • For  $B \in \mathcal{F}$ , we define  $\mathbb{P}(B|\mathcal{A}) = \mathbb{E}[\mathbb{1}_B|\mathcal{A}]$ : it is an  $\mathcal{A}$ -measurable r.v. (defined a.s.)

• If  $Y: (\Omega, \mathcal{F}) \rightarrow (\mathcal{E}, \mathcal{E})$  is a r.v., we define  $\mathbb{E}[X|Y] = \mathbb{E}[X|\sigma(Y)]$

(here  $X$  is always  $\mathbb{R}$ -valued integrable, but  $Y$  is not necessarily real-valued)

## Remark

Recall the Doob-Dynkin lemma: If  $Y$  is  $\mathbb{R}^n$ -valued, then a  $\sigma(Y)$  measurable function is of the form  $g(Y)$  with  $g$  measurable. (We did the proof for  $n=1$ , but the proof is very similar). As a consequence:

- there is a measurable function  $\varphi: \mathbb{R}^n \rightarrow \mathbb{R}$  such that  $\mathbb{E}[X|Y] = \varphi(Y)$
- If  $\varphi: \mathbb{R}^n \rightarrow \mathbb{R}$  is a measurable function such that  $\mathbb{E}[|\varphi(Y)|] < \infty$  and for every  $g$  measurable bounded  $\mathbb{E}[X g(Y)] = \mathbb{E}[\varphi(Y) g(Y)]$ , then  $\mathbb{E}[X|Y] = \varphi(Y)$

Remark the definition is consistent with what we saw in the discrete case. For example, if  $U$  is  $\mathbb{Z}$ -valued, let's find  $\mathbb{E}[X|U]$ . By the Doob-Dynkin lemma,  $\mathbb{E}[X|U] = g(U)$  and we want to find  $g$ . Taking  $Z = \mathbb{1}_{U=n}$  gives  $\mathbb{E}[X \mathbb{1}_{U=n}] = \mathbb{E}[g(U) \mathbb{1}_{U=n}] = \mathbb{E}[g(n) \mathbb{1}_{U=n}]$ . Thus  $g(n) = \frac{\mathbb{E}[X \mathbb{1}_{U=n}]}{\mathbb{P}(U=n)}$  when  $\mathbb{P}(U=n) > 0$ .

Simple properties of conditional expectation Take  $X \in L^1(\Omega, \mathcal{F}, \mathbb{P})$  and let  $\mathcal{A} \subset \mathcal{F}$  be a  $\sigma$ -field.

- ①  $\mathbb{E}[\mathbb{E}[X|\mathcal{A}]] = \mathbb{E}[X]$  (very very useful)
- ②  $\mathbb{E}[X|\{\emptyset, \Omega\}] = \mathbb{E}[X]$
- ③ If  $X$  is  $\mathcal{A}$ -measurable,  $\mathbb{E}[X|\mathcal{A}] = X$ . In particular, we always have  $\mathbb{E}[X|\mathcal{F}] = X$ .
- ④  $X \mapsto \mathbb{E}[X|\mathcal{A}]$  is linear
- ⑤  $X \geq 0 \Rightarrow \mathbb{E}[X|\mathcal{A}] \geq 0$ . As a consequence,  $X_1 \geq X_2 \Rightarrow \mathbb{E}[X_1|\mathcal{A}] \geq \mathbb{E}[X_2|\mathcal{A}]$ .
- ⑥  $|\mathbb{E}[X|\mathcal{A}]| \leq \mathbb{E}[|X|\mathcal{A}]$

NB: In every statement, the "almost sure" is implicit (recall that conditional expectations are defined uniquely almost surely)

Proof ① Just take  $Z = \mathbb{1}_\Omega$  in the characteristic property

② A  $\{\emptyset, \Omega\}$ -measurable function is constant, so  $\mathbb{E}[X|\{\emptyset, \Omega\}]$  is constant, and thus equal to its expectation, which by ① is  $\mathbb{E}[X]$ .

③ Follows from uniqueness:  $X$  then satisfies ① and ②

④ If  $X_1, X_2 \in L^1(\Omega, \mathcal{F}, \mathbb{P})$ ,  $X = \alpha \mathbb{E}[X_1|\mathcal{A}] + \beta \mathbb{E}[X_2|\mathcal{A}]$  satisfies

for every  $Z: \Omega \rightarrow \mathbb{R}$ ,  $A$ -measurable, bounded,  $\mathbb{E}[XZ] = \mathbb{E}[(\alpha X_1 + \beta X_2)Z]$ .

so  $\mathbb{E}[aX_1 + bX_2 | A] = a\mathbb{E}[X_1 | A] + b\mathbb{E}[X_2 | A]$

⑤ Fix  $\varepsilon > 0$ . Take  $Z = \mathbb{1}_{\mathbb{E}[X|A] \leq -\varepsilon}$ ,  $A$ -measurable. Then

$$0 \leq \mathbb{E}[X \mathbb{1}_{\mathbb{E}[X|A] \leq -\varepsilon}] = \mathbb{E}[\mathbb{E}[X|A] \mathbb{1}_{\mathbb{E}[X|A] \leq -\varepsilon}] \leq -\varepsilon \mathbb{P}(\mathbb{E}[X|A] \leq -\varepsilon)$$

Hence  $\mathbb{P}(\mathbb{E}[X|A] \leq -\varepsilon) = 0$ , so a.s.  $\mathbb{E}[X|A] > -\varepsilon$ .

Thus  $\forall n \geq 1$ , a.s.  $\mathbb{E}[X|A] > -\frac{1}{n}$ . By interchanging "a.s." and " $\forall \varepsilon$  countable set":

a.s.  $\forall n \geq 1$ ,  $\mathbb{E}[X|A] > -\frac{1}{n}$ . Thus a.s.  $\mathbb{E}[X|A] \geq 0$ .

⑥ Write  $X = X^+ - X^-$  with  $X^+, X^- \geq 0$ . Then by ①

$$|\mathbb{E}[X|A]| = |\mathbb{E}[X^+|A] - \mathbb{E}[X^-|A]| \leq \mathbb{E}[X^+|A] + \mathbb{E}[X^-|A] = \mathbb{E}[X^+ + X^-|A] = \mathbb{E}[|X||A]$$

## END OF LECTURE 15

### Proof of the theorem

Uniqueness Assume that  $X'$  and  $X''$  satisfy ① and ② above. Take  $Z = \mathbb{1}_{\varepsilon X' > X''}$ ,  $A$ -measurable.

$$\text{Then } \mathbb{E}[(X' - X'') \mathbb{1}_{\varepsilon X' > X''}] = \mathbb{E}[X' \mathbb{1}_{\varepsilon X' > X''}] - \mathbb{E}[X'' \mathbb{1}_{\varepsilon X' > X''}] = \mathbb{E}[X' \mathbb{1}_{\varepsilon X' > X''}] - \mathbb{E}[X' \mathbb{1}_{\varepsilon X' > X''}] = 0$$

But a.s.  $(X' - X'') \mathbb{1}_{\varepsilon X' > X''} > 0$ . Thus a.s.  $(X' - X'') \mathbb{1}_{\varepsilon X' > X''} = 0$ . Thus  $\mathbb{P}(\varepsilon X' > X'') = 0$ , so a.s.  $X' \leq X''$ . By symmetry,  $X'' \leq X'$  a.s., so  $X' = X''$  a.s.

existence There are essentially two approaches using results from measure theory: one with Radon-Nikodym, one with Hilbert  $L^2$  spaces. We present the second approach and underline in blue facts from measure theory that we do not prove.

Step 1: Assume  $X \in L^2(\Omega, \mathcal{A}, \mathbb{P})$ . We can equip  $L^2(\Omega, \mathcal{A}, \mathbb{P})$  with a scalar product: for  $Y, Z \in L^2(\Omega, \mathcal{A}, \mathbb{P})$ , we set  $\langle Y, Z \rangle = \mathbb{E}[YZ]$ , so that the norm  $\|\cdot\|$  defined by  $\|Y\|^2 = \langle Y, Y \rangle$  defines a normed vector space (here we identify two r.v. which are equal a.s.), which is complete:  $L^2(\Omega, \mathcal{A}, \mathbb{P})$  is a Hilbert space. Then  $L^2(\Omega, \mathcal{A}, \mathbb{P})$  is a closed subset of  $L^2(\Omega, \mathcal{A}, \mathbb{P})$

so we can consider the orthogonal projection  $X'$  of  $X$  on  $L^2(\Omega, \mathcal{A}, \mathbb{P})$  characterized by the fact that  $\forall Z \in L^2(\Omega, \mathcal{A}, \mathbb{P})$ ,  $\langle X - X', Z \rangle = 0$ , which implies

$\mathbb{E}[XZ] = \mathbb{E}[X'Z]$ , and gives the result (since  $Z$   $A$ -measurable, bounded implies  $Z \in L^2(\Omega, \mathcal{A}, \mathbb{P})$ )

Step 2: Assume  $X \in L^1(\mathcal{F}, \mathcal{A}, \mathbb{P})$  and  $X \geq 0$ . We use a truncation argument:

for  $n \geq 1$ , set  $X_n = X \mathbb{1}_{X \leq n} \in L^1(\mathcal{F}, \mathcal{A}, \mathbb{P})$ . Set  $X'_n = \mathbb{E}[X_n | \mathcal{A}]$ . Since  $X_n \leq X_{n+1}$  we have  $X'_n \leq X'_{n+1}$  a.s. by property ③ above. By monotonicity, we can define

$X' = \lim_{n \rightarrow \infty} X'_n$  as an a.s. limit. Then for  $Z \geq 0$   $\mathcal{A}$ -measurable bounded, from  $\mathbb{E}[X'_n Z] = \mathbb{E}[X_n Z]$  we get  $\mathbb{E}[X' Z] = \mathbb{E}[X Z]$  by monotone convergence. When  $Z$  is not necessarily  $\geq 0$ , we write  $Z = Z^+ - Z^-$  and apply (\*) with  $Z^+$  and  $Z^-$  and get (\*\*) with  $Z$  by linearity.

By taking  $Z = 1$  we get  $\mathbb{E}[X'] = \mathbb{E}[X] < \infty$ .

Step 3 When  $X \in L^1(\mathcal{F}, \mathcal{A}, \mathbb{P})$ , write  $X = X^+ - X^-$ . We then set  $X' = \mathbb{E}[X^+ | \mathcal{A}] - \mathbb{E}[X^- | \mathcal{A}]$ , which satisfies the requirements

① and ②



### 3) Conditional expectation of non-negative r.v.

We can actually define conditional expectation for r.v. in  $[0, \infty]$  without integrability conditions

Theorem Fix a r.v.  $X$   $[0, \infty]$ -valued. Let  $\mathcal{A} \subset \mathcal{F}$  be a  $\sigma$ -field. Then there exists a random variable  $X'$  with

①  $X' \in [0, \infty]$ ,  $X'$  is  $\mathcal{A}$ -measurable

② For every r.v.  $Z$ ,  $\mathcal{A}$ -measurable and  $Z \geq 0$ ,  $\mathbb{E}[Z X] = \mathbb{E}[Z X']$

Moreover, if  $X''$  is another such random variable, then  $X' = X''$  a.s.

We denote by  $\mathbb{E}[X | \mathcal{A}]$  any such random variable, called a version of the conditional expectation of  $X$  given  $\mathcal{A}$ .

Proof: Existence As in step 2 above, if we define  $X_n = X \mathbb{1}_{X \leq n}$  and  $X'_n = \mathbb{E}[X_n | \mathcal{A}]$ ,  $X'$ , the a.s. limit of  $(X'_n)_{n \geq 1}$ , satisfies ① and ②

Uniqueness Assume that  $X'$  and  $X''$  satisfy ① and ②. Then for  $a, b \in \mathbb{Q}_+$ , by taking  $Z = \mathbb{1}_{\{X' \leq a < b \leq X''\}}$  we get

$$b \mathbb{P}(X' \leq a < b \leq X'') = \mathbb{E}[b Z] \leq \mathbb{E}[X'' Z] = \mathbb{E}[X' Z] \leq a \mathbb{P}(X' \leq a < b \leq X'')$$

Since  $a < b$ , we get  $\mathbb{P}(X' \leq a < b \leq X'') = 0$ . By taking a countable union



we get  $P(X' < X'') = 0$ . Thus  $P(X' \geq X'') = 1$ . By symmetry  $P(X' \leq X'') = 1$ ,  
 so  $P(X' = X'') = 1$



As for conditional expectations in  $L^1$ , we have the following properties which immediately follow from the definition:

Simple properties of conditional expectation Take  $X \in [0, \infty]$  valued r.v. and let  $\mathcal{A} \subset \mathcal{F}$  be a  $\sigma$ -field.

- ①  $E[X | \mathcal{F}, \mathcal{R}] = E[X]$ ,  $E[X | \mathcal{A}] = X$
- ② If  $X$  is  $\mathcal{A}$ -measurable,  $E[X | \mathcal{A}] = X$ .
- ③ if  $\alpha \in [0, \infty]$ ,  $\beta \geq 0$ ,  $E[\alpha X + \beta Y | \mathcal{A}] = \alpha E[X | \mathcal{A}] + \beta E[Y | \mathcal{A}]$
- ④  $E[E[X | \mathcal{A}]] = E[X]$  (Very very useful)
- ⑤  $X_1 \geq X_2 \geq 0 \Rightarrow E[X_1 | \mathcal{A}] \geq E[X_2 | \mathcal{A}]$ .

## 4) Convergence Theorems

Theorem Let  $\mathcal{A} \subset \mathcal{F}$  be a  $\sigma$ -field

① [Conditional monotone convergence] Let  $(X_n)_{n \geq 0}$  be an increasing sequence of  $[0, \infty]$  r.v. with  $X = \lim_{n \rightarrow \infty} \uparrow X_n$

Then  $E[X_n | \mathcal{A}] \xrightarrow{n \rightarrow \infty} E[X | \mathcal{A}]$

② [Conditional Fatou] Let  $(X_n)_{n \geq 0}$  be a sequence of  $[0, \infty]$  r.v. Then  $E[\liminf_{n \rightarrow \infty} X_n | \mathcal{A}] \leq \liminf_{n \rightarrow \infty} E[X_n | \mathcal{A}]$

③ [Conditional dominated convergence] Let  $(X_n)_{n \geq 1}$  be a sequence of integrable r.v. such that

- $X_n \xrightarrow{a.s.} X$
- $\exists Y \geq 0$  in  $L^1$  s.t.  $\forall n \geq 1$ , a.s.  $|X_n| \leq Y$

Then  $E[X_n | \mathcal{A}] \xrightarrow{n \rightarrow \infty} E[X | \mathcal{A}]$  almost surely and in  $L^1(\mathcal{R}, \mathcal{A}, P)$

④ [Conditional Jensen] Let  $f: \mathbb{R} \rightarrow \mathbb{R}_+$  be a convex function (if  $f'' \geq 0$ ,  $f$  is convex) such that  $X \in L^1(\mathcal{R}, \mathcal{A}, P)$

Then  $f(E[X | \mathcal{A}]) \leq E[f(X) | \mathcal{A}]$

By taking  $\mathcal{A} = \mathcal{F}, \mathcal{R}$  we get  $f(E[X]) \leq E[f(X)]$  (Jensen inequality)

(as usual for conditional expectations, all statements are to be understood almost surely)



Proof ① We know that  $\mathbb{E}[X_n | \mathcal{A}]$  is  $\mathcal{A}$ -measurable as a limit of  $\mathcal{A}$ -measurable functions.

Then for  $z \geq 0$   $\mathcal{A}$ -measurable,  $\mathbb{E}[X_n z] = \mathbb{E}[\mathbb{E}[X_n | \mathcal{A}] z]$ , so by monotone convergence we get  $\mathbb{E}[X z] = \mathbb{E}[X' z]$ . Thus  $X' = \mathbb{E}[X | \mathcal{A}]$  a.s.

② Observe that for  $i \geq n > 1$

$$\mathbb{E}[\inf_{k \geq n} X_k | \mathcal{A}] \leq \mathbb{E}[X_i | \mathcal{A}].$$

$$\text{Thus } \mathbb{E}[\inf_{k \geq n} X_k | \mathcal{A}] \leq \inf_{k \geq n} \mathbb{E}[X_k | \mathcal{A}].$$

But  $\inf_{k \geq n} X_k \uparrow \liminf_{n \rightarrow \infty} X_n$ , so the result follows from ①

③ Apply ① with  $z = X_n \geq 0$  and  $z = -X_n \geq 0$ .

$$\mathbb{E}[z | \mathcal{A}] - \mathbb{E}[X | \mathcal{A}] = \mathbb{E}[\liminf_{n \rightarrow \infty} (z - X_n) | \mathcal{A}] \leq \mathbb{E}[z | \mathcal{A}] - \limsup_{n \rightarrow \infty} \mathbb{E}[X_n | \mathcal{A}]$$

$$\mathbb{E}[z | \mathcal{A}] + \mathbb{E}[X | \mathcal{A}] = \mathbb{E}[\liminf_{n \rightarrow \infty} (z + X_n) | \mathcal{A}] \leq \mathbb{E}[z | \mathcal{A}] + \limsup_{n \rightarrow \infty} \mathbb{E}[X_n | \mathcal{A}]$$

Thus  $\mathbb{E}[X | \mathcal{A}] \leq \liminf_{n \rightarrow \infty} \mathbb{E}[X_n | \mathcal{A}] \leq \limsup_{n \rightarrow \infty} \mathbb{E}[X_n | \mathcal{A}] \leq \mathbb{E}[X | \mathcal{A}]$  which gives the a.s. convergence.

The  $L^1$  convergence follows from dominated convergence, since  $|\mathbb{E}[X_n | \mathcal{A}]| \leq \mathbb{E}[|X_n| | \mathcal{A}] \leq \mathbb{E}[Y | \mathcal{A}]$  which is in  $L^1$  sense  $\mathbb{E}[\mathbb{E}[Y | \mathcal{A}]] = \mathbb{E}[Y] < \infty$

④ Set  $E_f = \{(a, b) \in \mathbb{R}^2 : \forall x \in \mathbb{R}, f(x) \geq ax + b\}$ . Since  $f$  is convex,

$\forall z \in \mathbb{R}, f(z) = \sup_{(a, b) \in E_f} (az + b)$  (\*). In addition, we can find  $D \subset E_f$  countable and dense in  $E_f$  (\*\*)

$$\text{Thus } \mathbb{E}[f(X) | \mathcal{A}] = \mathbb{E}\left[\sup_{(a, b) \in D} (aX + b) | \mathcal{A}\right] \geq \sup_{(a, b) \in D} \mathbb{E}[aX + b | \mathcal{A}] = \sup_{(a, b) \in D} \underbrace{\mathbb{E}[aX | \mathcal{A}] + b}_{= f(\mathbb{E}[X | \mathcal{A}])}$$

(here we used countability because conditional expectations are only defined a.s.)



Justification of (\*) For  $x \in \mathbb{R}$  we clearly have  $f(x) \geq \sup_{(a, b) \in E_f} (ax + b)$ . The other inequality comes from the fact that by convexity of  $f$  we can find  $a, b \in \mathbb{R}$  such that  $f(x) = ax + b$  and  $f(y) \geq ay + b \forall y \in \mathbb{R}$ .

Justification of (\*\*) Denote by  $\{z_n : n \geq 1\}$  the elements of  $\mathbb{Q}^2$ . Let  $p_n \in E_f$  be such that  $|p_n - z_n| \leq d(E_f, z_n) + \frac{1}{n}$  where  $d(E_f, z_n) = \inf_{p \in E_f} |p - z_n|$ . Let us check that  $\{p_n : n \geq 1\}$  is dense in  $E_f$ . For  $\varepsilon > 0$ . By density, we can find  $n \geq \frac{1}{\varepsilon}$  with  $|z_n - x| \leq \varepsilon$

then  $|p_n - x| \leq |p_n - r_n| + |r_n - x| \leq \frac{1}{n} + d(r_n, E_g) + \varepsilon \leq 2\varepsilon + d(r_n, E_g)$ .

But  $d(r_n, E_g) \leq |r_n - x| \leq \varepsilon$  (because  $x \in E_g$ ), so  $|p_n - x| \leq 3\varepsilon$ .

Remark We often use Jensen's inequality with  $f(x) = |x|^p$  with  $p \geq 1$  and with the r.v.  $|X|$ :  
 $\mathbb{E}[|X|^p] \leq \mathbb{E}[|X|^p]$  and  $\mathbb{E}[|X|^p | \mathcal{A}] \leq \mathbb{E}[|X|^p | \mathcal{A}]$

### 5) Some useful properties

The "information contained in  $\mathcal{A}$ " can be factorized out of the conditional expectation. From now on by  $L^1$  we mean  $L^1(\Omega, \mathcal{F}, \mathbb{P})$ .

Proposition Let  $\mathcal{A} \subset \mathcal{F}$  be a  $\sigma$ -field. Let  $X, Y$  be real-valued r.v. such that  $X, Y \in [0, \infty]$  or  $X, XY \in L^1$ . Assume that  $Y$  is  $\mathcal{A}$ -measurable. Then  $\mathbb{E}[YX | \mathcal{A}] = Y\mathbb{E}[X | \mathcal{A}]$

Proof If  $X, Y \in [0, \infty]$  and  $Z \geq 0$  is  $\mathcal{A}$ -measurable, then  $ZY \geq 0$  is  $\mathcal{A}$ -measurable, so

$$\mathbb{E}[ZYX] = \mathbb{E}[ZY\mathbb{E}[X | \mathcal{A}]]$$

$Y\mathbb{E}[X | \mathcal{A}]$  is  $\mathcal{A}$ -measurable and satisfies the characteristic property of conditional expectation, so  $Y\mathbb{E}[X | \mathcal{A}] = \mathbb{E}[ZYX | \mathcal{A}]$ . The proof is similar when  $X, XY \in L^1$ .

One has the tower property (restricting information) **END OF LECTURE 16**

Proposition Let  $\mathcal{A}_1, \mathcal{A}_2 \subset \mathcal{F}$  be  $\sigma$ -fields. Then for  $X$  r.v. with  $X \in [0, \infty]$  or  $X \in L^1$ ,  
 $\mathbb{E}[\mathbb{E}[X | \mathcal{A}_2] | \mathcal{A}_1] = \mathbb{E}[X | \mathcal{A}_1]$

Proof Let  $Z \geq 0$  be  $\mathcal{A}_1$ -measurable bounded. By uniqueness it is enough to check that

$$\mathbb{E}[ZX] = \mathbb{E}[Z\mathbb{E}[\mathbb{E}[X | \mathcal{A}_2] | \mathcal{A}_1]]$$

To this end, write:  $\mathbb{E}[Z\mathbb{E}[\mathbb{E}[X | \mathcal{A}_2] | \mathcal{A}_1]] = \mathbb{E}[Z\mathbb{E}[X | \mathcal{A}_2]] = \mathbb{E}[ZX]$

because  $Z$  is also  $\mathcal{A}_2$ -measurable. The proof is similar for  $X \in L^1$ .

Adding independent information does not change the conditional expectation:

Lemma Let  $\mathcal{A}_1, \mathcal{A}_2 \subset \mathcal{F}$  be  $\sigma$ -fields and  $X$  a r.v. which is  $\in [0, \infty]$  or  $\mathbb{R}^2$ . If  $\mathcal{A}_2 \perp \sigma(\sigma(X), \mathcal{A}_1)$ , then  $\mathbb{E}[X | \sigma(\mathcal{A}_1, \mathcal{A}_2)] = \mathbb{E}[X | \mathcal{A}_1]$

Proof: We show that  $\mathbb{E}[\mathbb{1}_C X] = \mathbb{E}[\mathbb{1}_C \mathbb{E}[X | \sigma(\mathcal{A}_1)]]$  for every  $C$  in a generating  $\pi$ -system of  $\sigma(\mathcal{A}_1, \mathcal{A}_2)$ . Indeed, since  $\mathbb{E}[X | \sigma(\mathcal{A}_1)]$  is  $\sigma(\mathcal{A}_1, \mathcal{A}_2)$  measurable, this will imply the result (see Exercise 2 Sheet 8).

We use  $\{A_1 \cap A_2 : A_1 \in \mathcal{A}_1, A_2 \in \mathcal{A}_2\}$ , which is a generating  $\pi$ -system of  $\sigma(\mathcal{A}_1, \mathcal{A}_2)$ .

$$\begin{aligned} \text{For } A_1 \in \mathcal{A}_1 \text{ and } A_2 \in \mathcal{A}_2, \quad \mathbb{E}[\mathbb{1}_{A_1 \cap A_2} X] &= \mathbb{E}[\mathbb{1}_{A_2} \mathbb{1}_{A_1} X] \\ &= \mathbb{E}[\mathbb{1}_{A_2}] \mathbb{E}[\mathbb{1}_{A_1} X] \quad \text{by } \perp \\ &= \mathbb{E}[\mathbb{1}_{A_2}] \mathbb{E}[\mathbb{1}_{A_1} \mathbb{E}[X | \mathcal{A}_1]] \\ &= \mathbb{E}[\mathbb{1}_{A_2} \cdot \mathbb{1}_{A_1} \mathbb{E}[X | \mathcal{A}_1]] \end{aligned}$$

and we get  $\sim$

Corollary If  $X \perp Y$ ,  $\mathbb{E}[X | Y] = \mathbb{E}[X]$

(take  $\mathcal{A}_1 = \{\emptyset, \Omega\}$  and  $\mathcal{A}_2 = \sigma(Y)$ )

## b) Conditional density functions (not covered in class: optional)

Assume that  $X, Y$  take values in  $\mathbb{R}^m, \mathbb{R}^n$  and that  $(X, Y)$  has a density:

$f_{(X,Y)}(x,y) = f_{(X,Y)}(x,y)$ . Let  $f_Y(y) = \int_{\mathbb{R}^m} f_{(X,Y)}(x,y) dx$  be a density of  $Y$ .

Then for  $h: \mathbb{R}^m \rightarrow \mathbb{R}_+$  and  $g: \mathbb{R}^n \rightarrow \mathbb{R}_+$  measurable, we have

$$\mathbb{E}[h(X)g(Y)] = \int_{\mathbb{R}^m \times \mathbb{R}^n} h(x)g(y) f_{(X,Y)}(x,y) dx dy \quad (\text{transfer theorem})$$

$$= \int_{\mathbb{R}^n} g(y) f_Y(y) dy \int_{\mathbb{R}^m} \frac{h(x) f_{(X,Y)}(x,y)}{f_Y(y)} \mathbb{1}_{f_Y(y) > 0} dx$$

$$= \mathbb{E}[h(X)g(Y)]$$

With 
$$\varphi(y) = \begin{cases} \frac{1}{f_Y(y)} \int_{\mathbb{R}^m} h(x) f_{X|Y}(x|y) dx & \text{if } f_Y(y) > 0 \\ 0 & \text{otherwise} \end{cases}$$

Thus  $E[h(X) | Y] = \varphi(Y)$ .

We interpret this result by writing

$$E[h(X) | Y] = \int_{\mathbb{R}^m} h(x) \nu(y, dx)$$

where  $\nu(y, dx) = \frac{1}{f_Y(y)} \mathbb{1}_{f_Y(y) > 0} dx = f_{X|Y}(x|y) dx$ . The measure

$\nu(y, dx)$  is called conditional distribution given  $Y=y$  and  $f_{X|Y}(x|y)$  is the conditional density function of  $X$  given  $Y=y$ . Notice that this function is defined only up to a zero-measure set.