

**PROBABILITY THEORY (D-MATH)
EXERCISE SHEET 1 – SOLUTION**

Exercise 1. This exercise shows that a simple random walk on \mathbb{Z} cannot be confined to a strip forever. Let $(X_n)_{n \geq 1}$ be an iid sequence of random variables defined by

$$P(X_1 = 1) = P(X_1 = -1) = 1/2.$$

For $n \geq 1$, define $S_n = X_1 + \dots + X_n$. Let $k \geq 1$ be a fixed integer. Show that

$$P(\forall n \geq 1 \ 0 \leq S_n \leq k) = 0.$$

Solution. For each $i \geq 1$ define

$$A_i = \{X_{2ik+1} = X_{2nk+2} = \dots = X_{2ik+k+1} = 1\}.$$

We claim that for all $i \geq 1$,

$$A_i \subset \{\forall n \geq 1 \ 0 \leq S_n \leq k\}^c.$$

To check this, let $\omega \in A_i$. We show that then $\omega \in \{\forall n \geq 1 \ 0 \leq S_n \leq k\}^c$. If $S_{2ki}(\omega) < 0$, then we are done, and if $S_{2ki}(\omega) \geq 0$, then

$$S_{2ki+k+1} = S_{2ki} + X_{2ik+1} + \dots + X_{2ik+k+1} \geq k + 1,$$

which proves the claim. So we have $\{\forall n \geq 1 \ 0 \leq S_n \leq k\} \subset A_i^c$, and so

$$\{\forall n \geq 1 \ 0 \leq S_n \leq k\} \subset \bigcap_{i \geq 1} A_i^c.$$

We show that $P(\bigcap_{i \geq 1} A_i) = 0$, which would complete the proof. To see this, first note that by independence of (X_n) we have

$$\forall i \geq 1 \quad P(A_i) = (1/2)^{k+1}.$$

Next, observe that, by construction, $(A_i)_{i \geq 1}$ are independent events, so for all $n \geq 1$ we have

$$P\left(\bigcap_{i=1}^n A_i^c\right) = (1 - (1/2)^{k+1})^n.$$

Therefore,

$$P\left(\bigcap_{i \geq 1} A_i^c\right) = \lim_{n \rightarrow \infty} P\left(\bigcap_{i=1}^n A_i^c\right) = 0,$$

as desired.

Exercise 2 [R]. Let $(X_n)_{n \geq 1}$ be an iid sequence of random variables uniformly distributed in $\{-1, 1, 2, -2\}$. For $n \geq 1$, let $S_n = X_1 + \dots + X_n$. Fix $n \geq 1$.

Solution.

- (i) Compute $E(S_n)$ and $\text{Var}(S_n)$.
- (ii) Prove that

$$P(|S_n| \geq 2\sqrt{n}) \leq \frac{3}{4}.$$

- (iii) Prove that

$$\forall k \in \mathbb{Z} \quad P(S_n = k) = P(S_n = -k).$$

- (iv) Prove that

$$\forall k \in \mathbb{Z} \quad P(X_1 + \dots + X_n = k) = P(X_{n+1} + \dots + X_{2n} = k).$$

- (v) Deduce that

$$\forall k \in \mathbb{Z} \quad P(S_{2n} = k) = \sum_{i \in \mathbb{Z}} P(S_n = i) \cdot P(S_n = k - i).$$

- (vi) Apply the Cauchy-Schwarz inequality to show that

$$\forall k \in \mathbb{Z} \quad P(S_{2n} = k) \leq P(S_{2n} = 0).$$

- (vii) Deduce that

$$P(S_{2n} = 0) \geq \frac{1}{50\sqrt{n}}.$$

Solution.

- (i) Since X_1 is symmetric, $E(X_1) = 0$. So, by linearity, we have

$$E(S_n) = E(X_1) + \dots + E(X_n) = 0.$$

We compute $\text{Var}(X_1) = E(X_1^2) - E(X_1)^2 = 5/2$. Since (X_n) are independent, we have

$$\text{Var}(S_n) = \text{Var}(X_1) + \dots + \text{Var}(X_n) = 5n/2.$$

- (ii) By Chebyshev's inequality, we have

$$P(|S_n| \geq 2\sqrt{n}) \leq \frac{(5n/2)}{4n} \leq 3/4.$$

- (iii) Define a sequence of random variables $(Y_i)_{i \geq 1}$ by $Y_i = -X_i$. Observe that $(Y_i)_{i \geq 1}$ has the same distribution as $(X_i)_{i \geq 1}$. So

$$\forall k \in \mathbb{Z} \quad P(X_1 + \dots + X_n = k) = P(Y_1 + \dots + Y_n = k) = P(X_1 + \dots + X_n = -k)$$

- (iv) This time define a sequence of random variables $(Y_i)_{i \geq 1}$ by $Y_i = X_{n+i}$. Again, $(Y_i)_{i \geq 1}$ has the same distribution as $(X_i)_{i \geq 1}$, so

$$\forall k \in \mathbb{Z} \quad P(X_1 + \dots + X_n = k) = P(Y_1 + \dots + Y_n = k) = P(X_{n+1} + \dots + X_{2n} = k).$$

(v) Fix $k \in \mathbb{Z}$. We have

$$\begin{aligned}
P(S_{2n} = k) &= P\left(\bigcup_{i \in \mathbb{Z}} \{S_n = i, S_{2n} - S_n = k - i\}\right) \\
(\text{the union is disjoint}) &= \sum_{i \in \mathbb{Z}} P(S_n = i, S_{2n} - S_n = k - i) \\
(S_n \text{ and } S_{2n} - S_n \text{ are ind.}) &= \sum_{i \in \mathbb{Z}} P(S_n = i) \cdot P(X_{n+1} + \cdots + X_{2n} = k - i) \\
(\text{part iv}) &= \sum_{i \in \mathbb{Z}} P(S_n = i) \cdot P(S_n = k - i).
\end{aligned}$$

(vi) We have

$$\begin{aligned}
P(S_{2n} = k) &= \sum_{i \in \mathbb{Z}} P(S_n = i) \cdot P(S_n = k - i) \quad (\text{part v}) \\
(\text{Cauchy-Schwarz}) &\leq \left(\sum_{i \in \mathbb{Z}} P(S_n = i)^2\right)^{1/2} \left(\sum_{i \in \mathbb{Z}} P(S_n = k - i)^2\right)^{1/2} \\
(\mathbb{Z} = k - \mathbb{Z}) &= \left(\sum_{i \in \mathbb{Z}} P(S_n = i)^2\right)^{1/2} \left(\sum_{i \in \mathbb{Z}} P(S_n = i)^2\right)^{1/2} \\
&= \sum_{i \in \mathbb{Z}} P(S_n = i)^2 \\
(\text{part iii}) &= \sum_{i \in \mathbb{Z}} P(S_n = i)P(S_n = -i) \\
(\text{part v}) &= P(S_{2n} = 0).
\end{aligned}$$

(vii) Using part (ii) we have $P(|S_{2n}| \in [-2\sqrt{2n}, 2\sqrt{2n}]) = 1 - P(|S_n| > 2\sqrt{n}) \geq 1/4$. Let $I = [-2\sqrt{2n}, 2\sqrt{2n}] \cap \mathbb{Z}$.

Then we have

$$\begin{aligned}
1/4 &\leq P(S_n \in I) \\
&= P\left(\bigcup_{i \in I} S_n = i\right) \\
(\text{the union is disjoint}) &= \sum_{i \in I} P(S_n = i) \\
(\text{part vi}) &\leq |I| \cdot P(S_{2n} = 0) \\
(|I| \leq 10\sqrt{n}) &\leq 10\sqrt{n} \cdot P(S_{2n} = 0).
\end{aligned}$$

So

$$P(S_{2n} = 0) \geq \frac{1}{40\sqrt{n}} > \frac{1}{50\sqrt{n}},$$

which completes the proof.

Exercise 3. Let $(X_n)_{n \geq 1}$ be iid $\text{Exp}(1)$ random variables. Show that

$$\limsup_{n \rightarrow \infty} \frac{X_n}{\log n} = 1 \quad a.s.$$

Solution. Let $\epsilon > 0$ be arbitrary. For $n \geq 1$, define the events

$$A_n = \{X_n \geq (1 + \epsilon) \log n\} \quad \text{and} \quad B_n = \{X_n \geq (1 - \epsilon) \log n\}.$$

First, note that

$$\sum_{n \geq 1} P(A_n) = \sum_{n \geq 1} 1/n^{1+\epsilon} < \infty,$$

so by the first Borel-Cantelli lemma, only finitely many of the A_n 's occur almost surely. Second, we have that

$$\sum_{n \geq 1} P(B_n) = \sum_{n \geq 1} 1/n^{1-\epsilon} = \infty.$$

Since the B_n 's are independent, the second Borel-Cantelli lemma implies that infinitely many of the B_n 's occur. So, we get

$$\limsup_{n \rightarrow \infty} \frac{X_n}{\log n} \in [1 - \epsilon, 1 + \epsilon] \quad a.s.$$

Now, we take a countable sequence $\epsilon_k \rightarrow 0$ to complete the proof. More precisely, we get for all $k \geq 1$,

$$P\left(\limsup_{n \rightarrow \infty} \frac{X_n}{\log n} \in [1 - 1/k, 1 + 1/k]\right) = 1,$$

which implies

$$P\left(\bigcap_{k \geq 1} \limsup_{n \rightarrow \infty} \frac{X_n}{\log n} \in [1 - 1/k, 1 + 1/k]\right) = 1.$$

This completes the proof because event in the display above is the same as

$$\{\limsup_{n \rightarrow \infty} X_n / \log n = 1\}.$$

Exercise 4. Let $(A_n)_{n \geq 1}$ be a sequence of events such that

$$\lim_{n \rightarrow \infty} P(A_n) = 0 \text{ and } \sum_{n \geq 1} P(A_n \setminus A_{n+1}) < \infty.$$

Prove that $P(\text{infinitely many } A_n \text{ occur}) = 0$.

Solution.

We want to show that

$$P\left(\bigcap_{n \geq 1} \bigcup_{m \geq n} A_m\right) = \lim_{n \rightarrow \infty} P\left(\bigcup_{m \geq n} A_m\right) = 0.$$

We claim that

$$\bigcup_{m \geq n} A_m \subset \left(\bigcup_{m \geq n} A_m \setminus A_{m+1}\right) \cup \left(\bigcup_{m \geq n} \bigcap_{k \geq m} A_k\right). \quad (1)$$

To show this, let $\omega \in \bigcup_{m \geq n} A_m$ but $\omega \notin \bigcap_{m \geq n} A_m \setminus A_{m+1}$. Let $m' \geq n$ be such that $\omega \in A_{m'}$. Since $\omega \notin A_{m'} \setminus A_{m'+1}$, we must have $\omega \in A_{m'+1}$. Continuing like this, we get inductively that $\omega \in A_k$ for all $k \geq m'$. This shows $\omega \in \bigcap_{k \geq m'} A_k$, which proves the claim. Next, observe that for each $m \geq n$

$$P\left(\bigcap_{k \geq m} A_k\right) = 0.$$

since $\bigcap_{k \geq n} A_k$ is contained in each of the $(A_k)_{k \geq m}$ and $P(A_k) \rightarrow 0$. So

$$P\left(\bigcup_{m \geq n} \bigcap_{k \geq m} A_k\right) = 0,$$

as well. Using this and union bound in (1) we obtain

$$P\left(\bigcup_{m \geq n} A_m\right) \leq \sum_{k \geq n} P(A_k \setminus A_{k+1}),$$

which converges to 0 as $n \rightarrow \infty$ since $\sum_{n \geq 1} P(A_n \setminus A_{n+1}) < \infty$ by assumption. This completes the proof.