

PROBABILITY THEORY (D-MATH)
EXERCISE SHEET 10 – SOLUTION

Exercise 1. [R] Let A be an compact set in \mathbb{R}^2 and let $(X, Y) \sim \text{Unif}(A)$. Compute

$$\mathbb{E}(X^2|Y)$$

in the following cases:

- (1) $A = [-1, 1]^2$,
 (2) $A = \{(x, y) : |x| + |y| \leq 1\}$.

Solution.

(1) In this case, X and Y are independent $\text{Unif}[-1, 1]$ random variables. So

$$\mathbb{E}(X^2|Y) = \mathbb{E}(X^2) = \int_{-1}^1 x^2/2 \, dx = 1/3.$$

(2) In this case, (X, Y) has density given by

$$f_{(X,Y)}(x, y) = \frac{1_{|x|+|y|\leq 1}}{2}.$$

Therefore, the marginal density of y is given by

$$f_Y(y) = 1_{|y|\leq 1} \int_{|y|-1}^{1-|y|} 1/2 \, dy = (1 - |y|)1_{|y|\leq 1}.$$

We compute the conditional expectation as in section 5 of chapter 10. Let $y \in [-1, 1]$. We get

$$\begin{aligned} \phi(y) &= \int_{-1}^1 \frac{x^2 1_{|x|+|y|\leq 1}}{2(1 - |y|)} \, dx \\ &= \frac{1}{2(1 - |y|)} \int_{|y|-1}^{1-|y|} x^2 \, dx \\ &= \frac{(1 - |y|)^2}{3}. \end{aligned}$$

So

$$\mathbb{E}(X^2|Y) = \frac{(1 - |Y|)^2}{3} \quad a.s.$$

Exercise 2. Let X, Y be independent random variables and let $\psi : \mathbb{R}^2 \rightarrow \mathbb{R}_{\geq 0}$ be a measurable function such that

$$\mathbb{E}(|\psi(X, Y)|) < \infty.$$

Define $\phi : \mathbb{R} \rightarrow [0, \infty]$ by

$$\phi(y) = \mathbb{E}(\psi(X, y)).$$

Show that

$$\mathbb{E}(\psi(X, Y)|Y) = \phi(Y) \text{ a.s.}$$

Solution. First, since $\phi(y) = \int_{\mathbb{R}} \psi(x, y) \mu_X(dx)$, ϕ is a measurable function (this is part of Fubini's theorem). Hence, $\phi(Y)$ is $\sigma(Y)$ -measurable. Therefore, it suffices to check that for every non-negative random variable Z which is $\sigma(Y)$ -measurable, we have

$$\mathbb{E}(\psi(X, Y)Z) = \mathbb{E}(\phi(Y)Z).$$

To this end, using Fubini's theorem we get

$$\begin{aligned} \mathbb{E}(\psi(X, Y)Z) &= \int_{\mathbb{R} \times \mathbb{R} \times \mathbb{R}_{\geq 0}} \psi(x, y) z \mu_{(X, Y, Z)}(dx dy dz) \\ \text{(because } X \text{ and } (Y, Z) \text{ are independent)} &= \int_{\mathbb{R} \times \mathbb{R} \times \mathbb{R}_{\geq 0}} \psi(x, y) z \mu_X(dx) \mu_{(Y, Z)}(dy dz) \\ &= \int_{\mathbb{R} \times \mathbb{R}_{\geq 0}} z \left(\int_{\mathbb{R}} \psi(x, y) \mu_X(dx) \right) \mu_{(Y, Z)}(dy dz) \\ &= \int_{\mathbb{R} \times \mathbb{R}_{\geq 0}} z \phi(y) \mu_{(Y, Z)}(dy dz) \\ &= \mathbb{E}(\phi(Y)Z). \end{aligned}$$

Exercise 3. [R] Let $(Y_n)_{n \geq 1}$ be iid random variables which are uniform in $\{-1, +1\}$ and let X be a random variable in L^2 . Let $[n]$ denote $\{1, \dots, n\}$ and for a subset $S \subset [n]$, define

$$Y_S = \prod_{i \in S} Y_i,$$

where Y_\emptyset defined to be 1.

- (1) Show that $E(X|Y_1) = E(X) + E(XY_1)Y_1$.
(2) More generally, for all $n \geq 1$ show that

$$E(X|Y_1, \dots, Y_n) = \sum_{S \subset [n]} E(XY_S)Y_S.$$

Solution. We use the notation $Y = (Y_1, \dots, Y_n)$. We prove the general formula using the L^2 projection interpretation of conditional expectation, which is applicable because we assume that $X \in L^2$. Let

$$\mathcal{F} = \sigma(Y_1, \dots, Y_n)$$

and let $\mathcal{H}_{\mathcal{F}}$ be the vector space of \mathcal{F} -measurable random variables in L^2 . We first establish that $\mathcal{H}_{\mathcal{F}}$ has dimension at most 2^n . Consider the random variables

$$\{y \in \{-1, 1\}^n : 1_{Y=y}\}.$$

Being a functions of Y and bounded, these random variables are in $H_{\mathcal{G}}$. We show that they span $\mathcal{H}_{\mathcal{F}}$. Let Z be a \mathcal{F} -measurable random variable in L^2 . Then we know that there exists a measurable function $f : \mathbb{R}^n \rightarrow \mathbb{R}$ such that

$$Z = f(Y_1, \dots, Y_n).$$

But now

$$f(Y_1, \dots, Y_n) = \sum_{y \in \{-1, 1\}^n} f(y)1_{Y=y}.$$

This shows that the dimension of $H_{\mathcal{F}}$ is at most 2^n . Next, we claim that

$$\{S \subset [n] : Y_S\}$$

is an orthonormal basis of $H_{\mathcal{F}}$. First, since $Y_S \in \{-1, 1\}$ we have $E(Y_S^2) = 1$. Next let R, S be distinct subsets of $[n]$. Then

$$E(Y_R Y_S) = E(Y_{R \Delta S}) = E\left(\prod_{k \in R \Delta S} Y_k\right) = \prod_{k \in R \Delta S} E(Y_k) = 0,$$

where we used independence in the second last inequality, and the facts that $R \Delta S \neq \emptyset$ and $E(Y_k) = 0$ for all k in the last. This shows orthonormality. Since we have 2^n such functions, they form a basis. So finally, the formula we wanted to show is just the projection formula with respect to this orthonormal basis.

Exercise 4. [R] Let X be a real-valued random variable defined on $(\Omega, \mathcal{F}, \mathbb{P})$ that takes values in $[0, \infty]$ a.s. Let $\mathcal{G} \subset \mathcal{F}$ be a sigma-algebra. Define $\mathbb{E}(X|\mathcal{G})$ and show that it is unique (up to almost sure equivalence).

Solution. Refer to section 7 of chapter 10.