

CHAPTER 16.
BACKWARD MARTINGALES

Setup: (Ω, \mathcal{F}, P) proba. space.

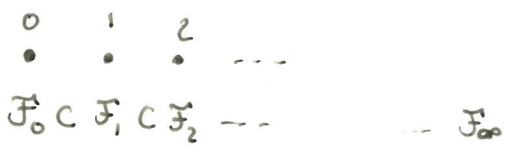
BACKWARD FILTRATION.

Def. a backward filtration is a sequence $(\mathcal{F}_n)_{n \in \mathbb{Z}_-}$ of σ -algebras $(\mathcal{F}_n \subset \mathcal{F})$ s.t.
$$\forall n \in \mathbb{Z}_- \quad \mathcal{F}_{n-1} \subset \mathcal{F}_n$$

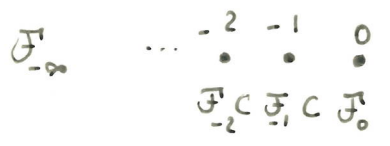
$(\mathbb{Z}_- = \{0, -1, -2, \dots\})$

“gain of information when time increases”
→

Filtration:



Backward filtration



←
“loss of information when we go back in time”

Notation:

$$\mathcal{F}_{-\infty} = \bigcap_{n \in \mathbb{Z}_-} \mathcal{F}_n$$

Example. $(X_n)_{n \in \mathbb{Z}_-}$ n.v. $\mathcal{F}_n = \sigma(X_n, X_{n-1}, X_{n-2}, \dots)$ back. filtr..

In the rest of the chapter, we fix a backward filtration $(\mathcal{F}_n)_{n \leq 0}$

2 BACKWARD MARTINGALES

Def. A sequence of r.v.s $(X_n)_{n \leq 0}$ is a backward martingale if

- $\forall n \leq 0$ X_n is \mathcal{F}_n -measurable and $X_n \in L^1$
- $\forall n \leq 0$ $E(X_n | \mathcal{F}_{n-1}) = X_{n-1}$ a.s.

Ex 1 Let $X \in L^1$. $X_n = E(X | \mathcal{F}_n)$ defines a backward martingale.

$$\begin{aligned}
 (\forall n \leq 0 \quad E(X_n | \mathcal{F}_{n-1}) &= E(E(X | \mathcal{F}_n) | \mathcal{F}_{n-1})) \\
 &= E(X | \mathcal{F}_{n-1})
 \end{aligned}$$

$\mathcal{F}_{n-1} \subset \mathcal{F}_n$

Ex 2 Let z_{-1}, z_{-2}, \dots iid $\text{Ber}(\frac{1}{2})$.

$$\mathcal{F}_n = \sigma(z_{n-1}, z_{n-2}, \dots)$$

$$X_0 := \sum_{n \leq -1} z_n 2^n = z_{-1} \cdot \frac{1}{2} + z_{-2} \cdot \frac{1}{4} + z_{-3} \cdot \frac{1}{8} + \dots$$

$$X_{-1} = \frac{1}{4} + \sum_{k \leq -2} z_k 2^k = \frac{1}{4} + z_{-2} \cdot \frac{1}{4} + z_{-3} \cdot \frac{1}{8} + \dots$$

replace by average

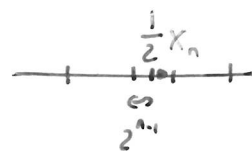
$$X_{-2} = \frac{1}{4} + \frac{1}{8} + \sum_{k \leq -3} z_k 2^k = \frac{1}{4} + \frac{1}{8} + z_{-3} \cdot \frac{1}{8} + \dots$$

$$X_n = \sum_{n \leq k \leq -1} \frac{1}{2} \cdot 2^k + \sum_{k \leq n} z_k 2^k$$

Rk: $X_0 \sim U([0, 1])$

$X_1 \sim U([\frac{1}{4}, \frac{3}{4}])$

$X_n \sim U([\frac{1}{2} - 2^{n-1}, \frac{1}{2} + 2^{n-1}])$



Exercise: check that $(X_n)_{n \leq 0}$ is a backward martingale.

Rk: Let Z_1, Z_2, \dots iid $Ber(\frac{1}{2})$ $\mathcal{F}_n = \sigma(Z_1, \dots, Z_n)$ filtration

$X_n = Z_1 * \frac{1}{2} + \dots + Z_n * \frac{1}{2^n}$ (forward) - martingale.

Prop. Let (X_n) be a backward martingale. Then for every $m, n \leq 0$ s.t. $m \leq n$, we have

$E(X_n | \mathcal{F}_m) = X_m$

Proof. Fix $m \leq 0$. let $k \in \{1, \dots, m\}$

$E(X_{m+k} | \mathcal{F}_m) = E(E(X_{m+k} | \mathcal{F}_{m+k-1}) | \mathcal{F}_m)$
 $\mathcal{F}_m \subset \mathcal{F}_{m+k-1}$

$= E(X_{m+k-1} | \mathcal{F}_m)$ a.s.

And the result follows by finite induction.

Rk: In particular $\forall n \leq 0$

$X_n = E(X_0 | \mathcal{F}_n)$

and $E(X_n) = E(X_0)$

3 CONVERGENCE THM

Thm. Let $(X_n)_{n \leq 0}$ be a backward martingale.

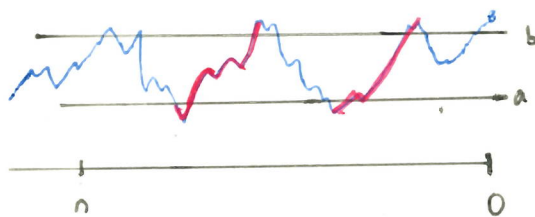
$$X_n \xrightarrow{n \rightarrow -\infty} E(X_0 | \mathcal{F}_\infty) \text{ a.s. and in } L^1.$$

Rk: no hypothesis because a back. martingale is automatically "closed", in particular, since $X_n = E(X_0 | \mathcal{F}_n)$, the sequence is automatically U.I.

Pp. Sketch: a.s. cv.

For $a < b$, $a, b \in \mathbb{Q}$, define

$$N_n(a, b) = \left| \left\{ \text{upcrossings of } (X_n, \dots, X_0) \right\} \right|$$



Doob upcrossing inequality:

$$\forall a < b, a, b \in \mathbb{R} \quad \underbrace{E(N_n(a, b))}_{\downarrow \text{ monotone cv}} \leq \frac{1}{b-a} E((X_0 - a)^-)$$

$$E(N_\infty(a, b))$$

"total number of upcrossings"

Hence a.s. we have

$$\forall a < b \quad N_{-\infty}(a, b) < \infty, \\ a, b \in \mathbb{R}$$

which deterministically implies the existence of $X_{-\infty}$ s.t.

$$X_n \xrightarrow[n \rightarrow -\infty]{} X_{-\infty} \quad \text{a.s.}$$

L^1 cv and identification of the limit.

Since $X_n = E(X_0 | \mathcal{F}_n)$, the sequence (X_n) is UI. Since it cv a.s., it also cv in L^1 .

We have $X_n \xrightarrow[n \rightarrow -\infty]{L^1} X_{-\infty}$. (NB: $X_{-\infty}$ is $\mathcal{F}_{-\infty}$ -meas.)

Let $A \in \mathcal{F}_{-\infty}$. For every $n \leq 0$, $A \in \mathcal{F}_n$, hence

$$\begin{aligned} E(X_0 1_A) &= E(E(X_0 | \mathcal{F}_n) 1_A) \\ &= E(X_n 1_A) \\ &\xrightarrow[n \rightarrow -\infty]{} E(X_{-\infty} 1_A) \quad (\text{because } X_n \rightarrow X_{-\infty}) \end{aligned}$$

Therefore $X_{-\infty} = E(X_0 | \mathcal{F}_{-\infty})$ a.s. □

Application: Let $Z \in L^1$

$$E(Z | \mathcal{F}_n) \xrightarrow{L^1, \text{a.s.}} E(Z | \mathcal{F}_{-\infty})$$

Appt. 2 : $X_n = \frac{1}{2} - 2^{-n-1} + \sum_{k \leq n-1} z_k 2^k$

$$X_n \xrightarrow{n \rightarrow -\infty} \frac{1}{2} \quad \text{a.s. and in } L^1$$

Pf. 1 : $|\sum_{k \leq n-1} z_k 2^k| \leq \sum_{k \leq n-1} 2^k = 2^n \rightarrow 0$ as $n \rightarrow -\infty$.

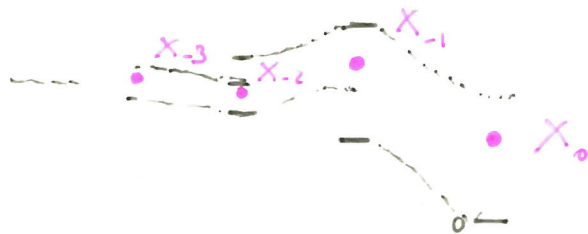
Pf 2 : $X_{-\infty}$ is $\mathcal{F}_{-\infty}$ -meas. Since $\mathcal{F}_{-\infty}$ is P -trivial

($\forall A \in \mathcal{F}_{-\infty} \quad P(A \in \{0, 1\})$) $X_{-\infty} = C$ a.s.

where $C = \text{cte.}$ and

$$C = E(X_{-\infty}) = E(X_0) = \frac{1}{2}.$$

$\frac{1}{2}$ ←



4 LAW OF LARGE NUMBERS REVISITED.

Let z_1, z_2, \dots iid $z_i \in L'$. For $k \geq 0$, $S_k = z_1 + \dots + z_k$

We give a new proof of $\frac{S_k}{k} \xrightarrow[k \rightarrow \infty]{} E(z_1)$ a.s. and in L' .

Define for $k \geq 0$

$$\mathcal{F}_{-k} = \sigma(S_{k+1}, z_{k+2}, z_{k+3}, \dots)$$

$$X_{-k} = \frac{S_{k+1}}{k+1}$$

$$\mathcal{F}_0 = \sigma(z_1, z_2, \dots) \quad X_0 = z_1$$

$$\mathcal{F}_{-1} = \sigma(S_2, z_2, \dots) \quad X_{-1} = \frac{z_1 + z_2}{2}$$

$$\mathcal{F}_{-2} = \sigma(S_3, z_3, \dots) \quad X_{-2} = \frac{z_1 + z_2 + z_3}{3}$$

⋮

Rk: $(\mathcal{F}_n)_{n \leq 0}$ is a backward filtration; Indeed, for all $k \geq 0$

$$\mathcal{F}_{-k-1} = \sigma(S_{k+2}, z_{k+3}, \dots)$$

$$\subset \sigma(S_{k+1}, z_{k+2}, z_{k+3}, \dots) = \mathcal{F}_{-k}$$

because $S_{k+2} = S_{k+1} + z_{k+2}$

Prop. $(X_n)_{n \leq 0}$ is a $(\mathcal{F}_n)_{n \leq 0}$ -backward martingale.

Proof We show that

$$\forall n \leq 0 \quad X_n = E(z_1 | \mathcal{F}_n).$$

Fix $k \geq 1$. We prove

$$\frac{S_k}{k} = E(z_1 | S_k, z_{k+1}, \dots)$$

Since z_{k+1}, z_{k+2}, \dots are indep of (z_1, S_k)
it suffices to show.

$$\frac{S_k}{k} = E(z_1 | S_k).$$

By symmetry, we have

$$\forall i \leq k \quad E(z_1 | S_k) = E(z_i | S_k)$$

(To prove this statement, consider $A \in \sigma(S_k)$

ie $1_A = \varphi(S_k) = \varphi(z_1 + \dots + z_k)$, for some φ meas. bounded.

Since $(z_1, S_k) \stackrel{(d)}{=} (z_i, S_k)$, we have

$$E(z_1 \varphi(z_1 + \dots + z_k)) = E(z_i \varphi(z_1 + \dots + z_k))$$

$$\text{ie } E(z_1 1_A) = E(z_i 1_A)$$

$$\begin{aligned}
 \text{Hence } E(z_i | S_k) &= E(z_1 | S_k) + \dots + E(z_k | S_k) \\
 &= E(S_k | S_k) \\
 &= S_k
 \end{aligned}$$

Hence $\exists X_{-\infty} \in L^1$ s.t. $X_n \xrightarrow[n \rightarrow \infty]{} X_{-\infty}$ a.s. and in L^1
 By Kolmogorov 0-1 Law,

$$X_{-\infty} = \lim_{k \rightarrow \infty} \frac{S_k}{k} \text{ is constant a.s.}$$

(because $\forall a \{ \lim_{k \rightarrow \infty} \frac{S_k}{k} \leq a \} \in \mathcal{G} = \bigcap_{k \geq 0} \sigma(z_k, z_{k+1}, \dots)$.)

$\exists c$ s.t. $X_{-\infty} = c$ a.s.

But $X_{-\infty} = E(z_1 | \mathcal{G}_{-\infty})$ so we must have

$$\boxed{E(z_1) = c}$$

This concludes

$$\frac{z_1 + \dots + z_k}{k} \xrightarrow[k \rightarrow \infty]{} E(z_1) \text{ a.s. and in } L^1$$