

CHAPTER 2  
LAW OF LARGE NUMBERS.

- Goals:
- precise understanding of the statement
  - complete proof of  $L^1$  LLN (Law of Large numbers)
  - strategy vs tactics/tricks in a mathematical proof.
  - methods: truncation / use of subsequence.

- Setup:
- $(\Omega, \mathcal{F}, P)$  probability space (fixed)
  - $L^p = \{X: \Omega \rightarrow \mathbb{R} \text{ measurable } E(|X|^p) < \infty\}$ ,  $p \geq 1$ .  
 $\hookrightarrow$  vector space ( $X, Y \in L^p \Rightarrow X + Y \in L^p$ )

## 1 STATEMENT

Thm (LLN)

Let  $X_1, X_2, \dots \in L^1$ , pairwise independent, identically distributed. Write  $m = E(X_1)$ .

$$\lim_{n \rightarrow \infty} \frac{X_1 + \dots + X_n}{n} = m \quad \text{a.s.}$$

Is it intuitive..?

Consider  $X_n$  iid  $X_n \sim \text{Ber}(\frac{1}{2})$  "coin flip"

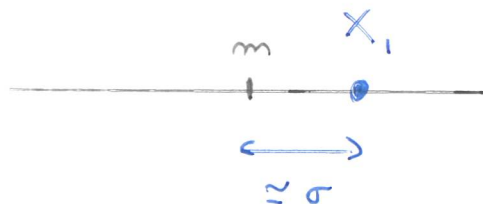
$$\frac{X_1 + \dots + X_n}{n} = \frac{\text{"number of heads in } n \text{ throws"}}{n}$$

$$\rightarrow \frac{1}{2} = E(X_i)$$

A theoretical explanation: "averaging reduces the Variance".

Assume  $X_i \in L^2$ .

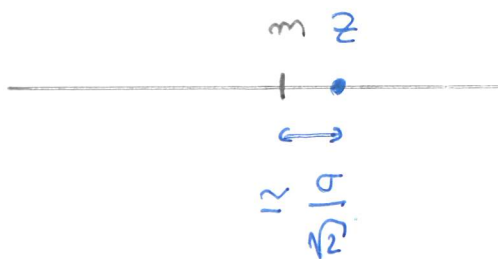
$$\begin{cases} m = E(X_i) & \text{"typical value"} \\ \sigma = \sqrt{\text{Var}(X_i)} & \text{"typical distance from } m \end{cases}$$



Consider  $Z = \frac{X_1 + X_2}{2}$  where  $X_2$  is an independent copy of  $X_1$

$$E(Z) = \frac{1}{2}(E(X_1) + E(X_2)) = m \quad \text{"same typical value as } X_i \text{"}$$

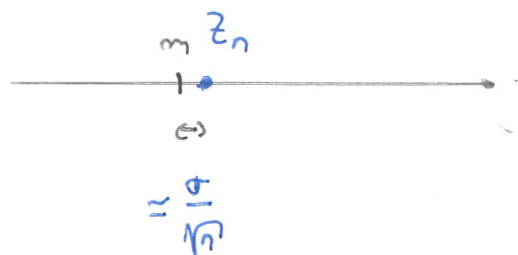
$$\text{Var}(Z) = \frac{1}{4}(\text{Var}(X_1) + \text{Var}(X_2)) = \frac{1}{2}\text{Var}(X_i) \quad \text{"typical distance from } m \text{ is smaller"}$$



$$\text{for } n \geq 1, \quad z_n = \frac{X_1 + \dots + X_n}{n}$$

$$\text{we have } E(z_n) = m$$

$$\sigma_n := \sqrt{\text{var}(z_n)} = \frac{1}{\sqrt{n}}$$



If  $X_i \in L^1 \setminus L^2$  the "typical distance" between  $m$  and  $z_n$  may be more than  $\frac{1}{\sqrt{n}}$  but stays  $o(1)$ .

### Applications:

see WUS Math (Spring 2024):

Monte-Carlo methods to approximate integrals, proof of Stone Weierstrass density theorem, ...

Strategy of the proof in this chapter:  $S_n = X_1 + \dots + X_n$

Idea 1: use the criterion  $\sum P(|\frac{S_n}{n} - m| > \varepsilon) < \infty$ .

Idea 2: use Markov / Chebitchev to estimate

$$P(|\frac{S_n}{n} - m| > \varepsilon).$$

2  $L^4$  - PROOF

Additional assumptions:  $X_1 \in L^4$ , and  $X_1, X_2, \dots$  are iid.

Define: for every  $n$   $Y_n := X_n - m$  ( $E(Y_n) = 0$ )

$$S_n = X_1 + \dots + X_n$$

Let  $\varepsilon > 0$ . For every  $n \geq 1$  we have

$$P\left(\left|\frac{S_n}{n} - m\right| \geq \varepsilon\right) = P\left(\left|\frac{Y_1 + \dots + Y_n}{n}\right| \geq \varepsilon\right)$$

$$\stackrel{\text{Markov}}{\leq} \frac{1}{\varepsilon^4 n^4} E\left((Y_1 + \dots + Y_n)^4\right)$$

$$= \frac{1}{\varepsilon^4 n^4} \sum_{1 \leq i, j, k, \ell \leq n} E(Y_i Y_j Y_k Y_\ell)$$

As soon as a factor  $X_\alpha$  appears exactly once in the product  $X_i X_j X_k X_\ell$ , then independence implies that the expectation of this term vanishes.

The only non-vanishing terms are of the form

$$E(X_i^2 X_j^2) \quad 1 \leq i, j \leq n.$$

Therefore, for every  $\varepsilon > 0$

$$\begin{aligned}
 P\left(\left|\frac{S_n}{n} - m\right| \geq \varepsilon\right) &\leq \frac{1}{\varepsilon^4 n^4} \sum_{1 \leq i, j \leq n} \underbrace{E(Y_i^2 Y_j^2)} \\
 &\stackrel{CS}{\leq} \sqrt{E(Y_i^4)} \sqrt{E(Y_j^4)} \\
 &= E(Y_1^4) \\
 &\leq \frac{3 E(Y_1^4)}{\varepsilon^4} \times \frac{1}{n^2}
 \end{aligned}$$

Since  $\sum_{n \geq 1} P\left(\left|\frac{S_n}{n} - m\right| \geq \varepsilon\right) < \infty$ , we have

$$\lim_{n \rightarrow \infty} \frac{S_n}{n} = m \quad \text{a.s.}$$

### 3 $L^2$ -PROOF

Assumption:  $X_1, X_2, \dots \in L^2$ , ident. distributed, pairwise independent.

For every  $\varepsilon > 0$ ,  $n \geq 1$ , we have

$$P\left(\left|\frac{S_n}{n} - m\right| \geq \varepsilon\right) \stackrel{\substack{\uparrow \\ \text{Chebyshev} \\ \text{inequality.}}}{\leq} \frac{1}{\varepsilon^2} \text{Var}\left(\frac{S_n}{n}\right) = \frac{1}{\varepsilon^2 n} \text{Var}(X_1) \stackrel{\substack{\uparrow \\ \text{pairwise} \\ \text{indep.}}}{=}$$

Pb:  $\sum_{n \geq 1} \frac{1}{n} = +\infty$ .

 Consider subsequences.

Lemma.

Let  $(u_n)_{n \geq 0}$  be a sequence of real numbers satisfying:

①  $\forall 1 \leq m \leq n \quad 0 \leq m u_m \leq n u_n$ ,

②  $\exists p \geq 0$  s.t.

$$\forall \alpha \in (1, \infty) \cap \mathbb{Q} \quad \lim_{i \rightarrow \infty} u_{\lfloor \alpha^i \rfloor} = p.$$

Then  $\lim_{n \rightarrow \infty} u_n = p$ .

Proof. Let  $\alpha > 1$ .

To every  $n$ , we associate  $i = i(n)$  satisfying

$$\lfloor \alpha^i \rfloor \leq n \leq \lfloor \alpha^{i+1} \rfloor$$

By ①, we have

$$\forall n \geq 1 \quad \lfloor \alpha^i \rfloor u_{\lfloor \alpha^i \rfloor} \leq n u_n \leq \lfloor \alpha^{i+1} \rfloor u_{\lfloor \alpha^{i+1} \rfloor}$$

Dividing by  $n$  and using  $\lfloor \alpha^i \rfloor \leq n \leq \lfloor \alpha^{i+1} \rfloor$  we get

$$\forall n \geq 1 \quad \frac{\lfloor \alpha^i \rfloor}{\lfloor \alpha^{i+1} \rfloor} u_{\lfloor \alpha^i \rfloor} \leq u_n \leq \frac{\lfloor \alpha^{i+1} \rfloor}{\lfloor \alpha^i \rfloor} u_{\lfloor \alpha^{i+1} \rfloor}.$$

Therefore, by taking  $n \rightarrow \infty$  and using ②, we get

$$\frac{p}{\alpha} \leq \liminf u_n \leq \limsup u_n \leq \alpha p.$$

Since  $\alpha > 1$  is arbitrary, this concludes

$$\lim_{n \rightarrow \infty} u_n = p \quad \blacksquare$$

We now give a proof of the LLN when  $X_i \in L^2$ ,  $X_i \geq 0$ .

For every  $\alpha > 1$ ,  $\alpha \in \mathbb{Q}$  and  $\varepsilon > 0$ , we have

$$\sum_{i \geq 1} P\left(\left|\frac{S_{L\alpha^i}}{L\alpha^i} - m\right| \geq \varepsilon\right) \leq \frac{\text{Var}(X_1)}{\varepsilon^2} \sum_{i \geq 1} \frac{1}{L\alpha^i} < \infty,$$

Hence, for every  $\alpha \in (1, \infty) \cap \mathbb{Q}$ , we have

$$\lim_{i \rightarrow \infty} \frac{S_{L\alpha^i}}{L\alpha^i} = m \quad \text{a.s.}$$

Since  $S_n \uparrow$ , by the lemma we get

$$\lim_{n \rightarrow \infty} \frac{S_n}{n} = m \quad \text{a.s.}$$

It remains to prove the result without the assumption  $X_i \geq 0$ .

Decompose for every  $n \geq 1$

$$X_n = X_n^+ - X_n^- \quad \text{where } X_n^+, X_n^- \geq 0.$$

The proof follows by applying the LLN to  $(X_n^-)_{n \geq 1}$  and  $(X_n^+)$ , which holds because the sequences are also in  $L^2$ , pairwise independent and identically distributed.

#### 4 $L^1$ - PROOF.

Assumption :  $X_1, X_2, \dots \in L^1$  pairwise independent, id. distributed

Notation :  $X = X_1$

Pb: If  $X \in L^1$   $\text{Var}(X) = E(X^2) - E(X)^2 = +\infty$

We cannot say  $\text{Var}\left(\frac{S_n}{n}\right) \ll 1 \dots$

We will use a truncation argument.

Step 0: WLOG  $X \geq 0$  a.s.

(as in the  $L^2$ -case)



Step 1: truncation.

Define for every  $n \geq 1$

$$Y_n := X_n \mathbb{1}_{X_n \leq n}$$

$$\sum_{n \geq 1} P(X_n \neq Y_n) = \sum_{n \geq 1} P(X_n > n) \\ \leq \int_{n-1}^n P(X \geq x) dx$$

$$\leq \int_0^{\infty} P(X \geq x) dx = E(X) < \infty.$$

By Borel-Cantelli I, almost surely

$$\exists n_0 \text{ s.t. } \forall n \geq n_0 \quad X_n = Y_n \dots$$

It suffices to prove

$$\frac{Y_1 + \dots + Y_n}{n} \xrightarrow[n \rightarrow \infty]{\text{a.s.}} m$$

Step 2: Cesaro

Define for every  $n \geq 1$

$$m_n := E(Y_n) = E(X \mathbb{1}_{X \leq n})$$

By dominated convergence  $m_n \xrightarrow[n \rightarrow \infty]{} m$

By Cesaro Theorem,

$$\frac{m_1 + \dots + m_n}{n} \xrightarrow{n \rightarrow \infty} m.$$

Defining  $\Sigma_n := Y_1 + \dots + Y_n$  for every  $n \geq 1$ ,

this implies

$$\frac{E(\Sigma_n)}{n} \xrightarrow{n \rightarrow \infty} m.$$

Step 3: Chebyshev inequality

Let  $\varepsilon > 0$ . For every  $n \geq 1$ , we have

$$P\left(\left|\frac{\Sigma_n - E(\Sigma_n)}{n}\right| \geq \varepsilon\right) \leq \frac{\text{Var}(Y_1) + \dots + \text{Var}(Y_n)}{n^2 \varepsilon^2}$$

$$\leq \frac{E(Y_1^2) + \dots + E(Y_n^2)}{n^2 \varepsilon^2}$$

$\text{Var}(Y_i) \leq E(Y_i^2) \rightarrow$

$$(\leq n \Rightarrow E(Y_i^2) \leq E(Y_n^2)) \rightarrow \leq \frac{1}{n \varepsilon^2} E(Y_1^2 \mathbb{1}_{Y_1 \leq n})$$

Step 4. using subsequences and Ritz's trick

Let  $\alpha > 1$ . Let  $\epsilon > 0$ .

$$\epsilon^2 \sum_{i \geq 0} P \left( \left| \frac{\sum_{L\alpha^i j} - E(\sum_{L\alpha^i j})}{L\alpha^i j} \right| \geq \epsilon \right)$$

$$\leq \sum_{i \geq 0} \frac{1}{L\alpha^i j} E \left( X^2 1_{X \leq L\alpha^i j} \right)$$

$$\leq \alpha * \sum_{i \geq 0} \frac{1}{\alpha^i} E \left( X^2 1_{X \leq \alpha^i} \right)$$

$$\stackrel{\text{Fubini}}{=} \alpha E \left( X^2 * \underbrace{\sum_{i: \alpha^i \geq X} \frac{1}{\alpha^i}} \right)$$

$$\stackrel{*}{\leq} \frac{\alpha}{\alpha-1} * \frac{1}{X} * 1_{X > 0}$$

$$\leq \frac{\alpha^2}{\alpha-1} E(X) < \infty$$

Hence,

$$\forall \alpha > 1 \quad \frac{\sum_{L\alpha^i j}}{L\alpha^i j} - \frac{E(\sum_{L\alpha^i j})}{L\alpha^i j} \xrightarrow[i \rightarrow \infty]{a.s.} 0$$

\* see Lemma at the end of the section

i.e. by step 2.

$$\forall \alpha > 1 \quad \frac{\sum \alpha^i j}{L \alpha^i} \xrightarrow[i \rightarrow \infty]{a.s.} m.$$

Conclusion.

By the geometric subsequence lemma (applied to  $(\frac{\sum_n}{n})$ ), we conclude

$$\lim_{n \rightarrow \infty} \frac{\sum_n}{n} = m \quad a.s.$$

Lemma (Rittik's trick)

Let  $x > 0, \alpha > 1$ .

$$\sum_{i: \alpha^i \geq x} \frac{1}{\alpha^i} \leq \frac{\alpha}{\alpha-1} \cdot \frac{1}{x}.$$

Proof. Let  $i_0 \in \mathbb{N}$  s.t.  $\alpha^{i_0} \geq x > \alpha^{i_0-1}$ .

$$\sum_{i: \alpha^i \geq x} \frac{1}{\alpha^i} = \sum_{i \geq i_0} \frac{1}{\alpha^i} = \frac{1}{\alpha^{i_0}} \cdot \frac{1}{1 - \frac{1}{\alpha}}$$

$$\leq \frac{1}{x} \cdot \frac{1}{\alpha - 1}$$