## **Mathematical Foundations for Finance Exercise Sheet 3**

**Exercise 3.1** Let  $(\Omega, \mathcal{F}, \mathbb{P})$  be a probability space endowed with the filtration  $\mathbb{F} = (\mathcal{F}_k)_{k=0}^T$  with  $\mathcal{F}_0$  trivial. Let  $X = (X_k)_{k=0}^T$  be a supermartingale. Show that  $X_0 \geq E[X_T]$  always, and that we have  $X_0 = E[X_T]$  if and only if X is a martingale.

**Solution 3.1** The process *X* is a supermartingale, so  $E[X_T | \mathcal{F}_0] \leq X_0$ , and since  $\mathcal{F}_0$  is trivial  $E[X_k] = E[X_k | \mathcal{F}_0] \leq X_0$ . So  $E[X_T] \leq E[X_k] \leq E[X_0]$  and  $E[X_T - X_k \mid F_k] \leq 0$  has expectation  $E[X_T] - E[X_k]$ . If X is a martingale, we have equality everywhere and hence  $E[X_T] = E[X_0]$ . Conversely, if  $E[X_T] = E[X_0]$ , then  $E[X_T] \leq E[X_k] = E[X_0]$  implies  $E[X_T] = E[X_k]$ ; so the nonpositive random variable  $E[X_T - X_k | \mathcal{F}_k]$  has expectation zero and hence must be zero *P*-a.s.. This gives  $E[X_T - X_k | \mathcal{F}_k] = 0$  *P*-a.s. and so *X* is a martingale.

**Exercise 3.2** Consider a filtered probability space  $(\Omega, \mathcal{F}, \mathbb{F}, P)$ , where  $\mathbb{F} = (\mathcal{F}_k)_{k \in \mathbb{N}_0}$ .

(a) Let *X* be a martingale. Show that for any bounded and convex function *f* :  $\mathbb{R} \to \mathbb{R}$ , the process  $f(X) = (f(X_k))_{k \in \mathbb{N}_0}$  is a submartingale.

Could we replace the request of *f* being bounded with a more general condition?

*Hint: You may use that finite-valued convex functions are continuous.*

(b) Let *X* be a submartingale, and let  $\theta = (\theta_k)_{k \in \mathbb{N}_0}$  be a bounded, nonnegative and predictable process. Show that the stochastic integral process  $\vartheta \bullet X$ , defined by

$$
\vartheta \bullet X_k = \sum_{j=1}^k \vartheta_j \Delta X_j = \sum_{j=1}^k \vartheta_j (X_j - X_{j-1}),
$$

is a submartingale.

Conclude that  $E[\vartheta \bullet X_k] \geq 0$  for all  $k \in \mathbb{N}_0$ .

(c) Let X be a submartingale and let  $\tau$  be a stopping time. Show that the stopped process  $X^{\tau} = (X^{\tau}_k)_{k \in \mathbb{N}_0}$  defined by  $X^{\tau}_k = X_{k \wedge \tau}$  is a submartingale.

## **Solution 3.2**

(a) The process  $f(X)$  is integrable because f is bounded. Since X is adapted (because it is a martingale) and *f* is continuous (since it is finite-valued and

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inequality. For  $0 \leq m < n$ , we write

$$
E[f(X_n) | \mathcal{F}_m] \geq f(E[X_n | \mathcal{F}_m]) = f(X_m),
$$

where the first step used the (conditional) Jensen's inequality, and the second step the martingale property. This concludes the proof.

A look at the proof shows that if we replace the condition "*f* is bounded" by  ${}^{\ast}f(X)$  is integrable", the result still holds.

(b) Since  $\vartheta$  is predictable and *X* is adapted, then  $\vartheta_i(X_i - X_{i-1})$  is  $\mathcal{F}_i$ -measurable for all  $j \in \mathbb{N}$ . It follows that  $\vartheta \bullet X_k$  is  $\mathcal{F}_k$ -measurable, so that  $\vartheta \bullet X$  is adapted. Also, since  $\vartheta$  is bounded and X is integrable, we have that  $\vartheta \bullet X$  is integrable. It remains to establish the submartingale inequality. Note that it suffices to show

$$
E[\vartheta \bullet X_{k+1} - \vartheta \bullet X_k \mid \mathcal{F}_k] \geq 0, \ \forall k \in \mathbb{N}_0.
$$

To this end, we write

$$
E[\vartheta \bullet X_{k+1} - \vartheta \bullet X_k \mid \mathcal{F}_k] = E[\vartheta_{k+1}(X_{k+1} - X_k) \mid \mathcal{F}_k]
$$
  
=  $\vartheta_{k+1} E[X_{k+1} - X_k \mid \mathcal{F}_k],$ 

where in the last step we used that  $\vartheta_{k+1}$  is  $\mathcal{F}_k$ -measurable and bounded. Since X is a submartingale, then  $E[X_{k+1} - X_k | \mathcal{F}_k] \geq 0$ . Since also  $\vartheta_{k+1}$  is nonnegative by assumption, we have

$$
E[\vartheta \bullet X_{k+1} - \vartheta \bullet X_k \mid \mathcal{F}_k] \geq 0,
$$

as required.

Since  $\vartheta \bullet X$  is a submartingale null at zero, we have for all  $k \in \mathbb{N}_0$  that

$$
E[\vartheta \bullet X_k] = E\big[E[\vartheta \bullet X_k \mid \mathcal{F}_0]\big] \geqslant E[\vartheta \bullet X_0] = 0.
$$

(c) For  $k \in \mathbb{N}_0$ , we have

$$
X_k^{\tau} = X_{k \wedge \tau} = X_0 + \sum_{j=1}^{k \wedge \tau} (X_j - X_{j-1}) = X_0 + \sum_{j=1}^k 1_{\{\tau \ge j\}} (X_j - X_{j-1}).
$$

So if we set  $\vartheta = (\vartheta_k)_{k \in \mathbb{N}}$  with  $\vartheta_k := 1_{\{\tau \geq k\}}$ , then

$$
X_k^{\tau} = X_0 + \vartheta \bullet X_k, \ \forall k \in \mathbb{N}_0.
$$

Since  $\tau$  is a stopping time, then  $\vartheta$  is a predictable process. Since  $\vartheta$  is also bounded and nonnegative, and *X* is a submartingale, we may apply part (b) to conclude that  $\vartheta \bullet X$  is a submartingale. Also, note that because  $X_0$  is  $\mathcal{F}_0$ -measurable and integrable, then the process  $(X_0)_{k \in \mathbb{N}_0}$  is a submartingale (in fact a martingale). Since the sum of two submartingales is a submartingale, we can conclude that  $X^{\tau}$  is a submartingale, as required.

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**Exercise 3.3** Let  $(\tilde{S}^0, \tilde{S}^1)$  be a *trinomial model*. This is like a binomial model a special case of a *multinomial model*, and the distribution of  $Y_k$  under  $P$  is given by

$$
Y_k = \begin{cases} 1+d & \text{with probability } p_1 \\ 1+m & \text{with probability } p_2 \\ 1+u & \text{with probability } p_3 \end{cases}
$$

where  $p_1, p_2, p_3 > 0, p_1 + p_2 + p_3 = 1$  and  $-1 < d < m < u$ . Here *d*, *m* and *u* are mnemonics for *down*, *middle* and *up*.

(a) Assume that *d* = −0*.*5, *m* = 0, *u* = 0*.*25 and *r* = 0. For *T* = 1, consider an arbitrary self-financing strategy  $\varphi \cong (V_0, \theta)$ . Show that if the total gain  $G_1(\theta)$ at time  $T = 1$  is nonnegative  $P$ -a.s., then

$$
P[G_1(\theta) = 0] = 1.
$$

What does this property imply?

(b) Show that  $S<sup>1</sup>$  is arbitrage-free by constructing an *equivalent martingale measure*  $(EMM)$  for  $S^1$ .

*Hint:* A probability measure Q equivalent to P on  $\mathcal{F}_1$  can be uniquely described by a *probability vector*  $(q_1, q_2, q_3) \in (0, 1)^3$ , where  $q_k = Q[Y_1 = 1 + y_k]$ ,  $k = 1$ , 2, 3, using the notation  $y_1 := d$ ,  $y_2 := m$  and  $y_3 := u$ . (A *probability vector in*  $\mathbb{R}^n$ ,  $n \in \mathbb{N}$  *is a nonnegative vector in*  $\mathbb{R}^n$  *whose coordinates sum up to 1.*)

(c) Assume now that  $d = -0.01$ ,  $m = 0.01$ ,  $u = 0.03$  and  $r = 0.01$ . For  $T = 2$ , give a parametrisation of all *equivalent martingale measures* (EMMs) for *S* 1 .

*Hint:* A probability measure Q equivalent to P on  $\mathcal{F}_2$  can be uniquely described by four *probability vectors*  $(q_1, q_2, q_3), (q_{j,1}, q_{j,2}, q_{j,3}) \in (0, 1)^3, j = 1, 2, 3$ , where  $q_j = Q[Y_1 = 1 + y_j]$  and  $q_{j,k} = Q[Y_2 = 1 + y_k | Y_1 = 1 + y_j], j, k = 1, 2, 3,$ using the notation  $y_1 := d$ ,  $y_2 := m$  and  $y_3 := u$ .

## **Solution 3.3**

(a) Let us compute the total gain  $G_1(\theta)$  at time  $T=1$ :

$$
G_1(\theta) = \theta_1^1 \Delta S_1^1 = \theta_1^1 (S_1^1 - S_0^1) = \theta_1^1 S_0^1 \left( \frac{Y_1}{1+r} - 1 \right) = \theta_1^1 S_0^1 \times \begin{cases} \frac{u-r}{1+r} & \text{with probability } p_3, \\ \frac{m-r}{1+r} & \text{with probability } p_2, \\ \frac{d-r}{1+r} & \text{with probability } p_1. \end{cases}
$$

Recall that  $u - r = 0.25 > 0$  and  $d - r = -0.5 < 0$ . Hence  $P[G_1(\theta) > 0] = 1$  if and only if  $\theta_1^1 S_0^1 = 0$ . As a result, we can conclude that

 $P[G_1(\theta) \ge 0] = 1 \Leftrightarrow \theta_1^1 = 0 \Leftrightarrow P[G_1(\theta) = 0] = 1.$ 

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Assume now that  $V_0 = 0$  and note that in this case  $V_1(\varphi) = G_1(\theta)$ . The above argument proves that if  $V_1(\varphi) \geq 0$  *P*-a.s., then  $V_1(\varphi) = 0$  *P*-a.s., and by Proposition 1.1 3) in the lecture notes, we know that this is equivalent to saying that  $S^1$  is arbitrage-free.

(b) Let  $(q_1, q_2, q_3) \in (0, 1)^3$  be a probability vector and *Q* be defined by

 $Q[Y_1 = 1 + y_k] := q_k, \quad k = 1, 2, 3,$ 

where  $y_1 := d$ ,  $y_2 := m$  and  $y_3 := u$ . Then  $S^1$  is a *Q*-martingale if and only if  $S<sup>1</sup>$  is adapted to the considered filtration (take the filtration generated by  $S<sup>1</sup>$ itself), integrable (the probability space is finite here, so all random variables are integrable), and

$$
E_Q\left[S_1^1\right] = S_0^1 \Leftrightarrow E_Q\left[S_0^1 Y_1/(1+r)\right] = S_0^1 \Leftrightarrow E_Q\left[Y_1\right] = 1+r
$$
  
\n
$$
\Leftrightarrow q_1 \times (1+d) + q_2 \times (1+m) + q_3 \times (1+u) = 1+r
$$
  
\n
$$
\Leftrightarrow q_1 \times d + q_2 \times m + q_3 \times u = r
$$
  
\n
$$
\Leftrightarrow -0.5q_1 + 0q_2 + 0.25q_3 = 0
$$
  
\n
$$
\Leftrightarrow q_3 = 2q_1.
$$

Recall that in order to make *Q* a probability measure, we need to have  $q_1 + q_2 +$  $q_3 = 1$ ; hence choosing  $q_1 = 0.25$ , we obtain that  $q_3 = 0.5$  and  $q_2 = 0.25$ . Noting that  $q_1, q_2, q_3 \in (0, 1)$ , we can also observe that Q is a probability measure equivalent to *P* and thus an EMM for *S* 1 .

(c) Let  $(q_1, q_2, q_3), (q_{j,1}, q_{j,2}, q_{j,3}) \in (0, 1)^3$ ,  $j = 1, 2, 3$ , be probability vectors and  $Q \approx P$  on  $\mathcal{F}_2 = \sigma(Y_1, Y_2)$  be defined by

$$
Q[Y_2 = 1 + y_k, Y_1 = 1 + y_j] := q_j q_{j,k}, \quad j, k = 1, 2, 3,
$$

where  $y_1 := d$ ,  $y_2 := m$  and  $y_3 := u$ . Then  $S^1$  is a *Q*-martingale if and only if it is adapted, integrable and

$$
E_Q\left[S_1^1\right] = S_0^1 \text{ and } E_Q\left[S_2^1 \middle| \mathcal{F}_1\right] = S_1^1 \quad Q\text{-a.s.}
$$
  
\n
$$
\Leftrightarrow E_Q\left[S_0^1 Y_1 / (1+r)\right] = S_0^1 \text{ and } E_Q\left[S_0^1 Y_1 Y_2 / (1+r)^2 \middle| \mathcal{F}_1\right] = S_0^1 Y_1 / (1+r) \quad Q\text{-a.s.}
$$
  
\n
$$
\Leftrightarrow E_Q\left[Y_1\right] = 1+r \text{ and } E_Q\left[Y_2 \middle| \mathcal{F}_1\right] = 1+r \quad Q\text{-a.s.}
$$

Since  $\mathcal{F}_1 = \sigma(Y_1)$  and  $Y_1$  takes three values, the latter is equivalent to

$$
E_Q[Y_1] = 1 + r
$$
 and  $E_Q[Y_2|Y_1 = 1 + y_j] = 1 + r$ ,  $j = 1, 2, 3$ .

For the first equation we can compute

$$
E_Q[Y_1] = 1 + r \Leftrightarrow q_1 \times (1 + d) + q_2 \times (1 + m) + q_3 \times (1 + u) = 1 + r
$$
  
\n
$$
\Leftrightarrow q_1 \times d + q_2 \times m + q_3 \times u = r
$$
  
\n
$$
\Leftrightarrow -0.01q_1 + 0.01q_2 + 0.03q_3 = 0.01
$$
  
\n
$$
\Leftrightarrow -q_1 + q_2 + 3q_3 = 1.
$$

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<span id="page-4-0"></span>Since *Q* is a probability measure equivalent to *P*, the triplet  $(q_1, q_2, q_3)$  has to satisfy

$$
\begin{cases}\n-q_1 + q_2 + 3q_3 = 1 \\
q_1 + q_2 + q_3 = 1 \\
q_1, q_2, q_3 \in (0, 1)\n\end{cases}
$$

Subtracting the second equation from the first yields

$$
2q_3 - 2q_1 = 0 \quad \Leftrightarrow \quad q_1 = q_3 \, .
$$

This in turn implies  $q_2 = 1 - 2q_1$ , and by the positivity constraint  $0 < q_1 < 0.5$ . In conclusion,  $(q_1, q_2, q_3) \in (0, 1)^3$  satisfies all the required conditions if and only if it is of the form  $(\lambda, 1 - 2\lambda, \lambda)$ , where  $0 < \lambda < 0.5$ .

For the second equation note that we have again

$$
E_Q[Y_2|Y_1 = 1 + y_j] = 1 + r \iff q_{j1} \times (1 + d) + q_{j2} \times (1 + m) + q_{j3} \times (1 + u) = 1 + r.
$$

Using the first part, we may thus conclude that  $(q_1, q_2, q_3)$ ,  $(q_{i,1}, q_{i,2}, q_{i,3})$ ,  $j = 1$ , 2, 3, describe a EMM for  $S^1$  if and only if they are of the form  $(\lambda, 1 - 2\lambda, \lambda)$ ,  $(\mu_j, 1 - 2\mu_j, \mu_j)$ , where  $0 < \lambda, \mu_j < 0.5, j = 1, 2, 3$ .

Note that while the condition for the martingale property is the same in each node, this condition is satisfied by many triplets, and we are allowed to choose a different triplet for each node. In other words, transition probabilities for *Q* need not be homogeneous across nodes, or equivalently put, we may choose  $\lambda$ ,  $\mu_1$ ,  $\mu_2$ ,  $\mu_3$  all different.