

# Applied Stochastic Processes

## Solution sheet 2

### Solution 2.1

- (a)
1.  $\mathbb{P}_x(X_1 = y, X_0 = x) = \delta^x(x) \cdot p_{xy} = p_{xy}$ .
  2.  $\mathbb{P}_x(X_1 = y) = \mathbb{P}_x(X_1 = y, X_0 = x) = \delta^x(x) \cdot p_{xy} = p_{xy}$ .
  3.  $P_x(X_1 = x, X_0 = y) = p_{xx}p_{xy}$ , which is not equal to  $p_{xy}$  in general.
  4.  $P_x(X_1 = x, X_0 = y) = \delta^x(y) \cdot p_{yx}$ , which is not equal to  $p_{xy}$  in general.
  5. Under the measure  $\mathbb{P}_x$ , we have  $X_0 \sim \delta^x$ . Thus,  $\mathbb{P}_\mu(X_1 = y) = \mathbb{P}_x(X_1 = y)$ .

Thus, the correct answers are 1, 2 and 5.

- (b) By definition of the  $n$ -step transition probability, we obtain

$$\begin{aligned} \mathbb{P}_x(X_{n+2} = z, X_{n+1} = y, X_n = x) &= p_{xx}^{(n)} \cdot p_{xy} \cdot p_{yz} \\ &= \sum_{x_1, \dots, x_{n-1} \in S} p_{xx_1} \cdot \dots \cdot p_{x_{n-1}x} \cdot p_{xy} \cdot p_{yz}. \end{aligned}$$

Thus, 3. and 5. are equal.

2. is also equal due to the 1-step and homogeneity property of Markov Chains:

$$\begin{aligned} \mathbb{P}_x(X_{2n+2} = z, X_{2n+1} = y, X_{2n} = x | X_n = x) &= \\ &= \mathbb{P}_x(X_{2n+2} = z | X_{2n+1} = y) \mathbb{P}_x(X_{2n+1} = y | X_{2n} = x) \mathbb{P}_x(X_{2n} = x | X_n = x) \\ &= \mathbb{P}_x(X_{n+2} = z | X_{n+1} = y) \mathbb{P}_x(X_{n+1} = y | X_n = x) \mathbb{P}_x(X_n = x) \\ &= \mathbb{P}_x(X_{n+2} = z, X_{n+1} = y, X_n = x). \end{aligned}$$

On the other hand, we can apply the simple Markov property with  $k = n$ ,  $Z = 1$  and  $f((X_{k+m})_{m \geq 0}) = \mathbb{1}_{X_{k+2}=z, X_{k+1}=y}$  to obtain

$$\mathbb{P}_x(X_{n+2} = z, X_{n+1} = y | X_n = x) = \mathbf{E}_x(\mathbb{1}_{X_2=z, X_1=y}) = \mathbb{P}_x(X_2 = z, X_1 = y) = p_{xy} \cdot p_{yz}.$$

Thus, 1. and 4. are equivalent to each other but not to  $\mathbb{P}_x(X_{n+2} = z, X_{n+1} = y, X_n = x)$ .

- (c) We have:

1.  $\mathbb{P}_1(X_2 = 3) = p_{12} \cdot p_{23} = 1/4$ .
2.  $\mathbb{P}_1(X_3 = 3) = 0$ , since the SRW is at even values at odd times.
3.  $\mathbb{P}_1(X_4 = 3) = \binom{4}{1} \cdot 1/16 = 1/4$ . We used that every nearest-neighbor path on  $\mathbb{Z}$  from 1 to 3 of length 4 does exactly 3 steps "+1" and 1 step "-1". Each such path has probability  $1/16$  and there are  $\binom{4}{1}$  ways to choose the position of the step "-1".
4.  $\mathbb{P}_1(X_5 X_6 < 0) = 0$ . Two consecutive values of the SRW cannot have opposite signs.
5.  $\mathbb{P}_1(X_{2n+1} = 0) = \binom{2n+1}{n} 2^{-2n-1}$  for  $n \in \mathbb{N}$ . Indeed, we can get to 0 after  $2n + 1$  steps by summing  $n + 1$  times "-1" and  $n$  times "+1". Each particular way of achieving this has probability  $2^{-n}$  of happening, and in total there are  $\binom{2n+1}{n}$  ways of choosing when the "+1" happen.

Thus, the probabilities that are 0 are 2 and 5.

**Solution 2.2** Let us identify the set  $a, b, c$  with  $1, 2, 3$ . Then, from the diagram we can get the following transition matrix

$$P = \begin{pmatrix} 0 & 1 & 0 \\ 0 & 1/2 & 1/2 \\ 1/2 & 0 & 1/2 \end{pmatrix}$$

We know that  $\mathbb{P}(X_n = a) = p_{a,a}^{(n)} = P^n(1,1)$ . Then we need to calculate  $P^n$ . We see that this matrix is diagonalizable since it has different eigenvalues. Indeed, its characteristic equation is given by

$$0 = \det(\lambda I - P) = \lambda \left( \lambda - \frac{1}{2} \right)^2 - \frac{1}{4} = \frac{1}{4}(\lambda - 1)(4\lambda^2 + 1)$$

and its eigenvalues are  $1, i/2, -i/2$ . Hence, there exists an invertible matrix  $U$  such that

$$P = U \begin{pmatrix} 1 & 0 & 0 \\ 0 & i/2 & 0 \\ 0 & 0 & -i/2 \end{pmatrix} U^{-1}$$

and then

$$P^n = U \begin{pmatrix} 1 & 0 & 0 \\ 0 & (i/2)^n & 0 \\ 0 & 0 & (-i/2)^n \end{pmatrix} U^{-1}$$

This implies that  $P^n(1,1) = x + y(i/2)^n + z(-i/2)^n$  for some constants  $x, y, z$ . We can calculate the value of these constants by using the first steps of our chain

$$\begin{aligned} 1 &= P^0(1,1) = x + y + z \\ 0 &= P^1(1,1) = x + iy/2 - iz/2 \\ 0 &= P^2(1,1) = x - y/4 - z/4. \end{aligned}$$

This gives us  $x = 1/5, y = (i - 2)/5$  and  $z = (2 - i)/5$ . Therefore

$$\begin{aligned} P^n(1,1) &= \frac{1}{5} + \frac{i-2}{5} \left(\frac{i}{2}\right)^n + \frac{2-i}{5} \left(\frac{-i}{2}\right)^n \\ &= \frac{1}{5} + \frac{i-2}{5} \left(\frac{1}{2}\right)^n \left(\cos \frac{n\pi}{2} + i \sin \frac{n\pi}{2}\right) + \frac{2-i}{5} \left(\frac{1}{2}\right)^n \left(\cos \frac{n\pi}{2} - i \sin \frac{n\pi}{2}\right) \\ &= \frac{1}{5} + \left(\frac{1}{2}\right)^n \left(\frac{4}{5} \cos \frac{n\pi}{2} - \frac{2}{5} \sin \frac{n\pi}{2}\right). \end{aligned}$$

**Solution 2.3** Define  $\mu$  to be the law of  $X_0$  and set

$$p_{xy} = \begin{cases} \mathbb{P}(X_{n+1} = y | X_n = x) & \text{if } \exists n : \mathbb{P}(X_n = x) > 0, \\ \mathbb{1}_{x=y} & \text{otherwise.} \end{cases}$$

By homogeneity,  $p_{xy}$  is well-defined. Furthermore, for every  $x_0, \dots, x_n \in S$ , we have

$$\begin{aligned} \mathbb{P}(X_0 = x_0, \dots, X_n = x_n) &= \underbrace{\mathbb{P}(X_0 = x_0)}_{=\mu(x_0)} \cdot \prod_{i=1}^n \underbrace{\mathbb{P}(X_i = x_i | X_0 = x_0, \dots, X_{i-1} = x_{i-1})}_{=\mathbb{P}(X_i = x_i | X_{i-1} = x_{i-1}) = p_{x_{i-1}x_i}} \\ &= \mu(x_0) \cdot p_{x_0x_1} \cdot \dots \cdot p_{x_{n-1}x_n}, \end{aligned}$$

where we used the 1-step Markov property and the definitions of  $\mu$  and  $P$ .

It remains to check that  $P$  is a transition probability. Let  $x \in S$ . If there exists  $n \geq 0$  such that  $\mathbb{P}(X_n = x) > 0$ , then

$$\sum_{y \in E} p_{xy} = \sum_{y \in E} \mathbb{P}(X_{n+1} = y | X_n = x) = 1.$$

Otherwise,

$$\sum_{y \in E} p_{xy} = \sum_{y \in E} \mathbb{1}_{x=y} = 1.$$

**Solution 2.4**

Under  $\mathbb{P}_0$ ,  $(X_n)_{n \geq 0}$  is a simple random walk (SRW) starting at 0. For  $i \in \mathbb{Z}$  and  $k \geq 0$ ,

$$\begin{aligned} \mathbb{P}_0(Z' = k | X_{10} = i) &= \mathbb{P}_0 \left[ \left( \sum_{n=10}^{20} \mathbb{1}_{X_n=i} \right) = k | X_{10} = i \right] = \mathbb{P}_i \left[ \left( \sum_{n=0}^{10} \mathbb{1}_{X_n=i} \right) = k \right] \\ &= \mathbb{P}_0 \left[ \left( \sum_{n=0}^{10} \mathbb{1}_{X_n=0} \right) = k \right] = \mathbb{P}_0(Z = k). \end{aligned}$$

Above, the second equality follows from the simple Markov property (Theorem 1.5 in the lecture notes with  $f(x_0, x_1, \dots) = \mathbb{1}_{\{\sum_{j=0}^{10} \mathbb{1}_{x_j=i} = k\}}$ ) and the third equality follows since  $(i + X_n)_{n \geq 0}$  is a SRW starting at  $i$  (under  $\mathbb{P}_0$ ). Since the right-hand side does not depend on  $i$ , it directly follows that

$$\mathbb{P}_0(Z' = k) = \sum_{i \in \mathbb{Z}} \mathbb{P}_0(Z' = k | X_{10} = i) \cdot \mathbb{P}_0(X_{10} = i) = \mathbb{P}_0(Z = k),$$

and so  $Z$  and  $Z'$  have the same distribution. Furthermore, we see that  $Z'$  and  $X_{10}$  are independent. The simple Markov property implies that  $Z$  and  $Z'$  are conditionally independent given  $\{X_{10} = i\}$ . Therefore,  $Z$  and  $Z'$  are independent as the following computation shows:

$$\begin{aligned} \mathbb{P}_0(Z = k, Z' = \ell) &= \sum_{i \in \mathbb{Z}} \mathbb{P}_0(Z = k, Z' = \ell | X_{10} = i) \cdot \mathbb{P}_0(X_{10} = i) \\ &= \sum_{i \in \mathbb{Z}} \mathbb{P}_0(Z = k | X_{10} = i) \cdot \mathbb{P}_0(Z' = \ell | X_{10} = i) \cdot \mathbb{P}_0(X_{10} = i) \\ &= \mathbb{P}_0(Z' = \ell) \cdot \left( \sum_{i \in \mathbb{Z}} \mathbb{P}_0(Z = k | X_{10} = i) \cdot \mathbb{P}_0(X_{10} = i) \right) \\ &= \mathbb{P}_0(Z' = \ell) \cdot \mathbb{P}_0(Z = k). \end{aligned}$$

**Solution 2.5**

- (a) We establish the inequality by induction on  $k$ . For  $k = 0$ , the inequality is trivial. For  $k \geq 1$ , it follows from the simple Markov property that

$$\begin{aligned} & \mathbb{P}_0(H_{-N,N} > k \cdot N) \\ &= \sum_{-N+1 \leq x_1, \dots, x_{(k-1)N} \leq N-1} \mathbb{P}_0(H_{-N,N} > k \cdot N, X_1 = x_1, \dots, X_{(k-1)N} = x_{(k-1)N}) \\ &= \sum_{-N+1 \leq x_1, \dots, x_{(k-1)N} \leq N-1} \mathbb{P}_{x_{(k-1)N}}(H_{-N,N} > N) \cdot \mathbb{P}_0(X_1 = x_1, \dots, X_{(k-1)N} = x_{(k-1)N}) \end{aligned}$$

Since the distance from any  $x \in \{-N+1, \dots, N-1\}$  to either  $N$  or  $-N$  is at most  $N$ , it follows that  $\mathbb{P}_x(H_{-N,N} \leq N) \geq 2^{-N}$ . Thus,

$$\begin{aligned} & \mathbb{P}_0(H_{-N,N} > k \cdot N) \\ & \leq (1 - 2^{-N}) \cdot \sum_{-N+1 \leq x_1, \dots, x_{(k-1)N} \leq N-1} \mathbb{P}_0(X_1 = x_1, \dots, X_{(k-1)N} = x_{(k-1)N}) \\ & = (1 - 2^{-N}) \cdot \mathbb{P}_0(H_{-N,N} > (k-1) \cdot N) \leq (1 - 2^{-N})^k, \end{aligned}$$

where we used the induction hypothesis in the last step.

We compute

$$\mathbf{E}_0(H_{-N,N}) = \sum_{\ell=0}^{\infty} \mathbb{P}_0(H_{-N,N} > \ell) \leq \sum_{k=0}^{\infty} N \cdot \underbrace{\mathbb{P}_0(H_{-N,N} > k \cdot N)}_{(1-2^{-N})^k} = N \cdot 2^N.$$

- (b) Assume towards a contradiction that  $\mathbf{E}_x(H_{-N,N}) = \infty$  for some  $x \in \{-N, \dots, N\}$ . Without loss of generality, let us assume that  $x$  is a non-negative integer. Then  $p_{0x}^{(x)} = 2^{-x}$ , and so by the simple Markov property

$$\begin{aligned} \mathbf{E}_0(H_{-N,N}) & \geq \mathbf{E}_0(H_{-N,N} \cdot \mathbb{1}_{X_1=1, \dots, X_x=x}) \\ & = \underbrace{\mathbf{E}_x((H_{-N,N} + x))}_{=\infty} \cdot \underbrace{\mathbb{P}_0(X_1 = 1, \dots, X_x = x)}_{=2^{-x}} = \infty, \end{aligned}$$

which contradicts the result of (a).

- (c) First, we note that by (b), the function  $f : \{-N, \dots, N\} \rightarrow \mathbb{R}_+$ , given by

$$f(x) = \mathbf{E}_x(H_{-N,N}),$$

is well-defined. Moreover,  $f$  is even (i.e.  $f(x) = f(-x)$ ) due to the symmetry of the SRW, and it has boundary values  $f(-N) = f(N) = 0$ . For  $x \in \{-N+1, \dots, N-1\}$ ,

$$\begin{aligned} f(x) &= \mathbf{E}_x(H_{-N,N}) = \mathbf{E}_x(H_{-N,N} \cdot \mathbb{1}_{X_1=x-1}) + \mathbf{E}_x(H_{-N,N} \cdot \mathbb{1}_{X_1=x+1}) \\ &= \mathbf{E}_{x-1}(H_{-N,N} + 1) \cdot \mathbb{P}_x(X_1 = x-1) + \mathbf{E}_{x+1}(H_{-N,N} + 1) \cdot \mathbb{P}_x(X_1 = x+1) \\ &= (f(x-1) + 1) \cdot \frac{1}{2} + (f(x+1) + 1) \cdot \frac{1}{2} = \frac{f(x-1) + f(x+1)}{2} + 1 \end{aligned}$$

Equivalently, for every  $x \in \{-N+1, \dots, N-1\}$ ,

$$f(x) - f(x-1) = f(x+1) - f(x) + 2.$$

Let  $n \geq 0$ . Summing over all  $x \in \{-n, \dots, n\}$ , it follows that

$$\begin{aligned} f(n) - f(-n-1) &= \sum_{x=-n}^n (f(x) - f(x-1)) \\ &= \sum_{x=-n}^n (f(x+1) - f(x) + 2) = f(n+1) - f(-n) + 2(2n+1). \end{aligned}$$

Thus, since  $f$  is even, we obtain

$$f(n) = f(n+1) + (2n+1).$$

Using  $f(N) = 0$ , we inductively obtain

$$\begin{aligned} f(n) &= \sum_{m=n}^{N-1} (2m+1) = 2 \cdot \left( \sum_{m=n}^{N-1} m \right) + (N-n) \\ &= 2 \cdot \left( \frac{N(N-1)}{2} - \frac{n(n-1)}{2} \right) + (N-n) \\ &= N^2 - n^2. \end{aligned}$$

In particular,  $f(0) = N^2$ , which is what we wanted to show.

*Remark:* Another strategy would be to show that the function  $g : \{-N, \dots, N\} \rightarrow \mathbb{R}_+$ , defined by

$$g(x) = f(x) + x^2,$$

is harmonic in the interior of  $\{-N, \dots, N\}$  and satisfies  $g(-N) = g(N) = N^2$ . Using the uniqueness of the solution to the Dirichlet problem, i.e. the fact that there is a unique harmonic function  $h : \{-N, \dots, N\} \rightarrow \mathbb{R}_+$  satisfying the boundary condition  $h(-N) = h(N) = N^2$ , it then follows that  $g(x) = N^2$  for every  $x \in \{-N, \dots, N\}$ . Thus,  $f(x) = N^2 - x^2$  for every  $x \in \{-N, \dots, N\}$ .